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Time Series Analysis by the Maximum Entropy Method

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THE UNIVERSITY OF CHICAGO
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TIME SERIES ANALYSIS BY THE MAXIMUM
ENTROPY METHOD

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ABSTRACT

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The principal subject of this report is the use of the Maximum Entropy method for spectral analysis of time series. The classical Fourier method is also discussed, mainly as a standard for comparison with the Maximum Entropy method. Examples are given which clearly demonstrate the superiority of the latter method over the former when the time series is short. The report also includes a chapter outlining the theory of the method, a discussion of the effects of noise in the data, a chapter on significance tests, a discussion of the problem of choosing the prediction filter length, and, most importantly, a description of a package of FORTRAN subroutines for making the various calculations. Cross-referenced program listings are given in the appendices. The report also includes a chapter demonstrating the use of the programs by means of an example. Real time series like the lynx data and sunspot numbers are also analyzed.

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INTRODUCTION

In the past, the spectral analysis of time series has for the most part utilized two classical approaches - the power spectrum method of R. B. Blackman and J. W. Tukey (1959) and the periodogram. The Blackman and Tukey approach involves the estimation of the autocorrelation function. A shortcoming of the estimator used is its assumption of a zero extension of the given data. On the other hand, the periodogram method assumes a periodic extension of the data, the period being equal to the length of the time series. John P. Burg (1968) has presented a new technique of time series analysis which seeks to minimize assumptions about the unavailable data. The method is called maximum entropy spectral analysis (which will be referred to as MESA throughout this report). MESA emerged in an effort to achieve better resolution from short time series.

A great amount of research on MESA has been done by T. J. Ulrych. His recent publications (1972b, 1973, 1974) indicate the wide extent to which MESA can be utilized. Ulrych's contention is that the classical Fourier transform methods are inferior to MESA because of their inability to analyze low frequency variations.

MESA and Fourier analysis are compared in this report. Generally speaking, MESA is superior for short time series, but it has one major disadvantage shared also by older methods. The problem involves the determination of the length of the vector of filter coefficients used for estimating the autocorrelation function. In an attempt to solve this problem, T. J. Ulrych (1975) suggested the use of the Akaike final

prediction error (FPE), (Akaike 1969), as applied to autoregressive decomposition. A brief critical discussion of this idea is also presented here in Chapter IV.

As much as possible, the mathematical presentation of time series analysis is simplified in this report. The derivation of the maximum entropy spectrum, however, employs theorems in complex analysis. Although the mathematics becomes rigorous in some parts, it is hoped that the many examples presented will impart a fair understanding of the methods discussed to less mathematically inclined readers.

Computer programs (written in FORTRAN) which calculate both the Fourier sample spectrum and maximum entropy spectrum plus some other useful information are described and explained in Chapter V. An example is presented in Chapter VI to illustrate the use of these programs. Program listings are given in the appendices.

CHAPTER I

THE DISCRETE FOURIER TRANSFORM METHOD

Consider the functions $y_1(t)$, $y_2(t)$, $y_3(t)$ in Figures 1.1-1.3, respectively.

Fig. 1.1 The function $y_1(t)$: amplitude = 2, period = 6 years.

Fig. 1.2. The function $y_2(t)$: amplitude = 4, period = 4 years.

Fig. 1.3. The function $y_3(t)$: amplitude = 3, period = 12 years.

The sum of $y_1(t)$, $y_2(t)$, $y_3(t)$ produces a new function $y(t)$ (Fig. 1.4).

Fig. 1.4. The new function $y(t) = y_1(t) + y_2(t) + y_3(t)$, sampled at intervals $\Delta t = 1$ year.

Since $y(t)$ is the sum of three functions whose periods are known, the analysis of its periodic components is clear. In practice, one is usually faced with the inverse problem, i.e., the analysis of a time series like $y(t)$ in order to determine what periodic components are present. This situation leads to the utilization of the Fourier transform method.

Although many natural processes are truly harmonic, in that they can be expressed exactly as a sum of sinusoidal terms, it is more often the case that a natural process will also contain aperiodic components. Such processes can still be analyzed by harmonic methods. In fact, any aperiodic function of finite length can be approximated to any desired degree of accuracy by a sum of sinusoids, provided one includes enough sinusoidal terms in the approximating function. In most real problems one is limited to a finite, discrete time series obtained by sampling the process at say n equally spaced points in time. If the process is denoted by $x(t)$, and if the sampling is arranged so that the sampling interval is one unit of time, i.e., $\Delta t = 1$, with the first sample point being taken at $t = 0$, then the time series can be written $x(0)$, $x(1)$, $x(2)$, ..., $x(n-1)$, or as is done more commonly, x_0 , x_1 , x_2 , ..., x_{n-1} . It can be shown that the terms of this series can be represented exactly by a finite Fourier expansion of the form

$$(1.1) \quad x_j = a_0 + \sum_{0 < k < n/2} \left(a_k \cos \frac{2\pi jk}{n} + b_k \sin \frac{2\pi jk}{n} \right) + (-1)^j a_{n/2},$$

$$j = 0, 1, 2, \dots, n-1,$$

where the $a_{n/2}$ term is included only if n is even. The Fourier coefficients are given by

$$(1.2) \quad a_0 = \frac{1}{n} \sum_{j=0}^{n-1} x_j ,$$

$$(1.3) \quad a_k = \frac{2}{n} \sum_{j=0}^{n-1} x_j \cos \frac{2\pi jk}{n} ,$$

$$(1.4) \quad b_k = \frac{2}{n} \sum_{j=0}^{n-1} x_j \sin \frac{2\pi jk}{n} ,$$

$$0 < k < \frac{n}{2}$$

$$(1.5) \quad a_{n/2} = \frac{1}{n} \sum_{j=0}^{n-1} (-1)^j x_j , \quad \text{if } n \text{ is even}$$

Equation (1.1) expresses the time series x_j in terms of the component frequencies $f_k = k/n$ and can also be written in terms of the equivalent periods $T_k = 1/f_k = n/k$ as

$$(1.6) \quad x_j = a_0 + \sum_{0 < k < n/2} (a_k \cos \frac{2\pi j}{T_k} + b_k \sin \frac{2\pi j}{T_k}) + (-1)^j a_{n/2} .$$

The Fourier spectrum of the time series is the discrete function represented by the set of points (f_k, F_k) , $k=0,1,\dots,[n/2]$, where $[n/2]$ is taken to be $n/2$ if n is even and $(n-1)/2$ if n is odd, and F_k is the Fourier amplitude defined by

$$F_k = (a_k^2 + b_k^2)^{1/2}$$

For each of the frequencies f_k one can also define an associated phase by

$$(1.8) \quad \theta_k = \tan^{-1} (a_k/b_k) \quad .$$

The expansion (1.6) can be written in terms of the Fourier amplitudes and phases as

$$(1.9) \quad x_j = a_0 + \sum_{0 \leq k < n/2} F_k \sin \left(\frac{2\pi j}{T_k} + \theta_k \right) + (-1)^j a_{n/2} \quad .$$

Again, this expansion is exact for all the sample points $x_0, x_1, \dots, x_{n-1}$, and if n is large enough one might reasonably hope that Equation (1.9) would give a good approximation to the underlying process $x(t)$ at other values of t also. That is, for any value of t in the interval $0 \leq t < (n-1)$, or even for t outside that interval, one might hope that

$$(1.10) \quad x(t) \approx a_0 + \sum_{0 \leq k < n/2} F_k \sin \left(\frac{2\pi t}{T_k} + \theta_k \right)$$

without too much error.

The function (f_k, F_k) is the discrete Fourier spectrum of the data set $x_0, x_1, \dots, x_{n-1}$. We note that continuous functions $x(t)$ have analogous continuous Fourier spectra, but we do not treat such cases here since we are concerned only with finite, discrete time series. For the continuous case the reader is referred to Jenkins and Watts (1968). In computing a discrete Fourier spectrum one hopes that

the result approximates the continuous Fourier spectrum of the underlying process $x(t)$. The approximation is not limited to the discrete frequencies $f_k = k/n$ but can be calculated for intermediate values of f by the formula

$$(1.11) \quad F(f) = \sqrt{a(f)^2 + b(f)^2},$$

where

$$(1.12) \quad a(f) = \frac{2}{n} \sum_{j=0}^{n-1} x_j \cos 2\pi j f,$$

$$(1.13) \quad b(f) = \frac{2}{n} \sum_{j=0}^{n-1} x_j \sin 2\pi j f.$$

We note in passing that some writers make a distinction between the continuous function $F(f)$ and the discrete function (f_k, F_k) , calling the latter a "line spectrum" and the former a "sample spectrum". Whatever it is called, the continuous function $F(f)$ hopefully approximates the Fourier spectrum of $x(t)$. The approximation fails to exactly represent the desired spectrum because the time series $x_0, x_1, \dots, x_{n-1}$ cannot exactly represent the process $x(t)$. The error arises for two reasons:

(1) The time series is restricted to a finite interval of time, so we do not know what the process $x(t)$ is like outside that interval, and

(2) The time series represents the process only at a finite number of points in the interval, adjacent points being separated by the sample spacing Δt .

The latter restriction limits the highest frequency that is resolvable - one cannot expect to isolate components with frequencies whose corresponding periods are smaller than the sampling interval. In fact it is necessary to have at least two samples per cycle of the highest frequency (shortest period) that can be resolved. The frequency

$$(1.14) \quad f_N = 1/(2 \Delta t)$$

is called the Nyquist frequency and is the highest frequency that can be resolved with a sample interval of Δt . The situation is actually worse than it appears on the surface because, if $x(t)$ does have important harmonic components at frequencies greater than the Nyquist frequency, then these components will distort the Fourier spectrum values at lower frequencies. This is because the approximation cannot distinguish between the frequency f and the frequency $(1/\Delta t - f)$. Thus, if $\Delta t = 1$, then the Nyquist frequency is 0.5 and if $x(t)$ has a harmonic component with frequency $f = 0.75$, the estimate of the transform will be distorted at $f = 0.25$. This phenomenon is called aliasing and must be kept in mind in designing a sampling procedure. One should always try to choose Δt small enough so that f_N exceeds the frequencies of all the important harmonic components in the process.

The first restriction, i.e., the restriction to a finite time interval, produces distortions in the calculated spectrum at the low frequencies (long periods). Obviously one cannot expect to resolve frequencies corresponding to periods much longer than the total time interval represented by the sample. Most users of the Fourier method do not trust the estimate at frequencies lower than that corresponding to

periods of at most one half the total time interval, and very conservative users demand that the total interval exceed a given period by a factor of 4 or more before they will trust the Fourier estimate at the corresponding frequency. The Fourier technique can detect the presence of long period components (components with periods comparable to the total length of record), but it cannot accurately pinpoint the frequency nor can it distinguish between two or more long period components at frequencies that are relatively close together.

It is convenient to analyze the effect of a truncated record in terms of a function called a data window. Consider a process $x(t)$ defined on the time interval $(-\infty, +\infty)$.

Fig. 1.5. The observed function $x_0(t)$, $\Delta t = 1$.

In general, one can only observe within a finite interval of length T , say $[-T/2, T/2]$ as in Figure 1.5. The available data are then confined to the interval $[-T/2, T/2]$. The rectangular data window for the interval $[-T/2, T/2]$ is defined by

$$(1.15) \quad w(t) = \begin{cases} 1, & \text{if } -T/2 \leq t \leq T/2, \\ 0, & \text{otherwise.} \end{cases}$$

The observed function corresponding to the available data is represented by

$$(1.16) \quad x_0(t) = x(t) w(t).$$

We temporarily ignore the effect of the discrete sampling and concentrate on the effect of applying the rectangular data window, i.e., the effect of truncating the record. Subjecting the observed function $x_0(t)$ to Fourier analysis will result in a transform $\mathcal{F}_{x_0}$ of x_0 which is really a smoothed version of the Fourier transform $\mathcal{F}_x$ of the original signal $x(t)$. That is, if we let

$$\begin{aligned} \mathcal{F}_{x_0} &= \text{Fourier transform of } x_0, \\ \mathcal{F}_x &= \text{Fourier transform of } x, \text{ and} \\ \mathcal{F}_w &= \text{Fourier transform of } w, \end{aligned}$$

then it can be shown that

$$(1.17) \quad \mathcal{F}_{x_0}(f) = \int_{-\infty}^{\infty} \mathcal{F}_x(f') \mathcal{F}_w(f - f') df'.$$

The Fourier transform of the truncated, observed time series is said to be the convolution of the desired Fourier transform with the Fourier transform of the data window. This latter function is called the spectral window and can be shown to be

$$(1.18) \quad \mathcal{F}_w(f) = T \frac{\sin 2\pi ft}{\pi ft} .$$

Applying the data window (1.15), which is sometimes called the boxcar window, necessarily distorts the Fourier transform. It can be shown that $\mathcal{F}_{x_0}(f)$ is necessarily smoother than $\mathcal{F}_x(f)$. A common practice, therefore, is to change the shape of the data window in order to reduce the distortion in the transform. This is usually done by rounding off the corners of the boxcar window, thus damping the contribution of the data at the two ends of the record. The idea is to design a window which decreases the smoothing and hence increases the resolution of the transform. This procedure is not without its own shortcomings. Spectral windows which give good resolution of the spectral peaks are often unstable to changes in the input data. On the other hand, those spectral windows that are stable with regard to changes in the data usually give poor resolution.

A more complete discussion of data windows and their corresponding spectral windows is given by Box and Jenkins (1970). We close our discussion of the subject by noting that using any data window involves some assumption about the data outside the available record (e.g., the boxcar window assumes a periodic extension of the data with period equal to the length of the record). Furthermore, using any window, other than

the boxcar window, distorts the data that are available. As we shall see in the next chapter, the Burg maximum entropy method is motivated by an attempt to overcome these two difficulties.

As an illustration of some of the ideas incorporated in this chapter, we present a very simple example. The number of points sampled is $n = 42$. The function sampled was the sum of a straight line trend and a sinusoid. The linear trend was included in the example because real time series data frequently contain trend components.

The data points in Figure 1.6 are generated from the equation

$$x_i = \int_{i-1}^i h(t) dt, \quad i = 1, 2, \dots, 42,$$

where

$$h(t) = 7 + t + 5 \cos \frac{2\pi t}{6}$$

The linear trend, $x_T(t) = 7 + t$, was first removed by performing a straight line, least squares fit to the data and subtracting out the result. The function subjected to Fourier analysis was then $x(t) - x_T(t)$. It is necessary to remove the trend line because Fourier analysis would interpret it as a component with infinite period. The spectrum of the detrended function is shown in Figure 1.7.

The highest peak in Figure 1.7 occurs at $f = 0.1675$ with a corresponding period of $P = 5.97$, which is a good estimate of $P = 6$, the true period of the function. The other smaller peaks are called sidelobes and do not represent significant spectral components. They are one of the results of truncating the record. The width of the main peak also depends on the record length and gives an idea of the uncertainty in the estimate of the peak frequency.

MATH TEST CASE

Fig. 1.6. Data $x(t)$ are histograms; yearly integrals of

$$h(t) = 5 \cos \frac{2\pi t}{6} + 7 + t.$$

FOURIER SPECTRUM

Fig. 1.7. Fourier Spectrum of detrended $x(t)$.

CHAPTER II

MAXIMUM ENTROPY SPECTRAL ANALYSIS

A. A Brief Explanation of the Entropy Function.

In English, when one expresses a simple, grammatically correct, logical statement, his choice of words is governed by probability laws. For example, if a statement starts with "The mountain", one cannot just append any English phrase to have the statement make sense. "The mountain sees" or "The mountain my" do not make sense. However, "The mountain is big" is a perfectly logical statement. Thus, a person's choice of words is strongly dependent upon the logical structure of the language, and hence there is a probability that certain words are more likely than other words to follow a phrase like "The mountain". The freedom of choice in selecting a message is what is scientifically termed as information or entropy.

The word "information" must not be confused with its ordinary meaning. In the physical sciences, the entropy or information associated with a situation is a measure of the degree of randomness in the situation. Entropy is expressed in terms of probabilities - for example, the probability of getting to a certain stage in the process of forming a message and when in that stage the probability that certain words will be chosen next (Shannon and Weaver 1949).

How then can one measure the entropy or information of a process? In particular, the processes treated here are Markoff processes; that is, stochastic processes in which the probabilities depend on previous

events. Shannon and Weaver (1949) have shown that such a measure for information can best be estimated using logarithms of the probabilities.

Suppose $p_1, p_2, \dots, p_n$ are the known probabilities of occurrence of a set of discrete events. Then

$$(2.1) \quad I = I(p_1, p_2, \dots, p_n) = -K \sum_{i=1}^n p_i \log_a p_i \quad ,$$

where K is a positive constant, and a is the chosen base for the logarithm, is called the entropy of the set of probabilities $p_1, p_2, \dots, p_n$. Similarly, the entropy of a continuous distribution with density distribution function $p(x)$ is defined by

$$(2.2) \quad I = I(p(x)) = - \int_{-\infty}^{\infty} p(x) \log p(x) dx \quad .$$

Assume for example that there are two possibilities with probabilities p and $q = 1-p$. Then the entropy is

$$(2.3) \quad I = I(p) = -p \log_e p - (1-p) \log_e (1-p) \quad ,$$

where for convenience the natural logarithm is used, and the constant K is taken to be 1. A plot of I in Equation (2.3) appears in Figure 2.1.

I is maximized when $q = 1-p$, that is $q=p=\frac{1}{2}$. In a physical situation this is the case of greatest uncertainty. In tossing a coin, for

Fig. 2.1. The entropy I as a function of the probability p when there are two possible outcomes to an event.

example, out of 1000 tosses, it is likely that heads will appear about 500 times and tails about 500 times, but the uncertainty for any given toss is maximized. For a biased coin the uncertainty for any given toss is less because if one simply predicted the most likely result on each toss, then one would be right more than 500 times out of 1000. The more uncertain the event is, the greater the freedom of choice in predicting the outcome. The function I in Equation (2.3) is 0 when $p=0$, $q=1$ or $q=0$, $p=1$. That is I assumes a minimum of 0 when the event is certain to happen. In this case, one has little freedom of choice in making a prediction. In all other cases, I assumes a positive value.

The logarithm is used in the definition of entropy since it is mathematically more convenient to handle. With a logarithmic definition, information is additive, as would be expected. Assume for example that there are two independent events with probabilities p_1 and p_2 . Then the probability that both events happen is $p_1 p_2$. The measured information or entropy of the two events is

$$(2.4) \quad I = -K \log_a(p_1 p_2) = I_1 + I_2 \quad ,$$

where

$$(2.5) \quad I_1 = -K \log p_1 \quad , \quad I_2 = -K \log p_2 \quad .$$

The most common logarithmic bases employed in the definition of entropy are 2, 10 and e . The natural logarithm is the most convenient for use in mathematical analysis and is the one we adopted here.

