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INDICES OF QUALITATIVE VARIATION

Allen R. Wilcox

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Director's Division CIVIL DEFENSE RESEARCH PROJECT*

INDICES OF QUALITATIVE VARIATION

Allen R. Wilcox

OCTOBER 1967

OAK RIDGE NATIONAL LABORATORY Oak Ridge, Tennessee operated by UNION CARBIDE CORPORATION for the U.S. ATOMIC ENERGY COMMISSION

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Prefatory Note

The statistical exploration reported in this memorandum was undertaken to help us to better implement our responsibilities for the analysis of American public opinion on defense policy in general, and continental defense in particular. As familiarity with surveys will certify, many of the most interesting questions and sets of responses only meet nominal criteria, i.e., cannot be placed in relative positions on a scale. Accordingly, the analyst who wishes to treat these questions with a desirable degree of rigor and sophistication needs measures appropriate to them, in the author's phrase, measures of qualitative variation. This memorandum is a modest effort to assist the analyst who desires to improve his ability to explore essentially qualitative data. As such, we hope it will be helpful to a variety of social scientists who confront such data in the course of empirical research.

Davis B. Bobrow

INDICES OF QUALITATIVE VARIATION

Allen R. Wilcox

I. INTRODUCTION

In textbook treatments of measures of variation, statistics are always mentioned which measure the variation of a univariate distribution when the variable under consideration satisfies the requirements of an ordinal, interval, or ratio scale. Most commonly, the range, the semi-interquartile range, the average deviation, the standard deviation, and the variance are presented and discussed. However, the presentation and discussion of measures of variation suitable for use with variables that satisfy only the requirements of a nominal scale is often completely absent. In addition, there appears to be no unified discussion of those few appropriate measures scattered through the literature. Based on the assumption that such measures may have considerable utility for the statistical handling of qualitative data. this paper represents a first attempt to gather together and to generate alternative indices of qualitative variation. The treatment is introductory throughout. A more intensive mathematical and empirical treatment of these measures will hopefully be provided by others who wish to probe more deeply into the characteristics and utility of the statistics.

II. FORMAL PROPERTIES OF THE INDICES

The discussion is limited to measures that satisfy certain formal conditions. The first three of these conditions represent what are considered to be desirable formal properties of such indices. First and second, the maximum and/or minimum values that such an index may obtain should not depend on the magnitude of either of the two basic parameters of a qualitative distribution--the number of cases and the number of categories. These conditions facilitate comparison of the values of a particular index, even when they are derived from radically different distributions. Third, they must all have a standard range of values: in this case, from 0 to 1. This condition facilitates comparison among indices for the same distribution.

The final condition specifies the forms that a distribution must have when an index based on it obtains the maximum and minimum values. With the number of cases placed at 100 and the number of categories at 4, Fig. 1 illustrates with histograms the minimum-related and maximumrelated forms, respectively. Thus, the minimum value of 0 occurs if, and only if, all cases fall within <u>one</u> category. Conversely, the maximum value of 1 occurs if, and only if, an identical number of cases fall within <u>each</u> category.¹ This condition delineates the general "type" of measure to which this paper is being addressed. It is assumed that an index that does not satisfy this condition can be more usefully related to some other concept or idea. Measures that satisfy this condition can be thought of as indices of "generalized qualitative variation."²



Fig. 1

¹As a consequence of this condition, the indices discussed herein reach 1 only when the number of categories divides evenly into the number of cases. This circumstance impares the comparability of the indices in certain cases but does not diminish their general usefulness.

²For a description of a measure of qualitative variation that does not satisfy this condition, see George J. McCall and J. L. Simmons, "A New Measure of Attitudinal Opposition," <u>Public Opinion Quarterly</u>, 30:2 (Summer 1966) 271-278. A critique of this measure can be found in David Gold, "Critical Note on a New Measure of Attitudinal Opposition," Public Opinion Quarterly, 31:1 (Spring 1967) 76-79.

III. DESCRIPTION OF THE INDICES

MODVR

The first index that we have developed, MODVR, is an <u>index of devia-</u> tion from the mode, analogous to the standard deviation as a measure of deviation from the <u>mean</u>. The <u>core</u> of this index is the sum of the differences between the non-modal frequencies and the modal frequency, i.e.,

$$\sum_{i=1}^{K} (f_{i} - f_{i})$$

where $f_i =$ the frequency of the <u>i</u>th category, and

 $f_m =$ the frequency of the modal category.

Since $\sum_{i=1}^{K} f_i = N$ and $\sum_{m=1}^{K} f_m = K f_m$, the formula can be simplified to $K f_m - N$.