From the preceding discussions, we note that I is indeed a reasonable measure of entropy. To maximize I is to maximize the uncertainty in the event. The constraints involved depend on the probabilities $p_1, p_2, \dots, p_n$ with $0 \leq p_i \leq 1$, $i = 1, 2, \dots, n$ and $\sum_{i=1}^n p_i = 1$.

In time series analysis, the relevant probabilities are not known. Instead, one is faced with the analysis of the power spectrum. In this case, the entropy is proportional to

$$(2.6) \quad I = \int_{-\infty}^{\infty} \ln P(f) df \quad ,$$

where $P(f)$ is the power spectrum (see Burg 1975, Shannon and Weaver 1949).

In looking at the power spectrum of a time series, one asks how much of the original information has been lost or gained. The degree of uncertainty is measured in the entropy function I of Equation (2.6).

R. Maximum Entropy Spectral Analysis (MESA)

The basic assumption in MESA is that the stationary time series under analysis is the most random or least predictable. The classical approach involves the estimation of the power spectrum from known values of the autocorrelation function ρ_k , $|k| \leq M$, where M is less than or equal to the number of observations in the given time series. The standard assumption is that $\rho_k = 0$ for $|k| > M$, but usually a weighting function is also introduced and multiplied by ρ_k and the result is Fourier transformed. Given a set of autocorrelation function values ρ_k with the imposed condition that the power spectrum be non-negative definite, there usually exist infinitely many power spectra which will be consistent with the given data. Maximum entropy spectral analysis is based on the idea of choosing the spectrum which corresponds to the most random time series whose autocorrelation function agrees with given values.

C. Statement of the Problem

Find a real positive function $P(f)$ which maximizes the entropy function

$$(2.7) \quad I = \int_{-f_M}^{f_M} \ln P(f) df$$

under the constraints

$$(2.8) \quad \rho_k = \int_{-f_M}^{f_M} P(f) \exp(i2\pi f k \Delta t) df, \quad$$

where $-M \leq k \leq M$, Δt is the sampling interval, and ρ_k the autocorrelation with time lag $k\Delta t$, $i = \sqrt{-1}$ and $f_M = \frac{1}{2\Delta t}$ is the Nyquist frequency.

D. Solution to the Problem

Wiener has shown that the power spectrum $P(f)$ is the discrete Fourier cosine transform of the autocorrelation ρ_k . So $P(f)$ has a Fourier series representation

$$(2.9) \quad P(f) = \frac{1}{2f_M} \sum_{k=-\infty}^{+\infty} \rho_k e^{-i2\pi f k \Delta t}.$$

Substituting (2.9) into (2.7), we get

$$(2.10) \quad I = \int_{-f_M}^{f_M} \ln\left(\frac{1}{2f_M} \sum_{k=-\infty}^{+\infty} \rho_k e^{-i2\pi f k \Delta t}\right) df.$$

Since the ρ_k 's are known for $-M \leq k \leq M$, expression (2.10) will be maximized with respect to the ρ_k 's, with $|k| > M$. Setting the appropriate derivatives to zero gives

$$\begin{aligned}
 (2.11) \quad \frac{\partial I}{\partial \rho_\ell} &= \int_{-f_M}^{f_M} \frac{\frac{1}{2f_M} e^{-i2\pi f \ell \Delta t}}{\left[\frac{1}{2f_M} \sum_{k=-\infty}^{+\infty} \rho_k e^{-i2\pi f k \Delta t} \right]} df \\
 &= \frac{1}{2f_M} \int_{-f_M}^{f_M} \frac{e^{-i2\pi f \ell \Delta t}}{P(f)} df \\
 &= 0, \quad |\ell| > M.
 \end{aligned}$$

But $[P(f)]^{-1}$ has a Fourier expansion so that

$$(2.12) \quad [P(f)]^{-1} = \sum_{m=-\infty}^{+\infty} \lambda_m e^{-i2\pi f m \Delta t}.$$

Thus (2.11) becomes

$$\begin{aligned}
 (2.13) \quad \frac{\partial I}{\partial \rho_\ell} &= \frac{1}{2f_M} \int_{-f_M}^{f_M} \left(\sum_{m=-\infty}^{+\infty} \lambda_m e^{-i2\pi f m \Delta t} \right) e^{-i2\pi f \ell \Delta t} df \\
 &= \frac{1}{2f_M} \int_{-f_M}^{f_M} \sum_{m=-\infty}^{+\infty} \lambda_m (\cos 2\pi f \Delta t m - i \sin 2\pi f \Delta t m) \times \\
 &\quad (\cos 2\pi f \Delta t \ell - i \sin 2\pi f \Delta t \ell) df \\
 &= 0 \quad \text{if} \quad |\ell| > M.
 \end{aligned}$$

Because of the orthogonal properties of sines and cosines, (2.13)

implies that $\lambda_m = 0$ for $|m| > M$.

From (2.12) and the preceding, we get

$$(2.14) \quad P(f) = \frac{1}{\sum_{m=-M}^M \lambda_m e^{-i2\pi f m \Delta t}}.$$

Thus (2.8) becomes

$$(2.15) \quad \rho_k = \int_{-f_M}^{f_M} \frac{e^{i2\pi f k \Delta t}}{\sum_{m=-M}^M \lambda_m e^{-i2\pi f m \Delta t}} df.$$

Let $z = e^{-i2\pi f \Delta t}$. Under this transformation,
 $dz = -i2\pi \Delta t e^{-i2\pi f \Delta t} df$, $df = -dz / i2\pi \Delta t z$ and (2.15) changes to

$$(2.16) \quad \rho_k = \frac{1}{2\pi i \Delta t} \oint \frac{z^{-k-1}}{\sum_{m=-M}^M \lambda_m z^m} dz,$$

where the contour integral around the unit circle $|z| = 1$ is computed in the counterclockwise direction.

Now $\sum_{m=-M}^M \lambda_m z^m$ must be real and positive for $|z| = 1$ since we require $P(f)$ to be real and positive. So

$$(2.17) \quad \begin{aligned} \sum_{m=-M}^M \lambda_m z^m &= \left(\sum_{m=-M}^M \lambda_m z^m \right)^* \\ &= \sum_{m=-M}^M \lambda_m^* z^{*m} = \sum_{m=-M}^M \lambda_m^* (1/z)^m, \end{aligned}$$

where $*$ denotes complex conjugation and $z z^* = 1$.

The expanded form of the left hand side of (2.17) is

$\lambda_0 + \lambda_1 z + \dots + \lambda_M z^M + \lambda_{-1} z^{-1} + \dots + \lambda_{-M} z^{-M}$, whereas the right hand side is $\lambda_0^* + \lambda_1^* z^{-1} + \dots + \lambda_M^* z^{-M} + \lambda_{-1}^* z + \dots + \lambda_{-M}^* z^M$. By equating the coefficients in (2.17), we have the relationship

$$(2.18) \quad \lambda_j = \lambda_{-j}^*, \quad j = 1, 2, \dots, M,$$

and

$$(2.19) \quad \sum_{m=-M}^M \lambda_m z^m = \lambda_0 + \lambda_1 z + \dots + \lambda_M z^M \\ + \lambda_1^* z^{-1} + \dots + \lambda_M^* z^{-M}.$$

By the fundamental theorem of Algebra $\sum_{m=-M}^M \lambda_m z^m$ has $2M$ roots and furthermore is factorable in the form

$$(2.20) \quad \sum_{m=-M}^M \lambda_m z^m = (b_0 + b_1 z + \dots + b_M z^M) \times \\ (b_0 + b_{-1} z^{-1} + \dots + b_{-M} z^{-M}) \\ = (b_0 + b_1 z + \dots + b_M z^M) \times \\ (b_0 + b_1^* z^{-1} + \dots + b_M^* z^{-M}).$$

Factoring b_0 from (2.20) leads to

$$\begin{aligned}
 (2.21) \quad \sum_{m=-M}^M \lambda_m z^m &= b_0^2 (1 + a_1 z + \dots + a_M z^M) \times \\
 &\quad (1 + a_1^* z^{-1} + \dots + a_M^* z^{-M}) \\
 &= (b_0^2 / 2f_M \Delta t) \left(\sum_{m=0}^M a_m z^m \right) \left(\sum_{m=0}^M a_m^* z^{-m} \right),
 \end{aligned}$$

where $b_0 > 0$, $a_0 = 1$, $a_i = b_i / b_0$, $a_i^* = b_i^* / b_0$, $i = 1, 2, \dots, M$ and $2 f_M \Delta t = 1$.

Let $P_M = 2 f_M / b_0^2$. Then $P_M > 0$ and Equation (2.14) becomes

$$\begin{aligned}
 (2.22) \quad P(f) &= \frac{1}{\sum_{m=-M}^M \lambda_m z^m} \\
 &= \frac{P_M \Delta t}{\left(\sum_{m=0}^M a_m z^m \right) \left(\sum_{m=0}^M a_m^* z^{-m} \right)}.
 \end{aligned}$$

The autocorrelation values in (2.16) simplify into

$$(2.23) \quad \rho_k = \frac{P_M}{2\pi i} \oint \frac{z^{-k-1}}{\left(\sum_{m=0}^M a_m z^m \right) \left(\sum_{m=0}^M a_m^* z^{-m} \right)} dz,$$

$-M \leq k \leq M$. Replacing k by $k-j$ with $j \geq 0$ in (2.23) gives

$$(2.24) \quad \rho_{k-j} = \frac{P_M}{2\pi i} \oint \frac{z^{-k+j-1} dz}{\left(\sum_{m=0}^M a_m z^m \right) \left(\sum_{m=0}^M a_m^* z^{-m} \right)}$$

Multiplying (2.24) by a_k^* and forming summations result in

$$(2.25) \quad \sum_{k=0}^M a_k^* \rho_{k-j} = \frac{P_M}{2\pi i} \oint \frac{z^{j-1} \sum_{k=0}^M a_k^* z^{-k}}{\left(\sum_{m=0}^M a_m z^m \right) \left(\sum_{m=0}^M a_m^* z^{-m} \right)} dz$$

$$= \frac{P_M}{2\pi i} \oint \frac{z^{j-1}}{\sum_{m=0}^M a_m z^m} dz, \quad j \geq 0.$$

What needs to be done at this point is to simplify the contour integral on the right hand side of (2.25). To do so requires Cauchy's famous theorems in complex variables theory (see Nehari 1968).

Theorem A: Let $f(z)$ be analytic in a region R and on its boundary C .

Then:

$$(2.26) \quad \oint_C f(z) dz = 0.$$

Theorem B: Cauchy's Integral Formula. If $f(z)$ is analytic inside and on the boundary C of a simply-connected region R and a any point inside C , then

$$(2.27) \quad f(a) = \frac{1}{2\pi i} \oint_C \frac{f(z)}{z-a} dz .$$

For $j = 0$, Equation (2.25) becomes

$$(2.28) \quad \sum_{k=0}^M a_k^* \rho_k = \frac{P_M}{2\pi i} \oint \frac{1}{\sum_{m=0}^M a_m z^m} \frac{dz}{z} .$$

Let $f(z) = 1/\sum_{m=0}^M a_m z^m$. Then $f(z)$ is analytic inside and on the unit circle $|z| = 1$. Thus by Cauchy's integral formula,

$$(2.29) \quad \frac{P_M}{2\pi i} \oint \frac{1}{\sum_{m=0}^M a_m z^m} \frac{dz}{z} = P_M f(0) = P_M ,$$

since $f(0) = a_0 = 1$.

For $j \geq 1$,

$$(2.30) \quad \frac{P_M}{2\pi i} \oint \frac{z^{j-1}}{\sum_{m=0}^M a_m z^m} dz = \frac{P_M}{2\pi i} \oint g(z) dz ,$$

where $g(z) = z^{j-1} / \sum_{m=0}^M a_m z^m$, a function which is analytic inside and on the unit circle $|z| = 1$. By Theorem A, $\oint g(z) dz = 0$. So

$$(2.31) \quad \frac{P_M}{2\pi i} \oint \frac{z^{j-1}}{\sum_{m=0}^M a_m z^m} dz = 0, \quad j \geq 1.$$

By combining (2.25), (2.29) and (2.30), we get

$$(2.32) \quad \sum_{k=0}^M a_k^* \rho_{k-j} = \begin{cases} P_M & \text{if } j = 0, \\ 0 & \text{if } j \geq 1. \end{cases}$$

Or similarly, since $\rho_{k-j}^* = \rho_{j-k}$, (2.32) becomes

$$(2.33) \quad \sum_{k=0}^M \rho_{k-j} a_k = \begin{cases} P_M & \text{if } j = 0, \\ 0 & \text{if } j \geq 1. \end{cases}$$

The matrix equation corresponding to (2.33) is

$$(2.34) \quad \begin{bmatrix} \rho_0 & \rho_1 & \cdots & \cdots & \cdots & \rho_M \\ \rho_{-1} & \rho_0 & \cdots & \cdots & \cdots & \rho_{M-1} \\ \vdots & & & & & \vdots \\ \vdots & & & & & \vdots \\ \vdots & & & & & \vdots \\ \rho_{-M} & \rho_{-M+1} & \cdots & \cdots & \cdots & \rho_0 \end{bmatrix} \begin{bmatrix} 1 \\ a_1 \\ \vdots \\ \vdots \\ \vdots \\ a_M \end{bmatrix} = \begin{bmatrix} P_M \\ 0 \\ \vdots \\ \vdots \\ \vdots \\ 0 \end{bmatrix}$$

The elements a_i , $i = 0, 1, 2, \dots, M$ with $a_0 = 1$ are normally called filter or reflection coefficients and Equation (2.34) is the equation for designing an $M + 1$ term prediction error filter with mean square error P_M . The meaning of these terms will become clearer in the next section. The matrix in (2.34) is a Toeplitz matrix.

Briefly reviewing the results of this section, we see that the power spectrum which maximizes (2.7) subject to (2.8) is the one given in (2.22). We can rewrite (2.22) in the following way:

$$\begin{aligned}
 (2.35) \quad P(f) &= \frac{P_M \Delta t}{\left| 1 + \sum_{m=1}^M a_m z^m \right|^2} \\
 &= \frac{P_M \Delta t}{\left| 1 + \sum_{m=1}^M a_m e^{-i2\pi f m \Delta t} \right|^2}
 \end{aligned}$$

since $z = e^{-i2\pi f \Delta t}$ and $a_0 = 1$. The terms P_M and a_m , $m = 1, \dots, M$ are given by the solution to the matrix equation (2.34). Once these are known, the power spectrum $P(f)$ can be computed for values of the frequency f in the range $-\frac{1}{2\Delta t} \leq f \leq \frac{1}{2\Delta t}$.

E. Predictive Filtering and the Mean Square Error P_M

As in Chapter I, let the given real time series be $x_0, x_1, \dots, x_{n-1}$. Assume further that the mean of the process has been approximated and subtracted from the x_i 's. The mathematical derivations in the preceding sections do not directly involve the x_i 's. Instead, the autocorrelation function values if known provide the basic information

in calculating the power spectrum $P(f)$. However, by working with the x_i 's directly, one can arrive at the same matrix equation (2.34). The idea is to apply the filter coefficients a_i , $i = 1, \dots, M$, with $M < n$ to the time series through a linear prediction relationship.

Let the time series $x_0, x_1, \dots, x_{n-1}$ be written in the form $x_0, x_1, x_2, \dots, x_{t-M}, x_{t-M+1}, \dots, x_{t-1}, x_t, \dots, x_{n-1}$. Consider the expression

$$(2.36) \quad \sum_{s=0}^M a_s x_{t-s} = x_t - (-a_1 x_{t-1} - \dots - a_M x_{t-M}) ,$$

where $a_0 = 1$. Furthermore let

$$(2.37) \quad \hat{x}_t = \sum_{s=1}^M (-a_s) x_{t-s} .$$

Then

$$(2.38) \quad e_t = x_t - \hat{x}_t = \sum_{s=0}^M a_s x_{t-s}$$

is actually the error in predicting x_t from the previous values $x_{t-1}, \dots, x_{t-M}$ by means of the prediction coefficients $a_1, a_2, \dots, a_M$. The error e_t , therefore, has to be kept close to a minimum for a good prediction.

The mean square error is defined to be

$$(2.39) \quad E\{e_t^2\} = E\{(x_t - \hat{x}_t)^2\} = E\left\{\left(\sum_{s=0}^M a_s x_{t-s}\right)^2\right\}$$

where E stands for the expectation operator. Now

$$(2.40) \quad \begin{aligned} E\{e_t^2\} &= E\left\{x_t - \sum_{s=1}^M (-a_s) x_{t-s}\right\}^2 \\ &= E\{x_t^2\} - 2 \sum_{s=1}^M (-a_s) E\{x_t x_{t-s}\} \\ &\quad + E\left\{\left[\sum_{s=1}^M (-a_s) x_{t-s}\right]^2\right\}. \end{aligned}$$

Since the x_i 's are assumed to be given, we would like to minimize the mean square error with respect to a_j , $j = 1, \dots, M$.

$$(2.41) \quad \begin{aligned} \frac{\partial E\{e_t^2\}}{\partial a_j} &= 2E\{x_t x_{t-j}\} + 2 \sum_{s=1}^M (-a_s) E\{x_{t-s} (-x_{t-j})\} \\ &= 0. \end{aligned}$$

By definition, the autocorrelation is

$$(2.42) \quad \rho_t = E\{x_s x_{s-t}\},$$

for real x values. So (2.41) becomes

$$(2.43) \quad 2\rho_j + 2 \sum_{s=1}^M (a_s) \rho_{j-s} = 0 ,$$

or

$$(2.44) \quad \rho_j + \sum_{s=1}^M a_s \rho_{j-s} = 0, \quad j = 1, \dots, M .$$

We thereby have the system of equations

$$(2.45) \quad \begin{array}{rcl} \rho_1 + a_1 \rho_0 + \dots + a_M \rho_{1-M} & = & 0 \\ \rho_2 + a_1 \rho_1 + \dots + a_M \rho_{2-M} & = & 0 \\ \cdot & & \cdot \\ \cdot & & \cdot \\ \cdot & & \cdot \\ \cdot & & \cdot \\ \rho_M + a_1 \rho_{M-1} + \dots + a_M \rho_0 & = & 0 . \end{array}$$

The above system of equations shows that each ρ_j , $j=1, \dots, M$ is a linear combination of the others. To complete the set, one adds the equation

$$(2.46) \quad \rho_0 + a_1 \rho_1 + \dots + a_M \rho_M = E \left\{ e_t x_t \right\}$$

which follows from (2.38) since

$$\begin{aligned}
 (2.47) \quad E\{e_t x_t\} &= E\left\{\sum_{s=0}^M a_s x_{t-s} x_t\right\} \\
 &= \sum_{s=0}^M a_s E\{x_t x_{t-s}\} = \sum_{s=0}^M a_s \rho_s.
 \end{aligned}$$

We can actually show that P_M whose value appears in Equation (2.33) can be expressed as

$$(2.48) \quad P_M = E\{e_t x_t\} = E\{e_t^2\}.$$

That is, P_M is just the least mean square error in the linear prediction. From (2.33) and (2.47), it follows that

$$(2.49) \quad P_M = \sum_{s=0}^M a_s \rho_s = E\{e_t x_t\}.$$

To show that $P_M = E\{e_t^2\}$ we first note that from (2.49) it follows that (2.46) can be rewritten with the quantity P_M on the right hand side and when this equation is appended to the system (2.45) the result is identical to the system (2.34). This follows from the fact that $\rho_{-s} = \rho_s$. Equation (2.34) is pre-multiplied by the row vector $[1 \ a_1 \ a_2 \ \dots \ a_M]$ to get

$$(2.50) \quad [1 \ a_1 \ a_2 \ \dots \ a_M] \begin{bmatrix} \rho_0 & \rho_1 & \dots & \rho_M \\ \rho_{-1} & \rho_0 & \dots & \rho_{M-1} \\ \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot \\ \rho_{-M} & \rho_{-M+1} & \dots & \rho_0 \end{bmatrix} \begin{bmatrix} 1 \\ a_1 \\ \cdot \\ \cdot \\ a_M \end{bmatrix} = P_M,$$

from which it follows that

$$(2.51) \quad \begin{aligned} P_M &= \sum_{v=0}^M \sum_{s=0}^M a_v \rho_{v-s} a_s \\ &= \sum_{v=0}^M \sum_{s=0}^M a_v E\{x_{t-s} x_{t-v}\} a_s = E\left\{\left(\sum_{s=0}^M a_s x_{t-s}\right)^2\right\}. \end{aligned}$$

Applying (2.39) gives

$$(2.52) \quad P_M = E\left\{\left(\sum_{s=0}^M a_s x_{t-s}\right)^2\right\} = E\{e_t^2\}.$$

For a fixed M and set of M filter coefficients, every x_t in the set $\{x_0, x_1, \dots, x_{n-1}\}$ can be predicted linearly as long as one stays within the set of x - values. We can think of this as the process in Diagram 2.1 with $a_0 = 1$. The above process is called filtering. In particular, it is called forward filtering, x_t being predicted from its M previous values. Backward filtering, on the other hand, is described in Diagram 2.2.

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Diagram 2.1. Forward filtering.

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Diagram 2.2. Backward filtering.

In Diagram 2.2, x_{t-M} is predicted from its M future values x_{t-M+1} , ..., x_t . The error in the prediction is h_t and

$$\begin{aligned}
 (2.53) \quad h_t &= \sum_{s=0}^M a_s x_{t-M+s} \\
 &= a_M x_t + a_{M-1} x_{t-1} + \dots + a_1 x_{t-M+1} + x_{t-M} .
 \end{aligned}$$

Let

$$(2.54) \quad \hat{x}_{t-M} = \sum_{s=1}^M (-a_s) x_{t-M+s} .$$

Then

$$\begin{aligned}
 (2.55) \quad E\{h_t^2\} &= E\{(x_{t-M} - \hat{x}_{t-M})^2\} \\
 &= E\left\{\left[x_{t-M} - \sum_{s=1}^M (-a_s) x_{t-M+s}\right]^2\right\}.
 \end{aligned}$$

By taking the partial derivative of $E\{h_t^2\}$ with respect to a_j , $j = 1, \dots, M$ and equating the results to 0, we get

$$\begin{aligned}
 (2.56) \quad \frac{\partial E\{h_t^2\}}{\partial a_j} &= \frac{\partial}{\partial a_j} \left[E\{x_{t-M} x_{t-M}\} - 2 \sum_{s=1}^M (-a_s) E\{x_{t-M+s} x_{t-M}\} \right. \\
 &\quad \left. + E\left\{\left[\sum_{s=1}^M (-a_s) x_{t-M+s}\right]^2\right\} \right] \\
 &= 2E\{x_{t-M} x_{t-M+j}\} + 2 \sum_{s=1}^M (-a_s) E\{(x_{t-M+s})(-x_{t-M+j})\} \\
 &= 0.
 \end{aligned}$$

By applying the definition of the autocorrelation function to (2.56), the following formula is obtained:

$$(2.57) \quad \rho_j + \sum_{s=1}^M a_s \rho_{j-s} = 0, \quad j = 1, \dots, M.$$

Equation (2.57) is exactly Equation (2.44). This implies that forward and backward filtering lead to the same matrix equation and that

$$(2.58) \quad P_M = E\{e_t^2\} = E\{h_t^2\}$$

for all possible t .

The following simple example will help illustrate the ideas described in this section. Suppose the time series is the set of values $A = \{x_0, x_1, x_2, x_3\}$. Let $M = 2$. The set A has two $M + 1$ ordered subsets: $A_1 = \{x_0, x_1, x_2\}$ and $A_2 = \{x_1, x_2, x_3\}$. Denote the filter coefficients by $a_0 = 1, a_1, a_2$. Then both x_0 and x_1 can be predicted from their future values by backward filtering and

$$(2.59) \quad \hat{x}_0 = (-a_1) x_1 + (-a_2) x_2 ,$$

$$\hat{x}_1 = (-a_1) x_2 + (-a_2) x_3 ,$$

where $\hat{x}_0, \hat{x}_1$ are the predicted values of x_0, x_1 respectively.

The backward filtering errors are:

$$(2.60) \quad h_0 = x_0 - \hat{x}_0 = x_0 + a_1 x_1 + a_2 x_2 ,$$

$$h_1 = x_1 - \hat{x}_1 = x_1 + a_1 x_2 + a_2 x_3 .$$

On the other hand, x_2 and x_3 are predicted by forward filtering and

$$\begin{aligned} \hat{x}_2 &= (-a_1) x_1 + (-a_2) x_0, \\ (2.61) \quad \hat{x}_3 &= (-a_1) x_2 + (-a_2) x_1, \end{aligned}$$

where $\hat{x}_2$, $\hat{x}_3$ are the predicted values of x_2 and x_3 respectively.

The forward filtering errors are:

$$\begin{aligned} e_2 &= x_2 - \hat{x}_2 = x_2 + a_1 x_1 + a_2 x_0, \\ (2.62) \quad e_3 &= x_3 - \hat{x}_3 = x_3 + a_1 x_2 + a_2 x_1. \end{aligned}$$

Briefly summarizing the discussion in the preceding pages, in predicting a point in a time series, either from its future or past values, one should minimize the mean square error P_M with respect to the filter coefficients a_j . In the next section, we discuss an algorithm (developed by J. P. Burg, 1975) which determines the values of the a_j 's and P_M in Equation (2.34).

F. The Burg Algorithm for Estimating P_M and the Filter Coefficients

J. P. Burg's algorithm for finding the filter coefficients and P_M in Equation (2.34) depend on both forward and backward filtering, i.e., the points in the time series are predicted from their previous and

future values. The solution to Equation (2.34) is based on a modified version of the Levinson algorithm for solving a Toeplitz matrix equation. Since the algorithm is an iterative procedure, the filter coefficients $(1, a_1, a_2, \dots, a_M)$ will be written as $(1, a_{1M}, a_{2M}, \dots, a_{MM})$.

Starting with a one-point filter ($a_{00} = 1$), we have

$$(2.63) \quad [\rho_0] [1] = [P_0] \quad \text{or} \quad \rho_0 = P_0$$

So the value of P_0 is determined by the autocorrelation ρ_0 . The two-point filter has the matrix equation

$$(2.64) \quad \begin{bmatrix} \rho_0 & \rho_1 \\ \rho_1 & \rho_0 \end{bmatrix} \begin{bmatrix} 1 \\ a_{11} \end{bmatrix} = \begin{bmatrix} P_1 \\ 0 \end{bmatrix}$$

Assuming that a_{11} and P_1 have been determined (the actual computations are described in the next section) the three-point filter is obtained from the preceding equation by means of the following extension:

$$(2.65) \quad \begin{bmatrix} \rho_0 & \rho_1 & \rho_2 \\ \rho_1 & \rho_0 & \rho_1 \\ \rho_2 & \rho_1 & \rho_0 \end{bmatrix} \left\{ \begin{bmatrix} 1 \\ a_{11} \\ 0 \end{bmatrix} + c_1 \begin{bmatrix} 0 \\ a_{11} \\ 1 \end{bmatrix} \right\} = \begin{bmatrix} P_1 \\ 0 \\ \Delta_1 \end{bmatrix} + c_1 \begin{bmatrix} \Delta_1 \\ 0 \\ P_1 \end{bmatrix}$$

where it is required that

$$(2.66) \quad \Delta_1 + c_1 p_1 = 0 \quad .$$

We note that the vectors $c_1 \begin{bmatrix} 0 \\ a_{11} \\ 1 \end{bmatrix}$ and $c_1 \begin{bmatrix} \Delta_1 \\ 0 \\ p_1 \end{bmatrix}$ are just multiples of the reverses of $\begin{bmatrix} 1 \\ a_{11} \\ c \end{bmatrix}$ and $\begin{bmatrix} p_1 \\ 0 \\ \Delta_1 \end{bmatrix}$ respectively. The reason for also applying the filter in reverse is that each point in the time series $x_0, x_1, \dots, x_{n-1}$ is predicted forward and backward. The three-point filter then is $(1, a_{12}, a_{22})$ where

$$(2.67) \quad \begin{aligned} a_{12} &= a_{11} + c_1 a_{11} \quad , \\ a_{22} &= c_1 \quad , \quad p_2 = p_1 + c_1 \Delta_1 \quad . \end{aligned}$$

Extending the matrix equation in (2.65) leads to:

$$(2.68) \quad \begin{bmatrix} \rho_0 & \rho_1 & \rho_2 & \rho_3 \\ \rho_1 & \rho_0 & \rho_1 & \rho_2 \\ \rho_2 & \rho_1 & \rho_0 & \rho_1 \\ \rho_3 & \rho_2 & \rho_1 & \rho_0 \end{bmatrix} \left\{ \begin{bmatrix} 1 \\ a_{12} \\ a_{22} \\ 0 \end{bmatrix} + c_2 \begin{bmatrix} 0 \\ a_{22} \\ a_{12} \\ 1 \end{bmatrix} \right\} = \begin{bmatrix} p_2 \\ 0 \\ 0 \\ \Delta_2 \end{bmatrix} + c_2 \begin{bmatrix} \Delta_2 \\ 0 \\ 0 \\ p_2 \end{bmatrix}$$

where in this case the four-point filter is $(1, a_{13}, a_{23}, a_{33})$ with

$$(2.69) \quad \begin{aligned} a_{13} &= a_{12} + c_2 a_{22} , \\ a_{23} &= a_{22} + c_2 a_{12} , \quad a_{33} = c_2 . \end{aligned}$$

Again it is required that $\Delta_2 + c_2 P_2 = 0$.

Continuing in this manner, we get the general form:

$$(2.70) \quad \begin{bmatrix} c_0 & c_1 & \dots & c_M \\ c_1 & c_0 & \dots & c_{M-1} \\ c_2 & c_1 & \dots & c_{M-2} \\ \vdots & \vdots & \dots & \vdots \\ \vdots & \vdots & \dots & \vdots \\ c_M & c_{M-1} & \dots & c_0 \end{bmatrix} \begin{bmatrix} 1 \\ a_{1,M} \\ a_{2,M} \\ \vdots \\ \vdots \\ a_{M,M} \end{bmatrix} = \begin{bmatrix} p_M \\ 0 \\ 0 \\ \vdots \\ \vdots \\ 0 \end{bmatrix}$$

valid for any $M < n$.

Furthermore, for $i = 1, \dots, M-1$,

$$(2.71) \quad \begin{aligned} a_{i,M} &= a_{i,M-1} + c_{M-1} a_{M-i,M-1} , \\ a_{M,M} &= c_{M-1} , \quad p_M = p_{M-1} + c_{M-1} \Delta_{M-1} , \\ \Delta_{M-1} + c_{M-1} p_{M-1} &= 0 . \end{aligned}$$

Now since $\Delta_{M-1} = -c_{M-1} P_{M-1}$, then

$$\begin{aligned}
 P_M &= P_{M-1} + c_{M-1} \Delta_{M-1} \\
 (2.72) \quad &= P_{M-1} - c_{M-1}^2 P_{M-1} \\
 &= P_{M-1} (1 - c_{M-1}^2) \\
 &= P_{M-1} (1 - a_{MM}^2) .
 \end{aligned}$$

The least mean square error P_M can only be estimated in this algorithm. Let us go back to the time series $\{x_0, x_1, \dots, x_{n-1}\}$ and assume that there are L possible subsets of the form $\{x_{t_j-M}, x_{t_j-M+1}, \dots, x_{t_j}\}$, $j = 1, \dots, L$. By applying the filter $(1, a_{1,M}, \dots, a_{M,M})$ in a forward manner to each of the x_{t_j} , $j = 1, \dots, L$, we get

$$\begin{aligned}
 &x_{t_j} + a_{1,M} x_{t_j-1} + \dots + a_{M,M} x_{t_j-M} \\
 &= x_{t_j} + (a_{1,M-1} + c_{M-1} a_{M-1,M-1}) x_{t_j-1} \\
 (2.73) \quad &+ (a_{2,M-1} + c_{M-1} a_{M-2,M-1}) x_{t_j-2} \\
 &+ \dots + c_{M-1} x_{t_j-M} , \\
 &j = 1, \dots, L .
 \end{aligned}$$

Equation (2.73) gives the error in predicting x_{t_j} from its previous M values. Since there are L equations of the form (2.73), the mean square error P_M can be estimated by the squared averages

$$\begin{aligned}
 P_{1M} = \sum_{j=1}^L \gamma_j & \left[x_{t_j} + (a_{1,M-1} + c_{M-1} a_{M-1,M-1}) x_{t_{j-1}} \right. \\
 (2.74) \quad & + \dots + (a_{M-1,M-1} + c_{M-1} a_{1,M-1}) x_{t_{j-M+1}} \\
 & \left. + (c_{M-1} x_{t_{j-M}}) \right]^2,
 \end{aligned}$$

where the γ_j 's are weights such that $\sum_{j=1}^L \gamma_j = 1$ and $\gamma_j \geq 0$.

The terms in (2.74) can be re-grouped using the quantities

$$\begin{aligned}
 \alpha_j &= a_{M-1,M-1} x_{t_{j-M+1}} + \dots + a_{1,M-1} x_{t_{j-1}} + x_{t_j}, \\
 (2.75) \quad \beta_j &= x_{t_{j-M}} + a_{1,M-1} x_{t_{j-M+1}} + \dots + a_{M-1,M-1} x_{t_{j-1}}, \\
 j &= 1, \dots, L.
 \end{aligned}$$

which are seen to be a forward prediction error and a backward prediction error respectively. The new expression is then

$$(2.76) \quad P_{1M} = \sum_{j=1}^L \gamma_j (\alpha_j + c_{M-1} \beta_j)^2.$$

If on the other hand, the filter $(1 \ a_{1M} \ \dots \ a_{MM})$ were applied backward, then

$$\begin{aligned}
 & x_{t_j-M} + a_{1,M} x_{t_j-M+1} + a_{2,M} x_{t_j-M+2} + \dots + a_{M,M} x_{t_j} \\
 &= x_{t_j-M} + (a_{1,M-1} + c_{M-1} a_{M-1,M-1}) x_{t_j-M+1} \\
 &+ (a_{2,M-1} + c_{M-1} a_{M-2,M-1}) x_{t_j-M+2} \\
 &+ \dots + (a_{M-1,M-1} + c_{M-1} a_{1,M-1}) x_{t_j-1} \\
 &+ c_{M-1} x_{t_j} .
 \end{aligned}
 \tag{2.77}$$

As was noted earlier, the backward and forward filters should produce the same mean square error P_M . So P_M can also be estimated by

$$\begin{aligned}
 P_{2M} &= \sum_{j=1}^L \gamma_j \left[x_{t_j-M} + (a_{1,M-1} + c_{M-1} a_{M-1,M-1}) x_{t_j-M+1} \right. \\
 &\quad + (a_{2,M-1} + c_{M-1} a_{M-2,M-1}) x_{t_j-M+2} \\
 &\quad \left. + \dots + c_{M-1} x_{t_j} \right] \\
 &= \sum_{j=1}^L \gamma_j (\beta_j + c_{M-1} \alpha_j)^2 .
 \end{aligned}
 \tag{2.78}$$

Thus one class of estimates for P_M is

$$(2.79) \quad \hat{P}_M = \frac{1}{2}(P_{1M} + P_{2M})$$

$$= \frac{1}{2} \sum_{j=1}^L \gamma_j \left[(\alpha_j + c_{M-1} \beta_j)^2 + (\beta_j + c_{M-1} \alpha_j)^2 \right]$$

where it is required that the γ_j be specified.

G. Andersen's Version of the Burg Algorithm for Estimating Coefficients

In applying the above algorithm, one would start with an estimate of P_0 . Then at each stage M , the quantity $\partial \hat{P}_M / \partial a_{M1}$ is set equal to zero. For $M = 0$, P_0 is estimated by

$$\hat{P}_0 = r_0$$

$$(2.80) \quad = \left(\frac{1}{n-1} \right) \sum_{j=0}^{n-1} x_j x_j .$$

where r_0 is an estimate of the autocorrelation r_0 .