Dividing this expression by N(K-1) standardizes the range of the index (K = the number of categories and N = the number of cases).

This expression is directly related to Linton Freeman's variation ratio, \underline{v} .³ The formula for \underline{v} is $1 - \frac{f_m}{N}$. According to Freeman, the variation ratio is to be used as an aid in judging the adequacy of the mode as a summarizer of the distribution. It is simply the proportion of nommodal cases. The higher its value, the less adequate the mode is as a summarizer. As inspection of its formula will confirm, \underline{v} can achieve a minimum of 0, but it can never achieve a maximum of 1 because the modal frequency can never be 0. However, by dividing and subtracting by appropriate correction factors, $\underline{f_m}$ can be standardized to

 $\frac{\frac{1}{M} - 1}{\frac{N}{K} (K-1)}$ which reduces to $\frac{Kf_m - N}{M}$, the core of MODVR. MODVR, $\frac{N}{N}(K-1)$

in other words, is essentially a standardized variation ratio.

⁵Linton C. Freemen, <u>Elementary Applied</u> <u>Statistics</u>, New York: John Wiley and Sons, 1965, pp. 40-43. Finally, the core of this index (and of most of the others which follow) is subtracted from 1 so that low values will stand for low variation and high values for high instead of vice versa. Thus, $MODVR^4$ equals

$$1 - \frac{Kf_m - N}{N(K-1)}$$

The next five indices are all variations on a theme, for they are all basically analogs of measures that are used on ordinal or interval data. This is done by treating the frequencies of the K categories as values of a quantitative variable, calculating the particular statistic, and finally standardizing the result to conform to the criteria listed above. In this way indices are constructed that are analogous to the range, average deviation, mean difference, variance, and standard deviation.

RANVR

The first measure of this type that we have developed is based on the range and is designated RANVR. Its basic component is the difference between the highest or modal frequency and the lowest frequency. The formula⁵ follows: (f + f)

$$\frac{(f_m - f_{\ell})}{f_m}$$

where $f_m =$ the modal frequency, and $f_{\ell} =$ the lowest frequency.

AVDEV

The third index that we have generated, AVDEV, is an analog of the average or mean deviation, which is defined as the arithmetic mean of the absolute differences of each value from the mean. Utilizing this

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⁴A computational formula is $\frac{K(N-f_m)}{N(K-1)}$. ⁵A computational formula is $\frac{f}{f}$

concept on qualitative data, the formula 6 is

$$1 - \frac{\sum_{i=1}^{K} \left| f_{i} - \frac{N}{K} \right|}{2 \frac{N}{K} (K-1)}.$$

MNDIF

The next index that we have developed, MNDIF, is an analog of the mean difference, a measure of variation that is discussed much less frequently than the average deviation or standard deviation. It is defined as "the average of the differences of all the possible pairs of variate-values, taken regardless of sign."⁷ The mean difference differs from the average and standard deviations in that it is "dependent on the spread of the variate-values among themselves and not on the deviations from some central value."⁸ For qualitative data, this notion can be formulated as

$$L = \frac{\sum_{i=1}^{K-1} \sum_{j=i+1}^{K} \left| f_i - f_j \right|}{N(K-1)}.$$

⁶A somewhat simpler formula for computational purposes is

$$K\sum_{i=1}^{K} \left| f_{i} - \frac{N}{K} \right|$$

⁷Maurice G. Kendall and Alan Stuart, <u>The Advanced Theory of Statistics</u>, <u>Vol. 1</u>, New York: Hafner Publishing Company, 1958, p. 46.
⁸Tbid., p. 47. MNDIF can be approached in other ways mathematically. Suppose that the numbers 1 to K are assigned to the K categories <u>in ascending order</u> <u>of magnitude</u>. Given this weighting, two indices of variation can be derived. The first is based on subtraction and essentially sums the category weights multiplied by the differences between category proportions and the average proportion, i.e.,

$$1 - \frac{\sum_{i=1}^{K} \left(\frac{f_i}{N} - \frac{1}{K}\right)}{\kappa(1 - \frac{1}{K}) - \sum_{i=1}^{K-1} \frac{i}{K}}$$

The second is similar,⁹ but is based on division, i.e.,



We derived both of these indices from an intuitive conception of what seemed, on inspection of a number of distributions, to be useful measures of variation--specifically, measures that take into account differences among the frequencies of all K categories. Given the

 $\frac{9}{4}$ computational formula for both this and the previous formula is

$$2\sum_{i=1}^{K} i f_{i} - N(K+1)$$

$$1 - \frac{i=1}{N(K-1)}$$

similarity of this conception to the definition of a mean difference, it is in retrospect not surprising that they and MNDIF are algebraically equivalent. 10

VARNC

The fifth index, VARNC, is an analog of the variance, which is defined as the arithmetic mean of the squared differences of each value from the mean. For use with qualitative data, its formula¹¹ is

$$1 - \frac{\sum_{i=1}^{K} \left(f_i - \frac{N}{K} \right)^2}{\frac{N^2 (K-1)}{K}}$$

This formula is quite similar to AVDEV, the basic difference being that the differences are squared instead of having their absolute values taken.

¹⁰ It might be noted that another measure, Gini's index of concentration, is also equivalent to MNDIF when it is modified for qualitative data. For a discussion of Gini's index, see Hayward R. Alker and Bruce M. Russett, "On Measuring Inequality," <u>Behavioral Science</u>, 9:3 (July 1964) 207-218.

¹¹A computational formula is

ξ١.

$$1 - \frac{\sum_{i=1}^{K} f_i^2 - \frac{N^2}{K}}{N^2(K-1)}$$

which can be reduced even further to

$$\frac{K\left(N^{2}-\sum_{i=1}^{K}f_{i}^{2}\right)}{N^{2}(K-1)}$$

If MNDIF is approached in like manner and a squared version derived, the resulting formula is

$$L = \frac{\sum_{i=1}^{K-1} \sum_{j=i+1}^{K} (f_i - f_j)^2}{N^2(K-1)}$$

and is equal to VARNC.¹²

VARNC can be approached in another way through a discussion of Mueller and Schussler's Index of Qualitative Variation (IQV).¹³ They base this index on the idea that the total number of differences among the categories (or in their terminology, <u>items</u>) of a given variable provides a means of measuring qualitative variation. This total is obtained by counting the differences between each category and every other category and summing these differences. Thus, using sex as an example of qualitative variable, "if there were nine boys and three girls, each of the nine boys would differ from each of the three girls, producing 27 differences."¹⁴ They reduce this counting procedure to the following rule: multiply every category frequency by every other category frequency and sum these products. Although the authors do not

¹²This is similar to the case for quantitative data. As Kendall and Stuart note, a squared mean difference in the quantitative case is equal to twice the variance. Op cit., p. 47.

¹³John H. Mueller and Karl F. Schuessler, <u>Statistical Reasoning in</u> <u>Sociology</u>, Boston: Houghton Mifflin Company, 1961, pp. 177-179.

¹⁴<u>Ibid</u>., p. 177.

present a general formula,¹⁵ one can be constructed as follows:

$$\frac{\sum_{i=1}^{K-1} \sum_{j=i+1}^{K} \left(f_{i}f_{j} \right)}{\frac{K(K-1)}{2} \left(\frac{N}{K} \right)^{2}}$$

This formula is also algebraically equivalent to VARNC.

Finally, VARNC can also be derived by standardizing χ^2 . A typical formula for χ^2 is

$$\sum_{i=1}^{K} \frac{(O_i - E_i)^2}{E_i}$$

where O_i = the observed frequency for the <u>i</u>th category, and

 E_i = the expected frequency for the <u>i</u>th category. Substituting the notation used in this paper, the formula is

$$\sum_{i=1}^{K} \frac{\left(f_{i} - \frac{N}{K}\right)^{2}}{\frac{N}{K}}$$

¹⁵A computational formula:

$$\frac{2 \text{ K} \sum_{i=1}^{K-1} \sum_{j=i+1}^{K} (f_i f_j)}{N^2 (K-1)}$$

and, when standardized, it becomes



which reduces to VARNC.

STDEV

STDEV, the final analog index that we have developed, is based on the standard deviation, which is defined as the square root of the variance. As with VARNC, a formula can be derived by starting either with AVDEV or with MNDIF. From AVDEV, the formula¹⁶ is

$$1 - \sqrt{\frac{\sum_{i=1}^{K} (f_i - \frac{N}{K})^2}{(N - \frac{N}{K})^2 + (K-1)(\frac{N}{K})^2}}$$

From MNDIF, the formula is

$$1 - \sqrt{\frac{\sum_{i=1}^{K-1} \sum_{j=i+1}^{K} (f_i - f_j)^2}{N^2(K-1)}}$$

¹⁶ A computational formula:

$$1 - \sqrt{\frac{\sum_{i=1}^{K} f_i^2 - \frac{N^2}{K}}{\frac{N^2(K-1)}{K}}}$$

The seventh and last index of qualitative variation is HREL, a measure originally developed by engineers for use in specifying the properties of communications channels. The rationale for HREL is presented in terms of guessing by Virginia Senders (supplementing the mode as best guess):