For $M = 1$, the set $\{x_0, x_1, \dots, x_{n-1}\}$ would have $n-1$ subsets: $\{x_0, x_1\}$, $\{x_1, x_2\}$, $\dots$, $\{x_{n-2}, x_{n-1}\}$ and the average power is:

$$(2.81) \quad \hat{P}_1 = \frac{1}{2} \frac{1}{n-1} \sum_{j=0}^{n-2} \left[(x_j + a_{11} x_{j+1})^2 + (x_{j+1} + a_{11} x_j)^2 \right] ,$$

where the weights are $\gamma_j = \frac{1}{n-1}$, $j = 0, \dots, n-2$ (see N. Andersen, 1974). Equating the partial of $\hat{P}_1$ with respect to a_{11} to zero, we get

$$(2.82) \quad \frac{\partial \hat{P}_1}{\partial a_{11}} = \frac{1}{2(n-1)} \sum_{j=0}^{n-2} \left[2(x_j + a_{11} x_{j+1}) x_{j+1} + 2(x_{j+1} + a_{11} x_j) x_j \right] = 0 .$$

So

$$(2.83) \quad a_{11} = -2 \frac{\sum_{j=0}^{n-2} x_j x_{j+1}}{\sum_{j=0}^{n-2} (x_j^2 + x_{j+1}^2)} .$$

Since a_{11} is known at this point, one solves for $\hat{P}_1$ by using (2.72),

$$(2.84) \quad \hat{P}_1 = \hat{P}_0 (1 - a_{11}^2) = r_0 (1 - a_{11}^2) .$$

For $M=2$, the set $\{x_0, x_1, \dots, x_{n-1}\}$ has $n-2$ subsets with $M+1=3$ consecutive elements in each subset: $\{x_0, x_1, x_2\}$, $\{x_1, x_2, x_3\}$, $\dots$, $\{x_{n-3}, x_{n-2}, x_{n-1}\}$. The resulting average power is:

$$\begin{aligned}
(2.85) \quad \hat{p}_2 &= \frac{1}{2} \frac{1}{n-2} \sum_{j=0}^{n-3} \left[(x_j + a_{12} x_{j+1} + a_{22} x_{j+2})^2 \right. \\
&\quad \left. + (x_{j+2} + a_{12} x_{j+1} + a_{22} x_j)^2 \right] \\
&= \frac{1}{2} \frac{1}{n-2} \sum_{j=0}^{n-3} \left[x_j + (a_{11} + a_{22} a_{11}) x_{j+1} \right. \\
&\quad \left. + a_{22} x_{j+2} \right]^2 + \left[x_{j+2} + (a_{11} + a_{22} a_{11}) x_{j+1} \right. \\
&\quad \left. + a_{22} x_j \right]^2,
\end{aligned}$$

since $a_{12} = a_{11} + a_{22} a_{11}$ from (2.67). Thus

$$\begin{aligned}
(2.86) \quad \frac{\partial \hat{p}_2}{\partial a_{22}} &= \frac{1}{2} \frac{1}{n-2} \sum_{j=0}^{n-3} \left[2(x_j + a_{11} x_{j+1} + a_{22} a_{11} x_{j+1} \right. \\
&\quad \left. + a_{22} x_{j+2}) (a_{11} x_{j+1} + x_{j+2}) + 2(x_{j+2} \right. \\
&\quad \left. + a_{11} x_{j+1} + a_{22} a_{11} x_{j+1} + a_{22} x_j) (a_{11} x_{j+1} + x_j) \right] \\
&= 0.
\end{aligned}$$

The value for a_{22} is the following form:

$$\begin{aligned}
 a_{22} = & -2 \sum_{j=0}^{n-3} (a_{11} x_j x_{j+1} + a_{11}^2 x_{j+1}^2 + a_{11} x_{j+1} x_{j+2} \\
 (2.87) \quad & + x_j x_{j+2}) / \sum_{j=0}^{n-3} (x_j^2 + 2a_{11} x_j x_{j+1} + 2a_{11}^2 x_{j+1}^2 \\
 & + 2a_{11} x_{j+1} x_{j+2} + x_{j+2}^2) .
 \end{aligned}$$

Since a_{22} is known, $\hat{P}_2$ is given by the expression:

$$(2.88) \quad \hat{P}_2 = \hat{P}_1 (1 - a_{22}^2) .$$

Continuing in this manner, at the M^{th} stage, there will be $n-M$ subsets of $\{x_0, x_1, \dots, x_{n-1}\}$ and they are of the form: $\{x_0, x_1, \dots, x_M\}, \{x_1, x_2, \dots, x_{M+1}\}, \dots, \{x_{n-M-1}, x_{n-M}, \dots, x_{n-1}\}$. The average output power is then

$$\begin{aligned}
 P_M = & \frac{1}{2} \frac{1}{n-M} \sum_{j=0}^{(n-1)-M} \left[(x_j + \sum_{k=1}^M a_{k,M} x_{j+k})^2 \right. \\
 (2.89) \quad & \left. + (x_{j+M} + \sum_{k=1}^M a_{k,M} x_{j+M-k})^2 \right] \\
 = & \frac{1}{2} \frac{1}{n-M} \sum_{j=0}^{(n-1)-M} \left[\left(\sum_{k=0}^M a_{k,M} x_{j+k} \right)^2 \right. \\
 & \left. + \left(\sum_{k=0}^M a_{k,M} x_{j+M-k} \right)^2 \right] .
 \end{aligned}$$

where $a_{0,M} = 1$. From (2.71), we recall that $a_{i,M} = a_{i,M-1} + a_{MM} a_{M-i,M-1}$, $i = 1, 2, \dots, M-1$. If we impose the added condition that $a_{0,M} = 1$ and $a_{i,M} = 0$ for $i > M$, then (2.71) holds for all M and (2.89) simplifies to:

$$\begin{aligned}
 (2.90) \quad \hat{P}_M &= \frac{1}{2} \frac{1}{n-M} \sum_{j=0}^{(n-1)-M} \left\{ \left[\sum_{k=0}^M (a_{k,M-1} + a_{MM} a_{M-k,M-1}) x_{j+k} \right]^2 \right. \\
 &\quad \left. + \left[\sum_{k=0}^M (a_{k,M-1} + a_{MM} a_{M-k,M-1}) x_{j+M-k} \right]^2 \right\} \\
 &= \frac{1}{2} \frac{1}{n-M} \sum_{j=0}^{(n-1)-M} \left[\left(\sum_{k=0}^M a_{k,M-1} x_{j+k} + a_{MM} \sum_{k=0}^M a_{M-k,M-1} x_{j+k} \right)^2 \right. \\
 &\quad \left. + \left(\sum_{k=0}^M a_{k,M-1} x_{j+M-k} + a_{MM} \sum_{k=0}^M a_{M-k,M-1} x_{j+M-k} \right)^2 \right].
 \end{aligned}$$

Let

$$(2.91) \quad p_{jM} = \sum_{k=0}^M a_{k,M-1} x_{j+k} = \sum_{k=0}^M a_{M-k,M-1} x_{j+M-k},$$

$$q_{jM} = \sum_{j=0}^M a_{M-k,M-1} x_{j+k} = \sum_{k=0}^M a_{k,M-1} x_{j+M-k},$$

$$j = 0, 1, \dots, (n-1) - M.$$

Then

$$(2.92) \quad \hat{p}_M = \frac{1}{2} \frac{1}{n-M} \sum_{j=0}^{(n-1)-M} \left[(p_{jM} + a_{MM} q_{jM})^2 + (q_{jM} + a_{MM} p_{jM})^2 \right] .$$

By setting $\partial \hat{p}_M / \partial a_{MM} = 0$, we get the value for a_{MM} :

$$(2.93) \quad a_{MM} = -2 \frac{\sum_{j=0}^{(n-1)-M} p_{jM} q_{jM}}{\sum_{j=0}^{(n-1)-M} (p_{jM}^2 + q_{jM}^2)} .$$

The recursion formulas for the p_{jm} and q_{jm} are derived from (2.71) and (2.91) They are:

$$(2.94) \quad p_{jM} = p_{j,M-1} + a_{M-1,M-1} q_{j,M-1} ,$$

$$q_{jM} = q_{j+1,M-1} + a_{M-1,M-1} p_{j+1,M-1} .$$

The starting values for the equations in (2.94) for $M=1$ are:

$$(2.95) \quad p_{j1} = x_j , \quad q_{j1} = x_{j+1} , \quad j = 0, 1, \dots, n-2 .$$

The equations in (2.94) and (2.95) are implemented in the computer program which is discussed later in this report.

CHAPTER III

APPLICATIONS OF MESA AND FOURIER ANALYSIS

A. Comparative Study of MESA and Fourier Analysis.

EXAMPLE 3.1.

To study the effectiveness of the two methods, consider a mathematical test case consisting of sine functions. The equation is:

$$x_t = 3 \sin (.05(2\pi t)) + 10 \sin (0.1(2\pi t)) + 5 \sin (0.2(2\pi t)) + 8 \sin (0.21(2\pi t)),$$

$$t = 0, 1, \dots, 99.$$

No trend in the time series is removed in the analysis. The discretized function x_t and its spectra are shown in Figures 3.1A, 3.1B, and 3.1C. By examining the equation for x_t , one can tell that the true periods are at the frequencies: $f = 0.05$, $f = 0.1$, $f = 0.2$, $f = 0.21$, corresponding to periods: $P = 20$, $P = 10$, $P = 5$ and $P = 4.76$ respectively. Table 3.1 gives a listing of the periods as predicted by both forms of analysis.

As one can infer from Figures 3.1B, 3.1C and Table 3.1, MESA gives somewhat better predictions of the frequencies of the periodic components. The number of filter coefficients M used in MESA for this particular problem is 51.

Fig. 3.1. A. Mathematical test case of four sinusoids (100 data points).
 B. Fourier spectrum.
 C. Maximum entropy spectrum with $M = 51$.

TABLE 3.1.
FREQUENCY COMPONENTS OF THE SINE PROBLEM
OF EXAMPLE 3.1 (100 data points)

Frequencies at which true periods occur		Frequencies predicted by MESA - 51 filter coefficients used	Frequencies predicted by Fourier analysis
(1)	0.05	0.0500	0.0475
(2)	0.10	0.1000	0.1000
(3)	0.20	0.2000	0.1975
(4)	0.21	0.2125	0.2125

EXAMPLE 3.2

What happens in the spectrum of the time series x_t in Example 3.1 if the number of sampled points is reduced? Suppose only 20 points are used, say at $t = 0, 1, \dots, 19$. The shorter time series x_t is shown in Figure 3.2A and its spectra in Figures 3.2 B and 3.2 C. Table 3.2 lists the frequencies for both spectra.

TABLE 3.2.
FREQUENCY COMPONENTS OF THE SINE PROBLEM OF
EXAMPLE 3.1 (20 data points)

Frequencies at which true periods occur		Frequencies predicted by MESA - 13 filter coefficients used	Frequencies predicted by Fourier analysis
(1)	0.05	0.0475	0.0300
(2)	0.10	0.0975	0.0925
(3)	0.20	0.1925	
(4)	0.21	0.2075	0.2075

Fig. 3.2. A. Mathematical test case of four sinusoids (20 data points).
 B. Fourier spectrum.
 C. Maximum entropy spectrum with $M = 13$.

Since the number of data points is relatively small, one certainly can expect shifts in the spectral peaks. Such shifts are present in both Figures 3.2B and 3.2C, but the frequencies predicted by MESA are consistently as accurate or more accurate than those predicted by Fourier analysis. The true peak at $f = 0.2$ is missing in the Fourier spectrum, but is approximated by $f = 0.1925$ in the MESA spectrum. The fact that MESA is able to distinguish between two very close frequencies indicates one strong advantage this method has over Fourier analysis. One should also take note that the longest period $P = 20$ is approximated by $P = 21.053$ by MESA, but is given as $P = 33.33$ by the Fourier spectrum.

So far the applications we have been considering involve mathematical test cases, where the periods of the functions are known in advance. In the following three examples, we present real time series.

EXAMPLE 3.3

Figure 3.3A gives a plot of the MacKenzie River region lynx fur returns of the Hudson Bay Company from 1821-1934 (Elton and Nicholson 1942), a period covering 114 years. It is evident that a definite pattern exists. In fact, as one closely examines the points, a period of roughly 10 years emerges. Indeed, both Fourier and MESA spectra bring out this dominant period as can be seen in Figures 3.3B and 3.3C. Both lynx spectra agree quite closely in their estimates of the period of the major peak and of the two more important minor peaks, corresponding to periods of roughly 38 and 11 years. The cycle of about 38 years seems

Fig. 3.3. A. Lynx fur returns (1821 - 1934), (Hudson Bay Company data; Straight line represents the linear trend).
 B. Fourier spectrum.
 C. Maximum entropy spectrum with $M = 69$.

to be a harmonic of 9.66 years, the major period, for $9.66 \times 4 = 38.64$ years.

TABLE 3.3
FREQUENCY COMPONENTS OF LYNX DATA

Major frequencies predicted by MESA - 69 filter coefficients used	Corresponding periods	Major frequencies predicted by Fourier analysis	Corresponding periods
(1) 0.0258	38.55	0.0258	38.65
(2) 0.0959	10.42	0.0918	10.89
(3) 0.1035	9.66	0.1035	9.66
(4) 0.1310	7.63	0.1327	7.54
(5) 0.2086	4.79	0.2078	4.81

EXAMPLE 3.4

Another well-studied time series is the Zurich sunspot record from 1700 to 1976 (see Figure 3.4A) as given by M. Waldemeier. The sunspot time series shows a strong periodicity of 11 years. There also seems to be a period of about 10¹⁰ years if one divides the interval 1700-1976 into roughly 1700-1800, 1800-1900, 1900-1976. Table 3.4 gives a listing of the more dominant peaks in the spectra (Figures 3.4B and 3.4C).

The 99.83 year - period appears to be a harmonic of 9.983, since $9.983 \times 10 = 99.83$. Several papers have been written [see for example (Currie 1972), (Cole 1973)] in an attempt to explain the different periods (other than those appearing in Table 3.4) in the sunspot spectrum. Some authors have claimed that certain periods are the results of beat

Fig. 3.4. A. Waldmeier Sunspot data (1700 - 1976) (Straight line represents the linear trend).
 B. Fourier spectrum.
 C. Maximum entropy spectrum with $M = 160$.

TABLE 3.4
FREQUENCY COMPONENTS OF SUNSPOT DATA

Major frequencies predicted by MESA - 160 filter coefficients used	Corresponding periods	Major frequencies predicted by Fourier analysis	Corresponding periods
(1) 0.01001	99.83	0.01001	99.83
(2) 0.0175	57.05	0.0175	57.05
(3) 0.0834	11.98	0.0826	12.10
(4) 0.0909	10.99	0.0901	11.09
(5) 0.0943	10.60	0.0952	10.51
(6) 0.1001	9.98	0.1001	9.98

frequencies (Wolff 1975). It is commonly accepted that the cycle of about 11 years is a real one.

In the sunspot and lynx records, the number of available data points is more than 10 times the length of the most dominant periods (the MESA estimates of which are 9.66 years in the lynx and 10.99 in the sunspots). Both MESA and Fourier analyses are in agreement with regard to the frequency estimates in the lynx spectrum and the sunspot spectrum. The performance of Fourier analysis is comparable to that of MESA in cases where there are long records available for the time series under study.

EXAMPLE 3.5

We now consider a shorter time series consisting of 42 data points - the annual commercial landings of the striped bass from the Middle Atlantic region in the years 1930 - 1971 (Van Winkle et al., in press) (see Figure 3.5A). A summary of the frequency components is listed in Table 3.5.

TABLE 3.5
FREQUENCY COMPONENTS OF MIDDLE ATLANTIC STRIPED BASS COMMERCIAL LANDINGS

Major frequencies predicted by MESA - 26 filter coefficients used	Corresponding periods	Major frequencies predicted by Fourier analysis	Corresponding periods
(1) 0.0250	39.93		
(2) 0.0550	18.15	0.0509	19.64
(3) 0.0793	12.61		
(4) 0.1302	7.68	0.1310	7.63
(5) 0.2671	3.74	0.2621	3.82

The spectra in Figures 3.5B and 3.5C differ in the frequency interval (0.00 - 0.10). In the Fourier spectrum, only one peak at $f = 0.0509$ appears within the interval. A lot of unresolved power is also present. In the same interval the MESA spectrum has three peaks at $f = 0.0250$, $f = 0.0550$, and $f = 0.0793$. The frequency $f = 0.2621$ with corresponding period $P = 3.815$ in Figure 3.5B [denoted peak (5)] appears to be a harmonic of $f = 0.1310$, since $2 \times 0.1310 = 0.2620$.

Fig. 3.5. A. Middle Atlantic striped bass yearly landings (1930 - 1971) (Straight line represents the linear trend).
 B. Fourier spectrum.
 C. Maximum entropy spectrum with $M = 26$.

B. The Effect of Noise

To compare the performance of the two methods when there is noise in the time series, we consider the sine function of Example 3.1, but utilize only the first 20 points (from $t = 0$ to $t = 19$). There are two major reasons for doing so. First of all, the longest period of the sine function is at frequency $f = 0.05$, corresponding to a period of $P = 20$. In Fourier analysis, the longest period predicted is the length of the time series. Secondly, we would like to see how well MESA and Fourier analysis perform whenever the length of available data is relatively short. We take note that when the number of points is 100, the predictions given by both methods are in fairly good agreement (see Table 3.1). When the number of points decreases to 20, the difference between the MESA spectrum and the Fourier spectrum can easily be detected (see Table 3.2). The equation under study here is:

$$x_t = 3 \sin(0.05(2\pi t)) + 10 \sin(0.1(2\pi t)) + \\ 5 \sin(0.2(2\pi t)) + 8 \sin(0.25(2\pi t)) + \\ R[t; N(0, 6.5 \sigma)],$$

$t = 0, 1, \dots, 19$, where

$R[t; N(0, 6.5 \sigma)]$ = a random number drawn for each time t from a normal distribution $N(0, 6.5 \sigma)$;

$N(0, 6.5 \sigma)$ = the normal distribution with mean 0 and standard deviation 6.5σ ;

6.5 = $\frac{1}{4} (3 + 10 + 5 + 8)$, the average of the amplitudes;

σ = the noise level on a scale of 0.0 to 1.0.

TABLE 3.6

FOURIER ANALYSIS: FREQUENCY COMPONENTS OF EXAMPLE 3.1
(20 data points at different noise levels)

Noise Level (α)	Frequencies (f) of the three highest peaks		
	0.05	0.10	0.2100
0.00	0.0300	0.0925	0.2075
0.05	0.0300	0.0925	0.2075
0.10	0.0300	0.0925	0.2075
0.20	0.0275	0.0925	0.2075
0.30	0.0275	0.0900	0.2075

Table 3.6 gives the estimated Fourier frequencies for various noise levels. The peak at $f = 0.05$ is approximated by $f = 0.03$ when $\alpha = 0.0, 0.05$ and 0.10 . As α increases to 0.20 and 0.30 , the peak shifts even farther away from the true one. In the sine function under study, the largest amplitude is 10 and it belongs to the sinusoid with period at $f = 0.10$. This sinusoid is least affected by the increase in noise. The frequency estimate of $f = 0.10$ is $f = 0.0925$ for the first four levels of α and only drops to $f = 0.09$ when α increases to 0.30 . The third peak of the sine function is at $f = 0.2000$ and does not appear in Table 3.6 at any noise level. This is because the true peak at $f = 0.2100$ is very close to the true peak $f = 0.2000$. As a result, there is a lot of unresolved power at $f = 0.2075$, at all levels of α in Table 3.6 (see Figure 3.2B for $\alpha = 0.00$).

In Table 3.7, the results from MESA are given at different values of the number of filter coefficients M . For $M = 3$ and $\alpha = 0.0, 0.05, 0.10, 0.20, 0.30$, one obtains a very poor spectral representation indicating only one of the four frequencies. As M takes the values 4, 5, 6, 7, at least two of the periods are resolved at $\alpha = 0.0, 0.05$ and 0.10. As the level of noise goes up, the frequencies undergo greater shifts for these values of M . Moreover, when $\alpha = 0.30$, spurious peaks appear at $f = 0.4100$ (for $M = 6$) and $f = 0.3525$ (for $M = 7$). Using $M = 8, 9, 10, 11$ results in a spectrum with three of the four frequencies of x_t at low levels of noise. As M takes the values 12 - 19 and when the noise level stays relatively low, all four frequencies are estimated but only with the addition of spurious peaks as well.