> What we need is a measure of uncertainty, or of "poorness of a guess," which will be high when the number of alternative possibilities is high, and low when some of the possibilities are much more likely than others. One possible measure is the <u>average</u> <u>number of questions we have to ask to specify the</u> <u>correct alternative</u>.¹⁷

A formula for such a measure of uncertainty, H, is

$-\sum_{p_i} \log_2 p_i$

where p_i = the proportion of cases in the <u>i</u>th category.¹⁸ This is a measure of actual uncertainty. To obtain a measure of uncertainty which can be compared across distributions, actual uncertainty must be divided by the maximum possible uncertainty. Maximum possible uncertainty occurs when the alternatives are equally likely (i.e., when all category frequencies are the same) and is equal to $\log_2 N$. The resulting formula for a measure of relative uncertainty is

¹⁷Virginia L. Senders, <u>Measurement and Statistics</u>, New York: Oxford University Press, 1958, p. 79. For additional material on informational measures of this type, see Fred Attneave, <u>Applications of Information Theory to Psychology</u>, New York: Henry Holt and Co., 1959, and Henry Quastler (ed.), <u>Information Theory in Psychology</u>, Glencoe: The Free Press, 1955.

¹⁸The units in which H is measured are called bits. Senders has this to say about the term:

In (communications) channels and networks, relays and tubes are among the most important components. A typical relay can be either open or closed, a tube can fire or not. Thus, relays and tubes correspond to men giving answers to true-false or yes-or-no questions, and a binary (base 2) system seems to be the most convenient one to use. H is therefore measured in binary digits, which has been shortened to "bits."

 $\frac{-\sum_{p_i \log_2 p_i}}{\log_2 K}$

IV. SUMMARY AND CONCLUSION

The statistics described in the preceding pages provide a considerable range of alternative measures of generalized qualitative variation. Of these, only HREL and VARNC (differently formulated) have been taken entirely from the statistical literature. The idea behind MODVR was present in the form of a non-standardized index. The remaining four indices (RANVR, AVDEV, MNDIF, and STDEV) have not, to our knowledge, previously appeared in the literature. Because of this and because our objective has been one of preliminary presentation, a number of tasks remain for those who would continue the examination of indices of this type. First, inquiry into sampling distributions has evidently been conducted only for HREL. Second, further exploration into mathematical interrelationships might provide insights into the comparative theoretical relevance of the indices. Finally, and most importantly, an accumulation of knowledge on the comparative utility of these statistics in various research contexts would be highly desirable.

¹⁹For ease in computation using mathematical tables, one can use either of the following formulas:

 $- \underbrace{\sum_{i=1}^{p_i} \log_{10} p_i}_{\log_{10} K} - \underbrace{\sum_{i=1}^{p_i} \log_{e} p_i}_{\log_{e} K}$

APPENDIX A

Input Preparation, Sample Problem, and Output for the NOMSTAT Computer Program

This appendix presents a description of and sample input and output from NOMSTAT, a computer program which calculated the seven statistics described in the body of the text. Input--the input variables read by NOMSTAT are described in Table 1. The sequence in which these variables are read and the format by which they are read is given in Table 2.

Table 1. Nomstat Input Variable Description

Inpu	it Variable	Description
1.	TITLE	Alphanumeric title identification
2.	NDIST	Number of distributions
3.	NPROP	l if statistics are to be in proportion form, any other number if in percentage form
4.	NPCNT	l if category sizes are to be put out in frequency form, any other number if in percentage form
5.	NPUNCH	l if statistics are to be punched as well as printed, any other number if not
6.	NSUM	l if summary statistics (means, standard devia- tions, and coefficients of variation) are not to be calculated, any other number if they are
7.	NPCT	l if input is in percentage form, any other number if in frequency form
8.	NCAT	Vector containing the number of categories in each distribution
9.	NN	Vector containing the number of observations in each distributionused only if category sizes are in percentage form
10.	IVAR	Input FORMAT under which data matrix is read
11.	FREQ, or PCNT	Data Matrix. Which of these is read in depends on input variable 7

	Input Variables	Card Format	Number of Cards
1.	TITLE	FORMAT (20A4)	l
2.	NDIST, NPROP, NPCNT, NPUNCH, NSUM, NPCT	FORMAT (613)	1
3.	NCAT	FORMAT (4012)	5
4.	NN, only if NPCT = 1	FORMAT (2014)	10
5.	IVAR	FORMAT (20A4)	1
6.	PCNT (If NPCT = 1) or FREQ (If NPCT \neq 1)	IVAR	Variable

Table 2. Nomstat Input Variable Sequence and Format

Output - The output variables printed by NOMSTAT are described in Table 3. The sequence in which these variables are printed is given in Table 4. Table 3. Nomstat Output Variable Description

Outp	out Variable	Description
1.	TITLE	Alphanumeric title identification
2.	NDIST	Number of distributions
3.	NPROP	l if statistics are to be in proportion form, any other number if in percentage form
4.	NPCNT	l if category sizes are to be put out in frequency form, any other number if in percentage form
5.	NPUNCH	l if statistics are to be punched as well as printed, any other number if not
6.	NSUM	l if summary statistics (means, standard devia- tions, and coefficients of variation) are not to be calculated, any other number if they are
7.	NPCT	l if input is in percentage form, any other number if in frequency form
8.	NCAT	Vector containing the number of categories in each distribution
9.	NN	Vector containing the number of observations in each distributionused only if category sizes are in percentage form
10.	IVAR	Input FORMAT under which data matrix is read
11.	FREQ, or PCNT	Data matrix. Which of these is read in depends on input variable 7.
12.	AVDEV	Nominal statistic 1
13.	MNDIF	Nominal statistic 2
14.	STDEV	Nominal statistic 3
15.	VARNC	Nominal statistic 4
16.	HREL	Nominal statistic 5
17.	MODVR	Nominal statistic 6
18.	RANVR	Nominal statistic 7
19.	NN	Number of observations in each distribution
20.	FREQ, or PCNT	Data matrix. Which of these is read in depends on input variable 7.
21.	MEAN	Vector of the means of the nominal statistics
22.	SIGMA	Vector of the standard deviations of the nominal statistics
23.	COVAR	Vector of the coefficients of variation of the nominal statistics

Sequence	Output Variable	Sense Switch Control
1.	TITLE	
2.	NDIST, NPROP, NPCNT, NPUNCH, NSUM, NPCT	
3.	NCAT	
4.	NN	
5.	IVAR	
6.	FREQ, or PCNT	
7.	AVDEV, MNDIF, STDEV, VARNC, HREL, MODVR, RANVR	
8.	IVIN	
9.	FREQ, or PCNI	FREQ (If NPCNT=1), otherwise PCNT
10.	MEAN, SIGMA, COVAR	If NSUM ≠ 1

Table 4. Nomstat Output Variable Sequence and Sense Switch Control

Sample Problem - A sample problem has been used which is based on the responses of a cross-section of the American public to a question on foreign policy asked on 29 surveys over a period of several years.

NOMSTAT TEST CASE - ALLEN R. WILCOX

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	AVDEV	MNDIF	STDEV	VARNC	HREL	MODVR	RANVR	N	PERCE	NTAGE	<u>`S</u>
		•							,		
1	64.60	61.53	64.20	87.18	85.81	77.70	20.19	1305.	48.	42.	10.
· 2	68.03	63.08	66.90	89.04	88.44	76.59	24.56	1273.	49.	39.	12.
3	63.50	51.94	58.37	82.67	82.25	64.41	15.77	1211.	57.	34.	9.
4	53.09	43.55	50.26	75.26	76.54	53.09	12.62	1263.	65.	27.	8.
5	38.11	32.09	37.24	60.61	64.92	38.11	8.96	1287.	75.	19.	7.
6	62.28	49.88	56.60	81.16	80.63	62.55	14.04	1295.	58.	34.	8.
7	48.93	41.67	47.40	72.34	75.01	48.93	13.43	1260.	. 67.	24.	9.
8	50.56	43.80	49.19	74.18	76.90	50.56	15.23	1258.	66.	24.	10.
9	57.30	47.59	54.11	78.94	79.93	57.30	15.20	1288.	62.	29.	9.
10	65.26	60.57	64.33	87.27	86.17	75.60	20.51	1258.	50.	40.	10.
11	66.59	55.99	61.88	85.47	85.73	66.59	20.86	1302.	56.	33.	12.
12	64.64	56.60	61.99	85.56	84.73	70.26	18.36	1281.	53.	37.	10.
13	63.57	59.20	62.80	85.16	84.64	75.23	18.15	1282.	50.	41.	9.
14	66.89	62.85	65.49	88.09	86.98	78.39	22.19	1284.	48.	42.	11.
15	64.14	58.18	62.95	86.27	85.26	72.52	19.03	1272.	5ż.	39."	10.
16	62.55	52.58	58.76	82.99	83.62	62.55	18.66	1254.	58.	31.	11.
17.	65.32	55.34	61.25	84.98	84.62	67.69	18.61	1283.	55.	35.	10.
18	64.72	53.07	59.36	83.48	83.41	64.72	17.45	1270.	57.	33.	10.
19	56.07	46.87	53.80	78.65	78.89	58.07	13.30	1276.	61.	31.	8.
20	29.15	26.11	28.95	49.52	56.21	29.15	8.29	1302.	81.	13.	7.
21	39.37	33.10	38.40	62.06	66.11	39.37	9.29	1284.	74.	19.	7.
22	50.27	41.49	47.99	72.95	74.79	50.27	12.00	1304.	66.	26.	8.
23	31.29	28.94	31.17	52.63	59.29	31.29	10.21	1275.	79.	13.	8.
24	31.80	29.47	31.68	53.32	59.93	31.80	10.50	1269.	79.	13.	8.
25	26.00	24.44	25.95	45.16	52.65	26.00	8.60	1252.	83.	10.	7
- 26	38.65	32.33	37.69	61.17	65.25	38.65	8.84	1234.	74.	19.	7.
27	53.23	44.54	50.87	75.86	77.55	53.23	14.04	1237.	65.	26.	9.
28	44.72	39.21	43.90	68.53	72.37	44.72	13.38	1288.	70.	20.	9.
29	49.38	42.67	48.06	73.02	75.88	49.38	14.53	1282.	67.	23.	10.
MEAN	53.16	46.16	51.09	74.64	76.36	55.68	15.06				
SIGMA	12.88	11.56	11.98	12.64	10.01	15.67	4.42				
COVAR	0.24	0.25	0.23	0.17	0.13	0.28	0.29				