We notice that using too small a value of M results in low resolution, and using too high a value of M results not only in high resolution but also in the appearance of several spurious peaks. Furthermore, for both high levels of noise and large values of M , the number of spurious peaks tends to be high.

Relatively good spectra are obtained for values of M in the range 8 - 14. Although spurious peaks do appear, MESA still gives better estimates than Fourier analysis. Also the problem of sidebands in Fourier analysis is absent from the MESA spectrum (see Figures 3.1B, 3.1C, 3.2B and 3.2C).

Figures 3.6A, 3.6B and 3.6C illustrate the MESA spectrum of x_t at $\alpha = 0.0$, $\alpha = 0.10$, and $\alpha = 0.30$ respectively. The value of M for these plots is 10. Because of the linear plot scaling, the peaks at

TABLE 3.7

MESA: FREQUENCY COMPONENTS OF EXAMPLE 3.1
(20 data points at different noise levels)

Number of filter coefficients (M)	Noise Level (%)				
	0%	5%	10%	20%	30%
3	0.1925	0.1900	0.1900	0.1900	0.1775
4	0.0850	0.0850	0.0850	0.1950	
	0.2050	0.2050	0.2075		
5	0.0925	0.0900	0.0825	0.0875	0.1000
	0.2075	0.2075	0.2050	0.2075	0.2075
6	0.0550				
	0.0975	0.0900	0.0875	0.0825	0.0850
	0.2075	0.2075	0.2050	0.2050	0.2050
					0.4100*
7	0.0950	0.0925	0.0900	0.0850	0.0825
	0.2075	0.2075	0.2050	0.2050	0.2025
					0.3525*
8	0.0475	0.0550			
	0.0950	0.0950	0.0900	0.0900	0.0850
	0.2075	0.2075	0.2050	0.2050	0.2025

*Denotes frequencies at which spurious peaks occur.

TABLE 3.7: Cont'd.

Number of filter coefficients (M)	Noise Level (α)				
	0%	5%	10%	20%	30%
9	0.0525	0.0525			
	0.0950	0.0950	0.0950	0.0900	0.0900
	0.2075	0.2075	0.2050	0.2050	0.2025
					0.3275*
					0.4475*
10	0.0475	0.0500	0.0500		
	0.0950	0.0950	0.0950	0.0925	0.0900
	0.2075	0.2075	0.2050	0.2050	0.2025
					0.3125*
					0.4025*
11	0.0500	0.0475	0.0500	0.0525	
	0.0950	0.0950	0.0950	0.0950	0.0925
	0.2075	0.2075	0.2050	0.2050	0.2025
				0.3100*	0.3075*
				0.4050*	0.4025*
12	0.0425	0.0475	0.0575		
	0.0700*				
	0.0950	0.0950	0.0975	0.0975	0.0925

*Denotes frequencies at which spurious peaks occur.

TABLE 3.7: Cont'd.

Number of filter coefficients (M)	Noise Level (α)				
	0%	5%	10%	20%	30%
	0.1875*				
	0.2075	0.2075	0.2050	0.2050	0.2025
	0.2200*				
			0.3125*	0.3125*	0.3075*
			0.4100*	0.4075*	0.4025*
13	0.0475	0.0475	0.0475		
	0.0975	0.0950	0.0950	0.0975	0.0900
	0.1925				
	0.2075	0.2075	0.2050	0.2050	0.2025
	0.2200*				
			0.3150*	0.3100*	0.3050*
			0.4125*	0.4000*	0.3950*
14	0.0450	0.0500	0.0500		
	0.0950	0.0950	0.0950	0.0950	0.0925
	0.1925				
	0.2075	0.2075	0.2050	0.2050	0.2025
			0.3175*	0.3100*	0.3050*
			0.4125*	0.4000*	0.3950*

*Denotes frequencies at which spurious peaks occur.

TABLE 3.7: Cont'd.

Number of filter coefficients (M)	Noise Level (α)				
	0%	5%	10%	20%	30%
15	0.0475	0.0450	0.0525		
	0.0800*				
	0.0975	0.0975	0.0950	0.0950	0.0975
		0.1875			
	0.2075	0.2075	0.2050	0.2050	0.2025
	0.2225*				
			0.3175*	0.3125*	0.3075*
			0.4125*	0.4000*	0.3950*
16					0.0350*
	0.0500	0.0425	0.0525		
	0.0900*				
	0.0975	0.0975	0.0950	0.0950	0.0875
					0.1225*
	0.1875	0.1875			
	0.2075	0.2075	0.2050	0.2050	0.2025
					0.2675*
			0.3150*	0.3125*	0.3125*
			0.4125*	0.4000*	0.3975*
				0.4750*	0.4750*

*Denotes frequencies at which spurious peaks occur.

TABLE 3.7: Cont'd.

Number of filter coefficients (M)	Noise Level (α)				
	0%	5%	10%	20%	30%
17	0.0500	0.0425	0.0575	0.0550	0.0475
		0.0850*			
	0.0975	0.0975	0.0950	0.0950	0.0900
		0.1850	0.1600		
	0.2075	0.2075	0.2050	0.2050	0.2025
					0.2725*
			0.3150*	0.3150*	0.3150*
			0.4125*	0.4000*	0.3975*
				0.4750*	0.4750*
18	0.0325*		0.0375*		
	0.0500	0.0425	0.0650	0.0550	0.0425
	0.0875*	0.0825*			0.0900*
	0.0975	0.0975	0.0975	0.0975	0.0950
					0.1300*
		0.1800	0.1600		
	0.2075	0.2075	0.2050	0.2050	0.2025
	0.2225*				
				0.2675*	0.2725*
		0.3250*	0.3100*	0.3150*	0.3150*

*Denotes frequencies at which spurious peaks occur.

TABLE 3.7: Cont'd.

Number of filter coefficients (M)	Noise Level (α)				
	0%	5%	10%	20%	30%
			0.3350*	0.4000*	0.3975*
		0.4150*	0.4125*	0.4750*	0.4750*
19	0.0325*		0.0375*	0.0300*	0.0300*
	0.0500	0.0425			
	0.0875*	0.0825*	0.0650*	0.0625*	0.0800*
	0.0975	0.0975	0.0975	0.0950	0.1000
				0.1350*	0.1425*
	0.1800	0.1800	0.1600		
	0.2025*				0.2025*
	0.2075	0.2075	0.2050	0.2050	
	0.2250*				
				0.2700*	0.2725*
		0.3250*	0.3100*	0.3150*	
		0.4125*	0.4125*	0.4725*	0.4725*

*Denotes frequencies at which spurious peaks occur.

Fig. 3.6. Maximum entropy spectrum of four sinusoids (20 points).

- A. $M = 10, \alpha = 0.0.$
- B. $M = 10, \alpha = 0.10.$
- C. $M = 10, \alpha = 0.30.$

0.3075 and 0.4025 in Figure 3.6C cannot be seen, but do show up when a logarithmic scale is used.

The preceding studies gives some indications of the advantages of MESA over Fourier analysis. Below are the major points to be considered in choosing between the two methods.

Fourier Analysis	MESA
1. Assumes a zero extention of the data and uses data windows.	1. Makes no assumptions outside the given data and no data windows are used.
2. The longest period predicted is the length of available data.	2. Can predict periods longer than the length of available data.
3. Sideband effects are produced resulting from power leakage.	3. No sidebands are produced.
4. Frequency components very close to each other cannot be resolved.	4. Can resolve frequency components very close to each other.

Another basic difference between the MESA spectrum and the Fourier spectrum is that in the former, the power in the spectral peak is proportional to the area under the curve, whereas in the latter, the power is proportional to the height of the peak.

CHAPTER IV

GUIDELINES IN SPECTRAL STUDIES

A. The Akaike Final Prediction Error (FPE).

The determination of the number of filter coefficients (M) which produces the best spectrum is one of the major unsolved problems in MESA. As a general rule, using a very small value for M results in a spectrum with low resolution; on the other hand, a very large value of M , which is close to the total number of data points n results in spurious peaks in the spectrum. A common practice is to use values of M close to half the total number of data points.

T. J. Ulrych (1975) suggested the use of the Akaike final prediction error criterion. This criterion is expressed in terms of the mean square error P_M of the prediction filter used to estimate the MESA spectrum. The idea is to get the least value for P_M at which point the error in the prediction is minimum.

Again our time series will be $x_0, x_1, \dots, x_{n-1}$. The final prediction error is defined by:

$$(4.1) \quad FPE(M) = Q \cdot \hat{P}_M,$$

where $\hat{P}_M$ is the expression in Equation (2.92) and

$$(4.2) \quad Q = \begin{cases} \frac{n+M}{n-M} & , \text{ if the time series is not detrended.} \\ \frac{n+M+1}{n-(M+1)} & , \text{ if the mean is removed.} \\ \frac{n+M+2}{n-(M+2)} & , \text{ if the series is linearly detrended.} \end{cases}$$

To implement the FPE criterion, one computes Equation (4.1) for all $M \leq n-1$. The value of M which gives the smallest $FPE(M)$ determines the number of filter coefficients to use. Table 4.1 lists the FPE values of the sine function in Example 3.1. The FPE values are given for $k = 1, \dots, 97$. They indicate a final prediction error function which attains a minimum at $M = 77$.

TABLE 4.1
FPE VALUES FOR EXAMPLE 3.1 (100 data points)

Number of filter coefficients M	FPE	Number of filter coefficients M	FPE
1	6.67667D-01	21	4.69691D-12
2	2.93505D-01	22	8.43425D-13
3	1.65996D-01	23	7.30391D-13
4	8.86949D-01	24	2.63503D-14
5	2.94595D-01	25	9.82543D-15
6	1.92259D-01	26	6.92124D-15
7	9.21715D-02	27	4.30912D-15
8	1.60844D-03	28	3.76044D-15
9	7.56116D-04	29	1.37526D-15
10	2.98365D-04	30	9.00214D-16
11	2.36250D-04	31	9.20203D-16
12	4.74525D-06	32	4.59301D-16
13	1.09780D-06	33	2.44462D-16
14	5.55949D-07	34	2.46991D-16
15	3.85160D-07	35	2.12235D-16
16	6.04607D-08	36	1.97891D-16
17	2.05121D-08	37	1.08859D-16
18	3.47970D-09	38	1.09643D-16
19	3.10793D-09	39	8.10464D-17
20	1.06768D-11	40	7.90130D-17

TABLE 4.1: Cont'd.

Number of filter coefficients M	FPE	Number of filter coefficients M	FPE
41	7.69202D-17	71	6.31891D-18
42	6.86446D-17	72	6.05030D-18
43	6.86533D-17	73	5.54125D-18
44	6.74952D-17	74	5.77307D-18
45	6.87912D-17	75	5.96710D-18
46	3.21866D-17	76	5.85930D-18
47	2.15421D-17	77	4.79736D-18
48	2.16067D-17	78	5.04358D-18
49	2.01605D-17	79	5.31335D-18
50	2.06381D-17	80	5.38276D-18
51	2.07332D-17	81	5.68238D-18
52	1.95511D-17	82	5.97227D-18
53	1.64197D-17	83	5.76230D-18
54	1.05644D-17	84	6.00796D-18
55	1.01181D-17	85	6.44054D-18
56	1.02351D-17	86	6.10300D-18
57	9.43625D-18	87	5.21797D-18
58	9.29254D-18	88	5.61166D-18
59	9.39571D-18	89	6.01996D-18
60	9.39461D-18	90	6.24087D-18
61	8.66899D-18	91	6.85661D-18
62	7.06601D-18	92	7.35459D-18
63	7.21889D-18	93	7.65730D-18
64	7.07270D-18	94	7.14762D-18
65	7.22471D-18	95	8.58823D-18
66	7.42477D-18	96	1.06101D-17
67	7.28808D-18	97	1.39827D-17
68	6.61568D-18		
69	6.18102D-18		
70	6.39689D-18		

The spectrum for $M = 77$ (Figure 4.1) shows the peaks at $f = 0.05$, 0.10 , 0.20 , and 0.2125 . These are the same frequencies predicted with $M = 51$ (see Figure 3.1C). Although a minimum FPE is obtained for $M = 77$, we studied several spectra for M in the interval $(50, 80)$. The FPE

assumes a relatively flat valley in this interval. Hence, very good estimates of the true peaks are produced. In this case the FPE does not seem to be a very discriminating criterion.

Fig. 4.1. Maximum Entropy Spectrum of Example 3.1 with $M=77$.

As another example of the unreliability of the minimum FPE criterion, we have plotted the FPE for the Middle Atlantic striped bass landings (Example 3.5) in Figure 4.2A. The FPE minimum is at $M = 1$. We have already noted that using very small values of M results in a poorly resolved spectrum. However, as one examines the plot for increasing values of M , a major numerical difference is noted between $M = 25$ and $M = 26$. The FPE function increases at almost every step before $M = 25$ and then takes a sudden drop at $M = 26$. It is interesting to see what happens to the spectrum as M goes from 25 to 26.

Fig. 4.2. A. FPE of the Middle Atlantic striped bass landings.
B. Maximum entropy spectrum with $M = 25$.
C. Maximum entropy spectrum with $M = 26$.

Figure 4.2B shows the spectrum for $M = 25$ and Figure 4.2C for $M = 26$. A major change occurs in the frequency interval $(0.0, 0.1)$ as M changes from 25 to 26. Instead of just one major peak within this interval when $M = 25$, three appear when $M = 26$. It appears that the three peaks are brought into 'focus'.

For many of the other time series that we have analyzed, the same kind of behavior occurs - that is, when the FPE takes a sudden drop in value, the spectrum produces sharper peaks. Whether, of course, the other two peaks that come into 'focus' are significant is another question.

It is also good practice, when deciding what value of M to choose, to refer to the Fourier spectrum of the time series as a guide in determining what value of M would give a good MESA spectrum.

3. Fisher's Test Statistic - The Periodogram Test

The periodogram is one of the earliest forms of spectral estimates. It can actually be derived from the sample spectrum in Equations (1.11), (1.12), (1.13). Using the same notation as in Chapter I, the periodogram values $FF(f)$ are given by the formula:

$$(4.3) \quad FF(f) = \frac{n}{2} \left[a(f)^2 + b(f)^2 \right] ,$$

where $f = k/n$, $k = 1, \dots, m$ if $n = 2m$. If $n = 2m + 1$, Equation (4.3) applies for $f = k/n$, $k = 0, 1, \dots, m$. Since we are using discrete values for f , (4.3) can be simplified into:

$$(4.4) \quad FF_k = \frac{n}{2} \left[a_k^2 + b_k^2 \right] .$$

One of the questions often considered is whether the maximum value of FF_k is significantly above the noise level. With Fisher's test, we proceed the following way to answer the question. Let

FF_L = the largest periodogram ordinate, and

$\sum_{k=1}^m FF_k$ = the sum of the periodogram ordinates.

The test statistic used is

$$(4.5) \quad \bar{F} = FF_L / \left[(1/m) \sum_{k=1}^m FF_k \right] .$$

Fisher (1929) showed that for $g > 0$, the probability that the largest of m normalized terms ($FF_i / \sum_{k=1}^m FF_k$, $i = 1, \dots, m$) should exceed g is

$$(4.6) \quad P_r \left\{ m^{-1} \bar{F} > g \right\} = m (1-g)^{m-1} - \frac{m(m-1)}{2} \frac{(1-2g)^{m-1}}{2} + \dots +$$

$$(-1)^{m-1} \frac{m!}{s! (m-s)!} (1-sg)^{m-1}$$

where s is the largest integer less than $1/g$. Table 4.2 gives the 1 and 5 percentage points for the distribution of ξ as given by W. Fuller (1976).

TABLE 4.2
PERCENTAGE POINTS FOR THE RATIO OF LARGEST
PERIODOGRAM ORDINATE TO THE AVERAGE

Number of ordinates	Probability of larger value	
	0.05	0.01
2	1.950	1.990
3	2.613	2.827
4	3.072	3.457
5	3.419	3.943
6	3.697	4.331
7	3.928	4.651
8	4.125	4.921
9	4.297	5.154
10	4.450	5.358
15	5.019	6.103
20	5.408	6.594
25	5.701	6.955
30	5.935	7.237
40	6.295	7.663
50	6.567	7.977
60	6.785	8.225
70	6.967	8.428
80	7.122	8.601
90	7.258	8.750
100	7.378	8.882
150	7.832	9.372
200	8.147	9.707
250	8.389	9.960
300	8.584	10.164
350	8.748	10.334
400	8.889	10.480
500	9.123	10.721
600	9.313	10.916
700	9.473	11.079
800	9.612	11.220
900	9.733	11.344
1000	9.842	11.454

To illustrate Fisher's test, the lynx data will be used. The sample spectrum (Figure 3.3B) exhibits the tallest peak corresponding to a period of 9.66 years. Recall that the straight line trend was subtracted from the data before calculating the spectrum. The null hypothesis is

$$H_0: x_t = \mu + \beta t + e_t .$$

The alternative hypothesis is

$$H_A: x_t = \mu + \beta t + A \sin \left[(2\pi t / 9.66) + \phi \right] + e_t ,$$

where e_t is the stochastic error term.

The frequencies used for the sample spectrum are $f = i/114$, $i = 1, 2, \dots, 57$. Then

$$\begin{aligned} FF_L / \left[(1/57) \sum_{k=1}^{57} FF_k \right] &= (1.42571 \times 10^8) / \left[(1/57)(2.79734 \times 10^3) \right] \\ &= 29.0511 \end{aligned}$$

From Table 4.2, the 1% point for this value of ξ is about 8.20. So $29.055 > 8.20$ and H_0 is rejected. One concludes that the 9.66 period is statistically significant.

C. The Cumulative Periodogram Test

In analyzing a real time series like the lynx data, our normal course is to remove the linear trend in the data before analysis. That is, we really are fitting the model

$$x_t = \mu + \beta t + e_t .$$

The series that then undergoes analysis is $\hat{e}_t$, the estimated error. Thus a question that one might ask is whether the $\hat{e}_t$'s under study comprise just white noise.

It can be shown (Box and Jenkins 1970) that the power spectrum $P(f)$ for white noise has a constant value $2\sigma_e^2$ (σ_e^2 is the variance for e_t) over the frequency interval $[0.0, 0.5]$. Thus the integrated spectrum for white noise is

$$(4.7) \quad PP(f) = \int_0^f P(v) dv = 2 \sigma_e^2 f ,$$

with f defined in the interval $[0.0, 0.5]$.