:

APPENDIX B

;

Fortran List of the NOMSTAT Computer Program

**FT'1.G.L.A.P.E. PRIGRAM NOMSTAT
C COMPUTE SEVEN NUMINAL MEASURES OF VARIATION FOR USE WITH
C UNIVARIATE DISTRIBUTIONS.
C DESCRIPTION OF PARAMETERS
C IVAR - VARIABLE FORMAT VECTOR
C NCAT - NUMBER OF CATEGORIES VECTOR
C FREQ - FREQUENCY MATRIX
$\frac{1}{C} = \frac{1}{2} = \frac{1}$
C NN - NUMBER OF OBSERVATIONS VECTOR
C SUM - STATISTICS SUMMATION VECTOR
C SUMSO - STATISTICS SQUARED SUMMATION VECTOR
C STAT - STATISTICS VECTOR
<u>C AVER – MEAN VECTOR</u>
C VAR - VARIANCE VECTOR
C SIG - STANDARD DEVIATION VECTOR
C XCUM - NUMBER OF ORSERVATIONS
C XMEAN - AVERAGE NUMBER OF OBSERVATIONS PER CATEGORY
C AVDEV - NUMERATOR TERM USED IN STAT(1)
C XMN - NUMERATOR TERM USED IN STAT(2)
C XSQ - NUMERATOR TERM USED IN STAT(3) AND STAT(4)
<u>C H – NUMERATOR TERM (ABSOLUTE UNCERTAINTY) USED IN</u>
C STAT(5) C HMAX - DENOMINATOR TERM (MAXIMUM UNCERTAINTY) USED IN
C STAT(5)
DIMENSION IVAR(100).NCAT(200).EREQ(50.200).PCNT(50.200).NN(200).
1PCT(50,200),SUM(7),SUMSQ(7),AVER(7),VAR(7),SIG(7),COVAR(7),STAT(7)
C READ IN TITLE CARD
<u> </u>
READ 111. (TITLE(1).(=1.20) 111 FURMAT(2044)
PRINT 112, (TITLE(I), I=1,20)
112 FORMAT(1H1,2044)
PRINT 57
57 FORMAT (1H1,43HCONTROL CARD AND VARIABLE FORMAT CARD INPUT)
C READ IN CONTROL CARD WITH SIX PARAMETERS IN 613 FORMAT
C 1 NDIST = NUMBER OF DISTRIBUTIONS
<u>C 2 NPROP = 1 IF STATISTICS ARE TO BE IN PROPORTION FORM, ANY</u>
C 3 NPCNT = 1 IF CATEGORY SIZES ARE TO BE PUT OUT IN
C FREQUENCY FORM, ANY OTHER NUMBER IF IN PERCENTAGE
C FORM C 4 NPUNCH = 1 IF STATISTICS ARE TO BE PUNCHED AS WELL AS
C PRINTED, ANY OTHER NUMBER LE NOT
C 5 NSUM = 1 IF SUMMARY STATISTICS (MEANS, STANDARD C DEVIATIONS, AND COEFFICIENTS OF VARIATION) ARE
C NOT TO BE CALCULATED, ANY OTHER NUMBER IF THEY
<u>C</u> ARF
C 6 NPCT = 1 IF INPUT IS IN PERCENTAGE FORM, ANY OTHER

	24
•	
	READ 15.NDIST.NPROP.NPCNT.NPUNCH.NSUM.NPCT
	15 FURMAT (613) Print 56. Ndist. NPrup. NPCNT. NPUNCH. NSUM. NPCT
	56 FORMAT (1H0,613)
<u>-c</u>	READ IN FIVE CONTROL CARDS IN 4012 FORMAT CONTAINING THE NUMBER
<u> </u>	OF CATEGORIES IN EACH DISTRIBUTION - LIMIT OF 200
ر س	L = 1
	M = 40 DO 16 I = 1.5
	READ 17, (NCAT(K), K=L, M)
	17 FORMAT(4012) PRINT 59. (NCAT(K).K=1.M)
	59 FORMAT (1H +4012)
	L = L+40 16 M = M+40
C	
<u> </u>	2014 FORMAT CONTAINING THE NUMBER OF OBSERVATIONS IN EACH
<u>_</u>	DISTRIBUTION - LIMIT OF 200
L	IF(NPCT.EQ.1)44,45
	44 L = 1 H = 20
	M = 20 DQ 46 I = 1,10
	READ 47.(NN(K),K=L,M)
	PRINT 62, (NN(K),K=L,M)
	62 FORMAT (1H +2014)
	46 M = M + 20
<u> </u>	45 CONTINUE
<u>_</u>	READ IN ONE VARIABLE FORMAT CARD
C.	READ 20+ (IVAR(I)+I=1+20)
	20 FORMAT(20A4)
	63 FORMAT (1H , 20A4)
<u> </u>	READ IN DATA ACCORDING TO VARIAN E EDRMAT IN EITHER EREDUENCY
<u>č</u>	OR PERCENTAGE FORM
Ċ	PRINT 61
	61 FORMAT (1H1,15HDATA CARD INPUT)
·	1100000000000000000000000000000000000
	NCT = NCAT(L)
	KEAU IVAR, (FREQ(I,L),I=1,NCT) 40 print IVAR, (Freq(I,L),I=1,NCT)
	41 UU 42 L = 1.NUISI NCT = NCAT(L)
	READ IVAR. (PCNT(I,L),I=1,NCT)
	42 PRIME IVARA (PUNICIALIATEIANCI) 43 CONTINUE
C	
<u> </u>	FRINT OUT REAUING
	IF(NPCNT.NE.1)G() TO 23 PRINT 13
	TOTAL 12

.

-4

L.

	13 FORMAT(1)1.5X.69H AVDEV MNDIE STDEV VARNC HREL MODVR RANVR
·	
	23 DEINT 25
	25 FRINT 25
<u> </u>	1 N DECENTACES/1
c	I N PERCENTAGES///
<u></u>	THE TALLESS YO O TWO VESTORS - SUM ASUMMATION OF THE NONTHAL
L C	INITIALIZE TO O TWO VECTORS - SUM (SOMMATION OF THE NUMINAL
<u></u>	STATISTICS) AND SUMSU (SUMMATION OF THE NUMINAL STATISTICS
L	SOUARED
<u> </u>	
	24 DO 19 I = 1.7
	19 SUM(I) = SUMSO(I) = 0
С	
C	BEGIN MAJOR DO LOOP - CALCULATION OF STATISTICS
C	
	DO 50 K = 1 , NDIST
C	
č	CALCULATE NUMBER OF OBSERVATIONS (XCUM) AND PERCENTAGES (PCNT)
<u> </u>	TE EREQUENCIES ARE INPUT
č	
<u> </u>	NCT = NCAT(K)
	5 XCUM = XCUM + FREQ(1,K)
	GO TO 66
	49 xcum = NN(K)
	66 DO 30 I = 1.NCT
	30 PENT(I+K) = FREQ(I+K)/XCUM*100.
С	
C	SUM NUMERATOR TERMS FOR STATISTICS 1-5
С	
	XMEAN = XCUM/NCT
	AVDEV = XMN = XSQ = H = C
	$DO = 6$ M = 1 \star NCT
	x = x + p = x + p = x + p = x + (p = x + p =
	$F = \{\mu_{1}, \mu_{2}, \mu_{3}, \mu_$
	B AVDEV - 4VDEVTABSF(PCNTENTKEVTACON/IOOCANEANT
	$DD 6 MPP = MP \cdot NC1$
	$6 \times MN = \times MN + ABSF(PCNT(M,K) \neq XCUM/100 - PCNT(MPP,K) \neq XCUM/100)$
С	
<u> </u>	CALCULATE STATISTICS 1-5
C	
	NCT1 = NCT-1
	STAT(1) = 100-50*NCT*AVDEV/(XCUM*NCT1)
	STAT(2) = 100-100*XMN/(XCUM*NCT1)
	STAT(3) = 1CO-1CO*SORTF(XSC)/SORTF(XCUM**2*NCT1/NCT)
	STAT(4) = 100-100*NCT*XSQ/(XCUM**2*NCT1)
	XCT = NCT
	HMAX = LOGE(XCT)/IOGE(2a)
	$STAT(5) = H \pm 100/HMAX$
ſ	
ř	CALCULATE WORKING VECTOR TO RESERVE ODDED OF DEDCENTAGES FOR
r r	UNERVENTER WORKING VECTOR TO PRESERVE DRUCK OF FERGENTAGES FUR
<u> </u>	
L	
	DU 64 I = I+NCI
_	64 PCT(I,K) = PCNT(I,K)
<u> </u>	
C	FIND LARGEST AND SMALLEST CATEGORY PERCENTAGES FOR USE IN