Dividing both sides of (4.7) by σ_e^2 gives the equation

$$(4.8) \quad PP(f) / \sigma_e^2 = 2 f .$$

Thus the plot of $PP(f)/\sigma_e^2$ against f is a straight line joining (0,0) to (0.5, 1). For test data, the integrated spectrum can be approximated by the cumulative periodogram CP_k :

$$(4.9) \quad CP_k = \sum_{i=1}^k FF_i / \sum_{j=1}^m FF_j, \quad k = 1, \dots, m,$$

which is compared with the white noise straight line.

To determine the significance of the deviations of the cumulative periodogram of the test data from the theoretical straight line joining (0,0) to (0.5,1), limit lines can be drawn on each side of the theoretical line, using the Kolmogorov - Smirnov test (Box and Jenkins 1970). If indeed the series e_t were to comprise white noise, then the cumulative periodogram would cross these limit lines with a stated probability. It should be pointed out however, that the probabilities are only approximated. The limit lines in the frequency interval $[0.0, 0.5]$ are given by the equations:

$$(4.10) \quad y = 2f \pm K_{\alpha} / \sqrt{q}$$

where

$$(4.11) \quad q = \begin{cases} (n-2) / 2, & \text{if } n \text{ is even} \\ (n-1) / 2, & \text{if } n \text{ is odd,} \end{cases}$$

f is the frequency in the interval $[0.0, 0.5]$, n is the total number of data points, α the stated probability.

The cumulative periodogram of a truly random series would then lie outside the limit lines of Equation (4.10) for an approximate fraction α of the time. Table 4.3 lists K_α versus α . This table is adapted from Box and Jenkins (1970).

TABLE 4.3
VALUES OF α AND K_α NEEDED FOR CUMULATIVE PERIODOGRAM TEST

Probability α	K_α
0.01	1.63
0.05	1.36
0.10	1.22
0.25	1.02

Figure 4.3 is the cumulative periodogram of the lynx data after the linear trend is removed. In this case $n = 114$. One can see that the periodogram has certainly crossed the limit lines, $y = 2f \pm K_{0.05}/\sqrt{q}$, where $q = \sqrt{(114-2)/2} = 7.48$. Our conclusion therefore is that e_t is not white noise and the errors are not independent. The model $x_t = \mu + \beta t + e_t$ is thus inadequate.

D. The Integrated Maximum Entropy Spectrum

In the sample spectrum in Fourier analysis, the power density in a peak (i.e., the strength of contribution of the frequency f) is directly related to the peak amplitude. In the maximum entropy spectrum, the power density is directly related to the area under the spectral peak [see Lacoss (1971)]. So in order to judge the contribution of a spectral peak in MESA, the area under the peak has to be computed. This

Fig. 4.3. Cumulative periodogram of the Lynx Data.

Solid Curve - Cumulative periodogram.

Dashed lines - Limit lines $y = 2 f \pm K_{0.05} / \sqrt{q}$.

suggests the use of an integrated spectrum obtained from the convolution of the MESA spectrum with a rectangular window of height unity and width equal to the effective width of the peaks in the MESA spectrum. The heights of the peaks in the integrated spectrum will then be proportional to the corresponding power densities.

The maximum entropy spectrum as given by Equation (2.35) is

$$(4.12) \quad P(f) = \frac{P_M \Delta t}{\left| 1 + \sum_{m=1}^M a_m e^{-i2\pi f m \Delta t} \right|^2}$$

The integrated spectrum $IP(v)$ will then be of the form

$$(4.13) \quad IP(v_i) = \int_{f_i}^{f_{i+1}} P(f) df,$$

where $f_{i+1} - f_i$ is the bandwidth of integration and

$$v_i = (f_{i+1} + f_i)/2.$$

To illustrate the above ideas, consider the spectrum $P(f)$ in Figure 4.4A. By choosing an appropriate width of integration, one can find the areas under the two spectral peaks. Figure 4.4B shows the estimated integral of $P(f)$ where

$$(4.14) \quad \begin{aligned} IP(v_i) &= \int_{f_i}^{f_{i+1}} P(f) df. \\ IP(v_j) &= \int_{f_j}^{f_{j+1}} P(f) df. \end{aligned}$$

For an M - length prediction error filter, only $M/2 + 1$ discrete frequency spectral components can be obtained (Jensen and Ulrych 1973). Hence to obtain an estimate of the bandwidth of integration we take the Nyquist frequency f_M and divide it by $M/2 + 1$. Once the integrated spectrum is computed, the height of a peak would then be proportional to the power density.

Fig. 4.4. (A) A sample maximum entropy spectrum. (B) The integrated maximum entropy spectrum.

The integrated MESA spectrum for a given time series is very similar in appearance to the Fourier spectrum for the same series, but it is not possible to easily judge whether or not a peak in the integrated MESA spectrum is significantly above the noise level because there does not yet exist a MESA analogue of Fisher's test for the periodogram. The reason is that at the present time no one knows the statistical distribution of integrated spectral estimates for random time series. It is, however, possible to obtain a significance test by a rather cumbersome Monte-Carlo technique which has been described by Jensen and Ulrych (1973). Essentially what one does is to generate a large number

of random time series having the same number of sample points and the same sample spacing as the series being studied. One then calculates the integrated MESA spectrum of each of these random series using the same number of filter coefficients and the same integration window width used to construct the integrated spectrum of the original sample. One can then use this collection of integrated spectra to empirically construct a table giving, as a function of peak height, the probability of obtaining that peak height by chance simply by counting the number of spectra having peaks that high or higher and dividing by the total number of spectra used in the process. The quality of the approximate probabilities obtained in this manner will, of course, improve as more random time series are used in the procedure.

CHAPTER V

THE COMPUTER PROGRAM AND ITS USE

A. Explanation of the Subroutines

Essentially, the program consists of the following routines:

- (1) Main program - which initializes the data values and other parameters needed in spectral analysis.
- (2) Subroutine TRANS1 - which provides different calls to SPECTR.
- (3) Subroutine SPECTR - which does spectral analysis using both Fourier and MESA methods. In turn, SPECTR calls:
 - (a) Subroutine DETRND - a routine to be provided by the user if detrending is done.
 - (b) Subroutine SFT - computes the Fourier sine and cosine terms needed to produce the sample spectrum.
 - (c) DCADRE - a built-in IMSL subroutine used in computing the integrated spectrum if it is desired.
 - (d) Double Precision Function SPCVAL - computes the function used in integrating the spectrum.
 - (e) Subroutines DOPLT, DOPLT1, DOPLT2, DOPLT3, DOPLT4, DOPLT5, DOPLT6, DOPLT7, DOPLT8, DOPLT9, SINGLE, GPHBGN, GPHBG1, GPHEND, ANGTIC - perform the necessary plotting using the DISSPLA routines.

1. The Main Program

The parameters which are given initial values are read in this section of the program and are transferred to TRANS1 and SPECTR by the common statement:

COMMON/SPECBL/DELTA, YEAR(600), Y(600), XD(600), VMAX, Z(600), V(600),
 S(600), FPE(600), XDIFF(600), DATE(600), X(600),
 ACS(600), BSN(600), AMP(600), ASF(600), PERIOD(600),
 NPTS, NFILT, NV, MCOUNT(600), IPRINT, IPLOT, ITREND,
 IFPEPL, IINTSP, IAUTCV, IYWACV

Table 5.1 lists in alphabetical order the variables which are read, their dimensions, types and what they represent. Currently, the dimensions allow a maximum of 600 data points to be analyzed by SPECTR.

TABLE 5.1
 LIST OF VARIABLES READ IN THE MAIN PROGRAM

Variable	Dimension	Type	Explanation
DATE	600	REAL*8	time array (times at which sample values occur).
DELTA		REAL*8	sampling interval.
IAUTCV		I*4	indicator for option to plot MESA estimate of autocovariance function.
IFPEPL		I*4	indicator for option to print/plot final prediction error.
IINTSP		I*4	indicator for option to compute integrated MESA spectrum.
IPLOT		I*4	indicator for option to plot spectra.

TABLE 5.1: Cont'd.

Variable	Dimension	Type	Explanation
IPRINT		I*4	indicator for option to print output.
ITREND		I*4	indicator for option to detrend data.
IYWACV		I*4	indicator for option to compute and plot Yule-Walker estimate of the autocovariance function.
MCOUNT	600	I*4	contains the array of numbers which correspond to the number of filter coefficients desired in each MESA spectrum.
NAM	14	Real*4	title for the data set.
NFILT		I*4	total number of maximum entropy spectra to be calculated with different values of the number of filter coefficients.
NPTS		I*4	number of data points in time series.
NV		I*4	number of frequencies desired in output; should not exceed a maximum value of 599.

TABLE 5.1: Cont'd.

Variable	Dimension	Type	Explanation
VMAX		Real*8	maximum frequency for the output spectrum; has a maximum value of 0.5, the Nyquist frequency.
XLBL	2	Real*4	label for time axis (x-axis) used in plotting.
Y	600	Real*8	time series to be analyzed.
YLBL	2	Real*4	label for time series data (y-axis) used in plotting.

The indicators IAUTCV, IFPEPL, IINTSP, IPLOT, IPRINT, ITREND, IYWACV should be initialized to values of 0 or 1. In each case, 1 is a positive response and 0 is a negative response.

The variables NAM, XLBL, YLBL, which are read in the main program, appear in the common block BLKLBL. These are mainly used for identification of data and plots. BLKLBL is shared by the Main program, TRANS1, SPECTR and DOPLT. The form of the statement is

```
COMMON/BLKLBL,NAM(14), XLBL(2), YLBL(2) .
```

The input cards, all of which are read in the main program, must have the following formats. The formats appear in the order the cards are read.

- (1) Input card on first READ statement.

XLBL		YLBL		DELTA	NPTS	NFILT	NV	VMAX
\$		\$						
2A4	2X	2A4	2X	E10.0	15	15	15	E10.0

XLBL and YLBL must terminate with a dollar sign.

- (2) Input card on second READ statement.

NAM S

14A4

NAM must also end in a dollar sign.

- (3) Input cards for third READ statement.

MCOUNT(1) MCOUNT(2) MCOUNT(NFILT)

~~~~~   ~~~~~   ~~~~~

I5   I5   I5

- (4) Input cards for fourth READ statement. This set contains the time series to be analyzed.

DATE(1)	Y(1)
DATE(2)	Y(2)
.	.
.	.
.	.
DATE(NPTS)	Y(NPTS)
<u> </u>	<u> </u>
F4.0	F10.0

(5) Input card for the fifth READ statement.

IPRINT	IPLOT	ITREND	IFPEPL	IINTSP	IAUTCV	IYWACV
<u>I5</u>	<u>I5</u>	<u>I5</u>	<u>I5</u>	<u>I5</u>	<u>I5</u>	<u>I5</u>

A flowchart of the main program is given in Figure 5.1.

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Fig. 5.1. Flowchart of the Main Program. Numbers (e.g., 1 and 500) refer to statement numbers in the program.

2. SUBROUTINE TRANS1

The variables in this subroutine are those that appear in the common area SPECLBL. This subroutine allows the user to perform transformations on the raw data (e.g., take logarithms) before doing spectral analysis. Up to five different spectra with five different transformations are allowed.

The basic statements in TRANS1 are the following:

(1) SUBROUTINE TRANS1

Optional user supplied statements for desired transformation on data.

CALL SPECTR

.

.

.

RETURN

(2) ENTRY TRANS2

Optional user supplied statements for second desired transformation.

CALL SPECTR

.

.

.

RETURN

(3) ENTRY TRANS3

Optional user supplied statements.

CALL SPECTR

.

.

.

RETURN

(4) ENTRY TRANS4

Optional user supplied statements.

CALL SPECTR

.

.

.

RETURN

(5) ENTRY TRANS5

Optional user supplied statements.

CALL SPECTR

.

.

.

RETURN

If the user wants to analyze only the input data, he need not supply additional statements and any of the 4 additional entries can be effectively deactivated by omitting CALL SPECTR.

An example of data transformation using the second Entry call (i.e., ENTRY TRANS2) is the following:

```
ENTRY TRANS2
DO 20 I = 1, NPTS
  Y(I) = DLOG(Y(I))
20 CONTINUE
CALL SPECTR
RETURN
```

In the preceding statements, the natural logarithm of the input data $Y(I)$, $I=1, \dots, NPTS$ is taken and the result is stored in $Y(I)$ again. Thus when SPECTR is called, the logarithmically transformed data are analyzed. The program listing for SUBROUTINE TRANS1 given in Appendix III gives the configuration of statements needed to first analyze the original data and then its logarithm.

3. SUBROUTINE SPECTR

The main function of SPECTR is to perform the following:

- (1) Detrend the input data by calling DETRND if desired;
- (2) Compute:
  - (a) the Yule-Walker estimates of the autocovariance function if desired.
  - (b) the Fourier spectrum.
  - (c) the periodogram, ordinates and their sum.
  - (d) the cumulative periodogram.
  - (e) the maximum entropy spectrum.

- (f) the autocorrelation function resulting from the maximum entropy computations.
  - (g) the integrated maximum entropy spectrum if desired.
  - (h) the final prediction error if desired.
- (3) Print and plot computed values.

A flowchart of subroutine SPECTR appears in Figure 5.2. Table 5.2 lists the parameters in subroutine SPECTR which appear in common blocks on DIMENSION statements. This list includes the variables which must appear in the calling program and the arrays whose dimension the user might want to alter if he desires to make changes in the programs. Variables that are only used internally in SPECTR or in subroutines called by SPECTR are not included in the list.

4. SUBROUTINE DETRND

DETRND is called by SPECTR with the statement:

```
CALL DETRND (NPTS, DATE, YEAR, Y, XDIFF, X, SFPE)
```

where

- NPTS - number of data points, i.e., length of time series.
- DATE - time array that must not be destroyed.
- YEAR - user-evaluated time values; these are to be initialized in DETRND.
- Y - input time series values; these values must not be destroyed.
- XDIFF - array containing the detrended data; computed in DETRND.
- X - array containing the trend of the data; computed in DETRND.

TABLE 5.2
LIST OF PARAMETERS IN SUBROUTINE SPECTR

Variable	Dimension	Type	Explanation
A	600	Real*8	Vector of working storage.
ACFN	600	Real*8	Vector of working storage.
ACS	600	Real*8	Vector of working storage.
AMP	600	Real*8	Vector of working storage.
ASF	600	Real*8	Vector containing Fourier power spectral values.
B	600	Real*8	Vector of working storage.
BSN	600	Real*8	Vector of working storage.
CUMPER	300	Real*8	Array of cumulative periodogram values.
DATE	600	Real*8	Time values for input series.
DELTA		Real*8	Sampling interval ('delta T').
FPE	600	Real*8	Array containing final prediction error.
FPER	300	Real*8	Array of periodogram frequencies.
IAUTCV		I*4	Autocovariance plotting flag (Burg estimate).
IFPEPL		I*4	Final prediction error flag.
IINTSP		I*4	Integrated spectrum flag.

TABLE 5.2: Cont'd

Variable	Dimension	Type	Explanation
IPLOT		I*4	Plotting flag.
IPRINT		I*4	Printing flag.
ITREND		I*4	Detrending flag.
IYWACV		I*4	Yule-Walker autocovariance flag.
MCOUN	600	I*4	Vector of prediction error filter lengths to use.
NAM	14	I*4	Title for the data set.
NFILT		I*4	Number of prediction error filter lengths to try.
NPTS		I*4	Number of data points in time series.
NV		I*4	Number of frequencies at which spectrum is calculated.
PERDG	300	Real *8	Array of periodogram values.
PERIOD	600	Real*8	Vector of working storage. Upon return to the main program, this vector contains the periods corresponding to the frequencies in V.
PPER	300	Real*8	Array of periodogram periods corresponding to FPER.

TABLE 5.2: Cont'd

Variable	Dimension	Type	Explanation
S	600	Real*8	Maximum entropy spectral values are stored in this array.
STATE	2	Real*4	Label for dependent variable (y - axis)
V	600	Real*8	Output vector of frequencies at which spectrum is calculated.
VMAX		Real*8	Maximum frequency at which spectrum is to be calculated.
X	600	Real*8	Upon returning from DETRND (if detrending is done), it should contain the trend of the input data. It is later used as a vector of working storage.
XD	600	Real*8	If detrending is done, it should contain the detrended data; if not, it is the original data. Spectral analysis is applied on XD.
XDIFF	600	Real*8	Upon returning from DETRND, it is the detrended data. It is later used as a working vector.

TABLE 5.2: Cont'd

Variable	Dimension	Type	Explanation
Y	600	Real*8	Input time series values.
YEAR	600	Real*8	Vector of working storage (can be used as time variable in DETRND).
YTIME	2	Real*4	Label for the time axis.
Z	600	Real*8	Prediction error filter coefficients.

SFPE - number of components in the trend of the time series to be removed. For example, if the mean is removed, SFPE = 1; if the intercept and slope are to be removed, SFPE = 2, and so on. SFPE is given the value 0 in SPECTR if detrending is not done.

Should the user choose to analyze the original data without detrending, the dummy subroutine in Table 5.3 may be used.

If the mean of the input data is to be removed, i.e., the model $x_t = \mu + e_t$ is to be fitted, then the subroutine in Table 5.4 will be appropriate to use.

Table 5.5 contains a detrending subroutine in which the linear trend of a time series is eliminated. A least-squares routine BLSQ is called to perform the fitting of the straight line to the data. A write-up of BLSQ appears in Appendix II. SUBROUTINE BLSQ is available from disk at ORNL (see, for example, control card for "BROOKS.LOAD. MODULES" in Tables 6.3 and 6.4).

Table 5.3. Dummy subroutine DETRND

Statement number columns 1-5	Continuation column 6	Fortran statement columns 7-72
		SUBROUTINE DETRND(NPTS,DATE,YEAR,Y,XDIFF,X,SFPE) IMPLICIT REAL*8(A-H,O-Z) DIMENSION DATE(1), YEAR(1), Y(1), XDIFF(1), X(1) RETURN END

Table 5.4. A detrending subroutine for removal of the mean of a time series

Statement number columns 1-5	Continuation column 6	Fortran statement columns 7-72	Explanation
		SUBROUTINE DETRND(NPTS,DATE,YEAR,Y,XDIFF,X,SFPE)	
		IMPLICIT REAL*8(A-H,O-Z)	
		DIMENSION DATE(1), YEAR(1), Y(1), XDIFF(1), X(1)	
		SFPE = 1.0	Since only the mean is removed, SFPE = 1.
		AVRAGE = 0.0	
		DO 100 I = 1, NPTS	
		AVRAGE = AVRAGE + Y(I)	Compute the mean of the Y array.
100		CONTINUE	
		AVRAGE = AVRAGE / NPTS	
		DO 300 I = 1, NPTS	
		X(I) = AVRAGE	Store the trend of the data in X array.
		XDIFF(I) = Y(I) - X(I)	Store the detrended data in XDIFF array.
300		CONTINUE	
		RETURN	
		END	

Table 5.5. A detrending subroutine for the removal of the linear trend in a time series

Statement number columns 1-5	Continuation column 6	Fortran statement columns 7-72	Explanation
		SUBROUTINE DETRND(NPTS,DATE,YEAR,Y,XDIFF,X,SFPE)	
		IMPLICIT REAL*8(A-H,O-Z)	
		DIMENSION DATE(1),YEAR(1),Y(1),XDIFF(1),X(1),	
	1	A(601,3),IRUN(601),COEF(2),COEF2(2),T(2,2),T1(2,2),	
	2	VAR(2,2)	
		NROW = 601	NROW is the row dimension of the least squares matrix A.
		NLSCOL = 2	NLSCOL is the number of coefficients of the linear fit.
		NT = 2	NT is the row dimension of the matrices T and T1, which are computed in the least squares routine BLSQ.
		SFPE = NLSCOL	SFPE = NLSCOL since there are two linear coefficients - the intercept and slope - to be removed.
		DO 50 I = 1,NPTS	
		YEAR(I) = I - 0.5	Initialize the time variable
		A(I,1) = 1.0	Set up the least squares matrix.
		A(I,2) = YEAR(I)	

Table 5.5. (continued)

Statement number columns 1-5	Continuation column 6	Fortran statement columns 7-72	Explanation
50	1	CONTINUE CALL BLSQ(A,Y,COEF,RESID,IRUN,T,T1,NPTS,NLSCOL, NROW,NSLCOL) DO 80 I = 1,NLSCOL SUM = 0.0 DO 70 J = 1,NLSCOL SUM = SUM + T1(I,J) * COEF(J)	Call the least squares routine BLSQ to perform the linear fit. Compute the coefficients of the linear fit using values returned from BLSQ.
70		CONTINUE COEF2(I) = SUM	
80		CONTINUE DO 90 I = 1,NLSCOL COEF(I) = COEF2(I)	COEF(1) contains the intercept of the time series. COEF(2) contains the slope of the trend line.
90		CONTINUE DO 100 I = 1,NPTS X(I) = A(I,NLSCOL + 1)	The trend of the data comes back in A(I,NLSCOL + 1) and is stored in X array.

Table 5.5. (continued)

Statement number columns 1-5	Continuation column 6	Fortran statement columns 7-72	Explanation
100		$XDIFF(1) = Y(1) - X(1)$ CONTINUE RETURN END	The detrended data is stored in XDIFF array.

SUBROUTINE BLSQ returns (in the array T1) the information needed to compute the variance - covariance matrix of the coefficients (i.e. the slope and intercept) and these values could be stored in the array VAR which is not used in the preceding program (Table 5.5). A SUBROUTINE DETRND which does add this final step is given in the program listings in Appendix IV.

It should be obvious from studying Table 5.5 and Appendix II how to remove more complicated detrending functions using BLSQ or even a non-linear least squares program if needed.

5. SUBROUTINE SFT

The computation of the factors

$$\sum_{j=0}^{n-1} x_j \cos 2\pi jf, \quad \sum_{j=0}^{n-1} x_j \sin 2\pi jf,$$

from Equations (1.12) and (1.13) respectively, which are used in the evaluation of the sample Fourier spectrum, is done in SFT. Subroutine SFT is called by SPECTR with:

CALL SFT(XD, ACS, BSN, V, KF1, NYR, PI) .

The corresponding arguments in SFT are given by:

SUBROUTINE SFT(X, A, B, V, K, N, PI)

These arguments are described in Table 5.6.

6. DOUBLE PRECISION FUNCTION SPCVAL

The double precision function SPCVAL computes for any frequency the maximum entropy spectral value $P(f)$ as given by Equation (2.35). SPCVAL is used in the integrated maximum entropy spectrum calculations. It is the function which is integrated by the IMSL subroutine DCADRE, which is described in Appendix I. The IMSL subroutine DCADRE is available from disk at ORNL (see, for example, control card for "JDAIMSL1" in Tables 6.3 and 6.4). Table 5.7 lists the important parameters in SPCVAL.

TABLE 5.6
A LIST OF THE PARAMETERS IN SUBROUTINE SFT

Variable	Dimension	Type	Explanation
A	1	Real*8	The array containing the factors $\sum_{j=0}^{n-1} x_j \cos 2 \pi j f$ upon return. It has a maximum dimension of 600.
B	1	Real*8	The array containing the factors $\sum_{j=1}^{n-1} x_j \sin 2 \pi j f$ upon return. It has a maximum dimension of 600.
K		I*4	Number of frequencies at which spectrum is calculated.
N		I*4	Number of points in the series under analysis.
PI		Real*8	The real number π .

TABLE 5.6: Cont'd

Variable	Dimension	Type	Explanation
V	1	Real*8	Output vector of frequencies at which spectrum is calculated.
X	1	Real*8	Time series under analysis.

TABLE 5.7

A LIST OF PARAMETERS IN THE FUNCTION SUBROUTINE SPCVAL

Variable	Dimension	Type	Explanation
AMP	600	Real*8	The array containing P_M .
DTIME		Real*8	The time interval Δt .
FREQ		Real*8	The frequency f .
M		1*4	The number of filter coefficients M .
PI		Real*8	The real number π .
SPCVAL		Real*8	The spectral value $P(f)$.
Z	600	Real*8	The array of filter coefficients $a_{i,M}$.

7. SUBROUTINE DOPLT

Subroutine DOPLT is a plotting routine which calls other subroutines primarily from the plotting package DISSPLA (DISSPLA is a universal package that performs the necessary plotting). DISSPLA is available from disk at ORNL (see, for example, control card for "DISSPLA.LOAD" in Tables 6.3 and 6.4). There are nine ENTRY statements, the function of each are listed as follows:

(1) ENTRY DOPLT1 - plots the original data against the time values; if the data are detrended, the trend is also plotted on the same set of axes.

(2) ENTRY DOPLT2 - if detrending is performed, the detrended data are plotted against the time values; the line $y = 0$ is also included in the plot and serves as a reference line for the residuals. Furthermore, a plot of the residuals against the fitted values or trend is done.

(3) ENTRY DOPLT3 - plots the Yule-Walker autocorrelation function against the lag values if IYWACV = 1.

(4) ENTRY DOPLT4 - plots the Fourier sample spectrum against frequency; it also plots the cumulative Fourier power spectrum against frequency. The 95% Kolmogorov-Smirnov bounds are also plotted on the latter graph.

(5) ENTRY DOPLT5 - plots the Burg autocorrelation function against the number of filter coefficients if IAUTCV = 1.

(6) ENTRY DOPLT6 - plots the maximum entropy spectrum against frequency.

(7) ENTRY DOPLT7 - plots the integrated maximum entropy spectrum against frequency if IINTSP = 1.

(6) ENTRY DOPLT8 - plots the final prediction error against the number of filter coefficients if IFPEPL = 1.

(9) ENTRY DOPLT9 - plots the periodogram and cumulative periodogram as a function of frequency. The 95% Kolmogorov-Smirnov bounds are included in the latter plot.

The other subroutines called by DOPLT after each ENTRY statement are mainly for setting up the axes for plotting. They are:

(1) SUBROUTINE SINGLE - transforms double precision variables into single precision. This is necessary since DISSPLA routines handle only Real*4 information.

(2) SUBROUTINE ANGTIC - controls the number of tick marks and angular position of the numbers on the axes.

(3) SUBROUTINE GPHBGN - sets up the initial x and y axes for plotting according to the range of the data pairs.

(4) SUBROUTINE GPHBGL - sets up the initial x and y axes for plotting the FPE. This uses logarithmic scaling of the axes, as contrasted with GPHBGN which makes use of linear scaling.

(5) SUBROUTINE GPHEND - plots the specific curve and draws a frame about the picture.

All the necessary parameters for plotting are transferred to Subroutine DOPLT by the two common blocks SPECBL and BLKLBL.

CHAPTER VI

ILLUSTRATION OF THE USE OF THE PROGRAM

A. An Ecological Example

The time series we will consider is the striped bass catch per unit effort data for the commercial fishery in the Hudson River from 1955 - 1975, a series of length 21 years. The set of input cards appears in Table 6.1. The first card contains the labels for the x and y axes, the time interval, the number of points in the time series, the number of times MESA is to be performed, the number of frequencies and the maximum frequency. The second card has the label for the time series data. The third card is the number of filter coefficients used. The time series follows on cards 4 - 24. The last card has all the options to print, to plot, to detrend, to compute the FPE, to compute the integrated spectrum to compute the autocorrelation function from MESA, and to compute the Yule - Walker estimates of the autocorrelation. All these option indicators are given the value 1.

The data were detrended using a SUBROUTINE DETRND which removed a linear trend by calling BLSQ, the least squares procedure (DETRND is in Appendix IV and BLSQ in Appendix II). Spectral analysis was done on both the raw data and the natural logarithm of the raw data. However, only output for the raw data will be presented. The printed output appears in Figure 6.1 and the plots in Figure 6.2A thru Figure 6.2L. The printout on the residual sum of squares, coefficients of the linear fit, and variance-covariance matrix come from DETRND. DISSPLA plot messages are also inserted.

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Table 6.1. Input cards for the Hudson River catch per unit effort data

Card									
1	YEAR	__	\$	__	CATCH	__	\$	__	1.0
2	HUDSON RIVER CATCH PER UNIT EFFORT\$								
3	__ 18								
4	1955	_	1897.	} TIME SERIES DATA					
5	1956	_	2550.						
6	1957	_	2456.						
7	1958	_	2479.						
8	1959	_	3985.						
9	1960	_	4089.						
10	1961	_	2372.						
11	1962	_	1670.						
12	1963	_	2039.						
13	1964	_	1378.						
14	1965	_	1952.						
15	1966	_	2842.						
16	1967	_	3595.						
17	1968	_	3397.						

Table 6.1. (continued)

Card	
18	1969 _ 4852.
19	1970 _ 2608.
20	1971 _ 1743.
21	1972 _ 1213.
22	1973 _ 4273.
23	1974 _ 1429.
24	1975 _ 1480.
25	1 1 1 1 1 1 1 1 TIME SERIES DATA

NOTE: ' _ ' indicates a space.

ANALYSIS OF THE ORIGINAL INPUT DATA WITHOUT ANY PRIOR TRANSFORMATIONS

HUDSON RIVER CATCHE PER UNIT EFFORTS

YEAR S	C/E S
1.95500 03	1.8970000 03
1.95600 03	2.5500000 03
1.95700 03	2.4560000 03
1.95800 03	2.4790000 03
1.95900 03	3.9850000 03
1.96000 03	4.0890000 03
1.96100 03	2.3720000 03
1.96200 03	1.6700000 03
1.96300 03	2.0390000 03
1.96400 03	1.3780000 03
1.96500 03	1.9520000 03
1.96600 03	2.0420000 03
1.96700 03	3.5950000 03
1.96800 03	3.3970000 03
1.96900 03	4.8520000 03
1.97000 03	2.6080000 03
1.97100 03	1.7430000 03
1.97200 03	1.2130000 03
1.97300 03	4.2730000 03
1.97400 03	1.4290000 03
1.97500 03	1.4800000 03

RESIDUAL SUM OF SQUARES = 2.2522230 07

COEFFICIENTS OF THE LINEAR FIT

INTERCEPT = 2.7189260 03 SLOPE = -1.2681820 01

VARIANCE-COVARIANCE MATRIX

ROW	COLUMN	1	COLUMN	2
1		2.26171597707163800 05	-1.61642775445846000 04	
2		-1.61642775445846000 04	1.53945509424615200 03	

THE SPECTRAL ANALYSIS WILL BE APPLIED TO THE DETRENDED DATA

YEAR S	C/E S	DETRENDED C/E S
1.95500 03	1.8970000 03	-6.1548480 02
1.95600 03	2.5500000 03	-1.4940300 02
1.95700 03	2.4560000 03	-2.3112120 02
1.95800 03	2.4790000 03	-1.9543940 02
1.95900 03	3.9850000 03	1.3232420 03
1.96000 03	4.0890000 03	1.4399240 03
1.96100 03	2.3720000 03	-2.6439190 02
1.96200 03	1.6700000 03	-9.5171210 02
1.96300 03	2.0390000 03	-5.7203030 02
1.96400 03	1.3780000 03	-1.2203440 03
1.96500 03	1.9520000 03	-6.3166670 02
1.96600 03	2.0420000 03	2.6901520 02
1.96700 03	3.5950000 03	1.0341470 01
1.96800 03	3.3970000 03	4.4931480 02
1.96900 03	4.8520000 03	2.3170610 03
1.97000 03	2.6080000 03	4.4742420 01