С		STATISTICS 6-7
Č		
		DO 100 I = $1.NCT1$
		IF (PCNT(I+1.K) .GE. PCNT(I.K)) 100, 101
	101	EXCH = PCNT(I+1,K)
		P(N)(1+1)K) = P(N)(1)K
	100	
	100	NCT2 = NCT-2
		$DO \ 102 \ I = 1 \cdot NCT2$
		IF (PCNT(I+1,K) .LE. PCNT(I,K)) 102,103
	103	EXCH = PCNT(I+1,K)
		PCNT(I+1,K) = PCNT(I,K)
		$PCNT(I \cdot K) = EXCH$
	102	CONTINUE
لم م		ሮአር ምክር ቀላይ ይኖ ለተዋወተቸው ራ… ን
$\frac{l}{r}$		CALCULATE STATISTICS OF
U		STAT(6) = $NCT*(100-PCNT(NCT.K))/NCT)$
		$STAT(7) = 100 * PCNT(NCT1 \cdot K) / PCNT(NCT \cdot K)$
С		
. C		CALCULATE PROPORTIONS IF REQUESTED
<u> </u>		
		IF(NPROP.NE.1)GO TO 82
		$DO \ 80 \ I = 1,7$
~	ងប្	SIAI(I) = SIAI(I)/I00
-		SUM THE STATISTICS AND THE STATISTICS SOURED
č		SON THE STATISTICS AND THE STATISTICS SWORKED
	82	DO 21 I = 1,7
		SUM(I) = SUM(I) + STAT(I)
	21	SUMSQ(1) = SUMSQ(1)+STAT(1)**2
<u> </u>		
C	•	PRINT MAIN OUTPUT IN EITHER FREQUENCY OR PERCENTAGE FORM
Ľ	.	
		IF (NPCT, FO, 1)51, 52
•	51	DO 53 1 = 1.NCT
	53	FREQ(I,K) = PCT(I,K) * XCUM/100
	52	PRINT 73, K, (STAT(I), I=1, 7), XCUM, (FREQ(I,K), I=1, NCT)
	73	FORMAT(16,7F7.2,F6.C,2X,10F5.0,4(/1H+,62X,10F5.0))
		GO TO 72
		PRINT 12.K. (STAT(I), I=1, 7), XCUM, (PCT(I,K), I=1, NCT)
~	12	FURMAI(16+7F7+2+F6+0+2X+10F5+0+4(71H++62X+10F5+J))
		PUNCH STATISTICS IF REQUESTED
č		
	72	IF(NPUNCH.NE.1) GO TO 50
		PUNCH 14.K.(STAT(I),I=1.7)
	14	FORMAT(14,7(F7.2))
	50	CONTINUE
C		
<u>_</u>		AND COEEECCENTS OF VARIATION (COMARY DEVIATIONS (SIG)
с С		AND COEFFICIENTS OF VARIATION (COVAR) IF REQUESTED
		IE(NSUM_EQ.1)GO TO 110
	•	$DQ_{22}I = 1,7$
		AVER(I) = SUM(I)/NDIST
-		VAR(I) = (SUMSQ(I)/NDIST) - AVER(I)**2
		SIG(I) = SORTF(VAR(I))
	22	COVAR(I) = SIG(I)/AVER(I)
		PRINT 9, (AVER(I), I=1,7), (SIG(I), I=1,7), (COVAR(I), I=1,7)
	9	FORMAT(1H0.5H MEAN,7F7.2/X,5HSIGMA.7F7.2/X,5HCOVAR,7F7.2)
	110	
		END NUMSTAT

Ŀ

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