1.97100 03	1.7430000 03	-7.6657580 02
1.97200 03	1.2130000 03	-1.2434440 03
1.97300 03	4.2730000 03	1.7487880 01
1.97400 03	1.4290000 03	-1.0424300 03
1.97500 03	1.4800000 03	-9.7444440 02

Fig. 6.1. Standard computer output.

PLOT NO. 3 WITH THE TITLE
RESIDUALS VS. FITTED VALUES
HAS BEEN COMPLETED.

PLOT FILE BEARS
PLOT 3 12.50.01 TUES 13 JUN, 1978 JCS-244L . 00PL DISSEPER VER 6.0

DATA FOR PLOT

NO. OF CURVES DRAWN 1

HORIZ. AXIS LENGTH 7.0 INCH.
VERT. AXIS LENGTH 5.0 INCH.

HORIZ. ORIGIN 0.2000 ON VERT. ORIGIN 0.1500 ON

HORIZ. AXIS LINEAR
STEP SIZE 0.5000 02 UNITS/INCH

VERT. AXIS LINEAR
STEP SIZE 0.2000 03 UNITS/INCH

LOCATION OF CURRENT PHYSICAL ORIGIN -
X= 1.50 Y= 2.50 INCHES -
FROM LOWER LEFT CORNER OF PAGE -

YULE-WALKER ESTIMATE OF AUTOCORRELATION FUNCTION

LAG	AUTOCORR.
0.0	1.00000 00
1.0000 00	2.072630-01
2.0000 00	-1.052380-01
3.0000 00	-1.037380-01
4.0000 00	-1.688130-01
5.0000 00	-0.560050-01
6.0000 00	-1.508720-01
7.0000 00	0.226980-02
8.0000 00	9.850780-02
9.0000 00	2.190820-01
1.0000 01	2.067000-01
1.1000 01	-7.057590-02
1.2000 01	-1.067090-01
1.3000 01	5.227570-02
1.4000 01	-1.552020-02
1.5000 01	-1.201860-01
1.6000 01	-3.052390-02
1.7000 01	5.378180-02
1.8000 01	-0.770930-02
1.9000 01	0.025860-02
2.0000 01	3.500210-02

PERIODOGRAM

K	FREQ.	PERIOD	PERIODOGRAM ORDINATE	CUMULATIVE PERIODOGRAM
1	0.00762	21.00000	6.708750 05	0.02979
2	0.09520	11.50000	7.527130 06	0.36400
3	0.10286	7.00000	1.608220 06	0.52776
4	0.19000	5.25000	7.030060 05	0.55897
5	0.23810	0.20000	2.600110 06	0.67659
6	0.20571	3.50000	3.525960 06	0.03315
7	0.33333	1.00000	0.256650 05	0.05205
8	0.38095	2.62500	3.246100 05	0.06606
9	0.02057	2.33333	6.907120 05	0.09731
10	0.07619	2.10000	2.312070 06	1.00000

SUM OF PERIODOGRAM ORDINATES = 2.252220 07

MAXIMUM ORDINATE = 7.527130 06 OCCURRING AT FREQUENCY = 0.521810-02 PERIOD = 1.050000 01

FISHER STATISTIC = 3.302090 00

NUMBER OF ORDINATES = 10

Fig. 6.1. (continued).

SEA OF POSITIVE COORDINATES = 0.074010 00

WCOEFF = 10

AUTOCORRELATION FUNCTION VALUES

LAS	AUTOCORR.
1.0	1.000000 00
1.000000 00	2.1501070-01
2.000000 00	-1.0206540-01
3.000000 00	-2.1677930-01
4.000000 00	-1.5189900-01
5.000000 00	-6.0896050-01
6.000000 00	-1.2985870-01
7.000000 00	1.0230490-01
8.000000 00	2.0425320-01
9.000000 00	3.4944520-01
1.000000 01	4.5609720-01
1.100000 01	-1.6130180-01
1.200000 01	-3.6101340-01
1.300000 01	-9.6250570-02
1.400000 01	-1.6812050-01
1.500000 01	-3.9043610-01
1.600000 01	1.113420-01
1.700000 01	3.6761020-01
1.800000 01	2.5962050-01

MAXIMUM PEAKS AND POBBLE SPECTRA

NYQUIST FREQUENCY = 5.0000-01

RESOLUTION BANDWIDTH = 5.0000-02

FINAL PREDICTION ERROR = 1.1010 07

WCOEFF = 10	FREQUENCY	PERIOD	MAXIMUM PEAKS	POBBLE AMPLITUDE
I = 1	0.0	399.000	0.1649160 00	1.2747060-11
I = 2	0.0025126	199.000	0.1892100 00	1.9682480 00
I = 3	0.0050251	132.667	0.2631400 00	7.0156490 00
I = 4	0.0075377	99.500	0.3899350 00	1.6471800 01
I = 5	0.0100503	79.600	0.5752780 00	2.8754400 01
I = 6	0.0125628	66.333	0.8278010 00	4.1897070 01
I = 7	0.0150754	54.457	5.1594640 00	6.1420490 01
I = 8	0.0175879	49.750	5.5887340 00	8.0784950 01
I = 9	0.0201005	40.222	6.1383210 00	1.0140230 02
I = 10	0.0226131	39.800	6.8817410 00	1.2262440 02
I = 11	0.0251256	36.182	7.7450670 00	1.8380270 02
I = 12	0.0276382	33.167	8.9128200 00	1.6429260 02
I = 13	0.0301508	30.615	1.0435900 05	1.8348860 02
I = 14	0.0326633	28.029	1.2442640 05	2.0083290 02
I = 15	0.0351759	26.533	1.5117010 05	2.1548840 02
I = 16	0.0376884	24.875	1.8692680 05	2.2833580 02
I = 17	0.0402010	23.012	2.3494590 05	2.3403180 02
I = 18	0.0427136	22.111	2.9851490 05	2.4506500 02
I = 19	0.0452261	20.947	3.7936680 05	2.4974860 02
I = 20	0.0477387	19.900	4.7149150 05	2.5289600 02
I = 21	0.0502513	18.952	5.8566220 05	2.5545150 02
I = 22	0.0527638	18.091	6.3026770 05	2.5888670 02
I = 23	0.0552764	17.308	6.9702010 05	2.6491320 02
I = 24	0.0577889	16.591	6.1830990 05	2.7527230 02
I = 25	0.0603015	15.920	5.6441140 05	2.9141870 02
I = 26	0.0628141	15.304	5.0576200 05	3.1419630 02
I = 27	0.0653266	14.741	4.5394920 05	3.4170390 02
I = 28	0.0678392	14.214	4.1324120 05	3.7938210 02
I = 29	0.0703518	13.724	3.8416070 05	4.2022730 02
I = 30	0.0728643	13.267	3.6600780 05	4.6500400 02
I = 31	0.0753769	12.839	3.5804750 05	5.1241490 02
I = 32	0.0778894		3.6007510 05	5.6117090 02

Fig. 6.1. (continued)

I= 31	Q.0804020	12.438	3.726488 05	6.190372 02
I= 32	Q.3824146	12.041	3.973893 05	6.578773 02
I= 35	Q.9858271	11.704	4.178386 05	7.136888 02
I= 36	Q.0878397	11.771	4.483228 05	7.468225 02
I= 37	Q.0904423	11.056	4.496302 05	7.453812 02
I= 38	Q.0924688	11.757	7.238816 05	8.198481 02
I= 39	Q.0958778	10.874	9.867618 05	8.892519 02
I= 40	Q.0979899	10.205	1.308989 04	9.731923 02
I= 41	Q.1005025	9.959	1.957973 04	9.911986 02
I= 42	Q.1030151	9.707	3.288812 04	9.931720 02
I= 43	Q.1055276	9.476	6.173198 06	9.102207 02
I= 44	Q.1080402	9.256	1.729061 07	9.110136 02
I= 45	Q.1105528	9.045	2.216152 07	9.043922 02
I= 46	Q.1130651	9.888	1.535061 07	8.967527 02
I= 47	Q.1155779	9.652	7.953889 06	9.425891 02
I= 48	Q.1180905	8.868	4.686280 06	8.688713 02
I= 49	Q.1206030	8.292	1.110926 06	9.430199 02
I= 50	Q.1231156	8.122	2.310879 06	9.188770 02
I= 51	Q.1256281	7.940	1.852728 06	7.926781 02
I= 52	Q.1281407	7.808	1.572906 06	7.688998 02
I= 53	Q.1306531	7.658	1.392193 06	7.363670 02
I= 54	Q.1331658	7.509	1.267302 06	7.071732 02
I= 55	Q.1356784	7.370	1.170620 06	6.777085 02
I= 56	Q.1381910	7.236	1.082400 06	6.481078 02
I= 57	Q.1407035	7.107	9.989781 05	6.183686 02
I= 58	Q.1432161	6.982	8.880023 05	5.881598 02
I= 59	Q.1457286	6.862	7.693913 05	5.578818 02
I= 60	Q.1482412	6.746	6.530678 05	5.265127 02
I= 61	Q.1507538	6.633	5.481885 05	4.980278 02
I= 62	Q.1532663	6.525	4.491370 05	4.680570 02
I= 63	Q.1557789	6.419	3.701870 05	4.283199 02
I= 64	Q.1582915	6.317	3.072571 05	3.866198 02
I= 65	Q.1608040	6.219	2.577939 05	3.468768 02
I= 66	Q.1633166	6.121	2.198788 05	3.051635 02
I= 67	Q.1658291	6.030	1.900078 05	2.617536 02
I= 68	Q.1683417	5.940	1.678768 05	2.172151 02
I= 69	Q.1708543	5.851	1.503787 05	1.726528 02
I= 70	Q.1733668	5.768	1.376880 05	1.308158 02
I= 71	Q.1758794	5.686	1.288302 05	9.626892 01
I= 72	Q.1783920	5.606	1.221817 05	8.708872 01
I= 73	Q.1809045	5.528	1.185812 05	7.800015 02
I= 74	Q.1834171	5.452	1.173381 05	6.958302 02
I= 75	Q.1859296	5.378	1.185368 05	6.178579 02
I= 76	Q.1884422	5.307	1.223088 05	5.233599 02
I= 77	Q.1909548	5.237	1.290915 05	4.670228 02
I= 78	Q.1934673	5.169	1.392618 05	3.992723 02
I= 79	Q.1959799	5.101	1.581259 05	3.049592 02
I= 80	Q.1984925	5.038	1.752818 05	2.858719 02
I= 81	Q.2010050	4.975	2.052909 05	2.193567 02
I= 82	Q.2035176	4.918	2.887097 05	1.492711 02
I= 83	Q.2060302	4.854	3.131810 05	1.758871 02
I= 84	Q.2085427	4.795	4.128788 05	1.979306 02
I= 85	Q.2110553	4.738	5.781508 05	1.166255 02
I= 86	Q.2135678	4.682	8.530450 05	1.116268 02
I= 87	Q.2160804	4.628	1.167009 06	1.410287 02
I= 88	Q.2185930	4.575	2.180722 06	1.509167 02
I= 89	Q.2211055	4.523	3.863302 06	1.558676 02
I= 90	Q.2236181	4.472	4.585735 06	1.567179 02
I= 91	Q.2261307	4.422	3.850801 06	1.587631 02
I= 92	Q.2286432	4.378	2.201056 06	1.496503 02
I= 93	Q.2311558	4.326	1.855888 06	1.413990 02
I= 94	Q.2336683	4.280	1.037821 06	1.300888 02
I= 95	Q.2361809	4.238	7.929751 05	1.158728 02
I= 96	Q.2386935	4.189	6.823576 05	1.078016 02
I= 97	Q.2412060	4.188	5.858338 05	1.770589 02
I= 98	Q.2437186	4.193	8.811658 05	1.533793 02
I= 99	Q.2462312	4.061	8.378171 05	1.270688 02
I= 100	Q.2487437	4.020	8.088019 05	1.986138 02
I= 101	Q.2512563	3.980	3.488668 05	1.688827 02
I= 102	Q.2537688	3.981	3.789715 05	1.390168 02
I= 103	Q.2562814	3.982	3.657135 05	1.110153 02
I= 104	Q.2587940	3.868	1.590666 06	2.878788 02
I= 105	Q.2613065	3.827	3.538827 06	2.717028 02
I= 106	Q.2638191	3.790	3.895010 06	2.670275 02
I= 107	Q.2663317	3.755	3.857718 06	2.755939 02
I= 108	Q.2688442	3.720	3.810539 06	2.969109 02
I= 109	Q.2713568	3.685	3.821806 06	1.289731 02
I= 110	Q.2738693	3.651	3.888159 06	1.685988 02
I= 111	Q.2763819	3.618	3.518888 06	1.127789 02
I= 112	Q.2788945	3.586	3.657561 06	1.589907 02
I= 113	Q.2814070	3.558	3.907257 06	1.051618 02
I= 114	Q.2839196	3.522	8.120682 05	1.893985 02
I= 115	Q.2864322	3.491	4.999548 06	1.901926 02
I= 116	Q.2889447	3.461	6.187799 06	1.278597 02
I= 117	Q.2914573	3.431	8.229610 06	1.595312 02
I= 118	Q.2939698	3.402	1.251388 06	1.889151 02
I= 119	Q.2964828	3.371	2.358718 06	7.036410 02
I= 120	Q.2989950	3.345	6.712618 06	7.150280 02
I= 121	Q.3015075	3.317	2.583667 07	7.186610 02
I= 122	Q.3040201	3.289	5.991710 06	7.188159 02
I= 123	Q.3065327	3.262	1.687289 06	7.023168 02
I= 124	Q.3090452	3.236	6.951178 05	6.828418 02

Fig. 6.1 (continued).

I = 125	0.1115578	1.210	3.6779060	05	6.5502150	02
I = 126	0.1100704	1.194	2.2197210	05	6.2051150	02
I = 127	0.1165829	1.159	1.4639100	05	5.7947100	02
I = 128	0.1190955	1.134	1.0299260	05	5.3254970	02
I = 129	0.1215090	1.109	0.6156640	06	4.8067070	02
I = 130	0.1240296	1.085	0.3670750	06	4.2462840	02
I = 131	0.1266332	1.062	0.2174790	06	3.6547050	02
I = 132	0.1293457	1.040	0.1354285	06	3.0430320	02
I = 133	0.1321653	1.018	0.0799020	06	2.4236060	02
I = 134	0.1350970	0.992	0.0428635	06	1.8112950	02
I = 135	0.1381438	0.970	0.0249048	06	1.2206170	02
I = 136	0.1413050	0.948	0.0148397	06	0.6320750	01
I = 137	0.1445805	0.926	0.0071026	06	0.1664970	01
I = 138	0.1479721	0.905	0.0035595	06	0.0044240	01
I = 139	0.1514817	0.884	0.0017695	06	1.2210650	02
I = 140	0.1551102	0.863	0.0008372	06	1.6250500	02
I = 141	0.1588587	0.843	0.0004330	06	1.9847170	02
I = 142	0.1627272	0.823	0.0002290	06	2.2969290	02
I = 143	0.1667157	0.803	0.0001265	06	2.5240670	02
I = 144	0.1708242	0.783	0.0000633	06	2.6955160	02
I = 145	0.1750527	0.764	0.0000300	06	2.7977990	02
I = 146	0.1794012	0.745	0.0000139	06	2.8325400	02
I = 147	0.1838697	0.726	0.0000060	06	2.9024470	02
I = 148	0.1884582	0.707	0.0000020	06	2.7120100	02
I = 149	0.1931667	0.689	0.0000000	06	2.5676090	02
I = 150	0.1980052	0.671	0.0000000	06	2.3777590	02
I = 151	0.2029737	0.653	0.0000000	06	2.1536370	02
I = 152	0.2080722	0.636	0.0000000	06	1.9101370	02
I = 153	0.2133007	0.619	0.0000000	06	1.6477750	02
I = 154	0.2186592	0.603	0.0000000	06	1.3554160	02
I = 155	0.2241477	0.588	0.0000000	06	1.0311010	02
I = 156	0.2297662	0.568	0.0000000	06	0.6779470	02
I = 157	0.2355147	0.551	0.0000000	06	0.3094300	02
I = 158	0.2413932	0.535	0.0000000	06	0.0518930	02
I = 159	0.2474017	0.519	0.0000000	06	1.0075300	02
I = 160	0.2535402	0.503	0.0000000	06	1.9811520	02
I = 161	0.2598087	0.488	0.0000000	06	2.2164610	02
I = 162	0.2662072	0.472	0.0000000	06	2.4308440	02
I = 163	0.2727357	0.457	0.0000000	06	2.6119670	02
I = 164	0.2793942	0.442	0.0000000	06	2.7584910	02
I = 165	0.2861827	0.427	0.0000000	06	2.8604170	02
I = 166	0.2931012	0.412	0.0000000	06	2.9184730	02
I = 167	0.3001497	0.398	0.0000000	06	2.9324470	02
I = 168	0.3073282	0.383	0.0000000	06	2.9052390	02
I = 169	0.3146267	0.369	0.0000000	06	2.8318770	02
I = 170	0.3220452	0.355	0.0000000	06	2.7097450	02
I = 171	0.3295837	0.341	0.0000000	06	2.5318770	02
I = 172	0.3372422	0.327	0.0000000	06	2.3029550	02
I = 173	0.3450207	0.314	0.0000000	06	2.0182900	02
I = 174	0.3529192	0.301	0.0000000	06	1.6830690	02
I = 175	0.3609377	0.287	0.0000000	06	1.3144820	02
I = 176	0.3690762	0.274	0.0000000	06	0.9343630	02
I = 177	0.3773347	0.261	0.0000000	06	0.5514070	02
I = 178	0.3857132	0.249	0.0000000	06	0.1654790	02
I = 179	0.3942117	0.236	0.0000000	06	2.7963760	02
I = 180	0.4028302	0.223	0.0000000	06	3.0281010	02
I = 181	0.4115687	0.211	0.0000000	06	3.2754310	02
I = 182	0.4204272	0.199	0.0000000	06	3.5264040	02
I = 183	0.4294057	0.187	0.0000000	06	3.7682160	02
I = 184	0.4385042	0.175	0.0000000	06	3.9933260	02
I = 185	0.4477227	0.163	0.0000000	06	4.1915990	02
I = 186	0.4570612	0.151	0.0000000	06	4.3608270	02
I = 187	0.4665197	0.140	0.0000000	06	4.4964980	02
I = 188	0.4760982	0.128	0.0000000	06	4.5967070	02
I = 189	0.4857967	0.117	0.0000000	06	4.6610570	02
I = 190	0.4956152	0.106	0.0000000	06	4.6904020	02
I = 191	0.5055537	0.095	0.0000000	06	4.6884010	02
I = 192	0.5156122	0.084	0.0000000	06	4.6578650	02
I = 193	0.5257907	0.073	0.0000000	06	4.6004690	02
I = 194	0.5360892	0.062	0.0000000	06	4.5306530	02
I = 195	0.5465077	0.052	0.0000000	06	4.4557980	02
I = 196	0.5570462	0.041	0.0000000	06	4.3758760	02
I = 197	0.5677047	0.031	0.0000000	06	4.3029500	02
I = 198	0.5784832	0.020	0.0000000	06	4.2405110	02
I = 199	0.5893817	0.010	0.0000000	06	4.2067230	02
I = 200	0.6004002	0.000	0.0000000	06	4.1936510	02

Fig. 6.2. (continued).

INTEGRATED RADIANT SPECTRUM

WIDTH OF INVESTIGATION WINDOW = 2.5000-C2

WAVELENGTH (Å)	FREQUENCY	PERIOD	REL. INT. SPEC.	INTEGRATED SPECTRUM	WAVELENGTH (Å)
1	0.0	0.0	0.00000000	0.0	0.0
2	0.0025126	398.099	0.00000000	0.0	0.0
3	0.0050251	199.049	0.00000000	0.0	0.0
4	0.0075377	132.667	0.00000000	0.0	0.0
5	0.0100503	99.500	0.00000000	0.0	0.0
6	0.0125628	79.490	0.00000000	0.0	0.0
7	0.0150754	66.333	0.00000000	0.0	0.0
8	0.0175879	56.857	0.00000000	0.0	0.0
9	0.0201005	49.700	0.00000000	0.0	0.0
10	0.0226131	44.222	0.00000000	0.0	0.0
11	0.0251256	39.800	0.00000000	0.0	0.0
12	0.0276382	36.182	0.00000000	0.0	0.0
13	0.0301508	33.167	0.00000000	0.0	0.0
14	0.0326633	30.615	0.00000000	0.0	0.0
15	0.0351759	28.029	0.00000000	0.0	0.0
16	0.0376884	26.511	0.00000000	0.0	0.0
17	0.0402010	24.075	0.00000000	0.0	0.0
18	0.0427136	22.012	0.00000000	0.0	0.0
19	0.0452261	20.211	0.00000000	0.0	0.0
20	0.0477387	18.607	0.00000000	0.0	0.0
21	0.0502513	17.190	0.00000000	0.0	0.0
22	0.0527638	15.952	0.00000000	0.0	0.0
23	0.0552764	14.891	0.00000000	0.0	0.0
24	0.0577889	13.906	0.00000000	0.0	0.0
25	0.0603015	13.003	0.00000000	0.0	0.0
26	0.0628141	12.179	0.00000000	0.0	0.0
27	0.0653266	11.429	0.00000000	0.0	0.0
28	0.0678392	10.748	0.00000000	0.0	0.0
29	0.0703518	10.130	0.00000000	0.0	0.0
30	0.0728643	9.572	0.00000000	0.0	0.0
31	0.0753769	9.070	0.00000000	0.0	0.0
32	0.0778894	8.622	0.00000000	0.0	0.0
33	0.0804020	8.226	0.00000000	0.0	0.0
34	0.0829146	7.881	0.00000000	0.0	0.0
35	0.0854271	7.586	0.00000000	0.0	0.0
36	0.0879397	7.340	0.00000000	0.0	0.0
37	0.0904523	7.142	0.00000000	0.0	0.0
38	0.0929648	6.981	0.00000000	0.0	0.0
39	0.0954774	6.856	0.00000000	0.0	0.0
40	0.0979899	6.764	0.00000000	0.0	0.0
41	0.1005025	6.699	0.00000000	0.0	0.0
42	0.1030151	6.657	0.00000000	0.0	0.0
43	0.1055276	6.633	0.00000000	0.0	0.0
44	0.1080402	6.622	0.00000000	0.0	0.0
45	0.1105528	6.620	0.00000000	0.0	0.0
46	0.1130653	6.625	0.00000000	0.0	0.0
47	0.1155779	6.634	0.00000000	0.0	0.0
48	0.1180905	6.646	0.00000000	0.0	0.0
49	0.1206030	6.662	0.00000000	0.0	0.0
50	0.1231156	6.682	0.00000000	0.0	0.0
51	0.1256281	6.706	0.00000000	0.0	0.0
52	0.1281407	6.734	0.00000000	0.0	0.0
53	0.1306533	6.766	0.00000000	0.0	0.0
54	0.1331658	6.800	0.00000000	0.0	0.0
55	0.1356784	6.836	0.00000000	0.0	0.0
56	0.1381910	6.874	0.00000000	0.0	0.0
57	0.1407035	6.913	0.00000000	0.0	0.0
58	0.1432161	6.952	0.00000000	0.0	0.0
59	0.1457286	6.992	0.00000000	0.0	0.0
60	0.1482412	7.032	0.00000000	0.0	0.0
61	0.1507538	7.072	0.00000000	0.0	0.0
62	0.1532663	7.112	0.00000000	0.0	0.0
63	0.1557789	7.152	0.00000000	0.0	0.0
64	0.1582915	7.192	0.00000000	0.0	0.0
65	0.1608040	7.232	0.00000000	0.0	0.0
66	0.1633166	7.272	0.00000000	0.0	0.0
67	0.1658291	7.312	0.00000000	0.0	0.0
68	0.1683417	7.352	0.00000000	0.0	0.0
69	0.1708543	7.392	0.00000000	0.0	0.0
70	0.1733668	7.432	0.00000000	0.0	0.0
71	0.1758794	7.472	0.00000000	0.0	0.0
72	0.1783920	7.512	0.00000000	0.0	0.0
73	0.1809045	7.552	0.00000000	0.0	0.0
74	0.1834171	7.592	0.00000000	0.0	0.0
75	0.1859296	7.632	0.00000000	0.0	0.0
76	0.1884422	7.672	0.00000000	0.0	0.0
77	0.1909548	7.712	0.00000000	0.0	0.0
78	0.1934673	7.752	0.00000000	0.0	0.0
79	0.1959799	7.792	0.00000000	0.0	0.0
80	0.1984925	7.832	0.00000000	0.0	0.0
81	0.2010050	7.872	0.00000000	0.0	0.0
82	0.2035176	7.912	0.00000000	0.0	0.0
83	0.2060302	7.952	0.00000000	0.0	0.0

Fig. 6.1. (continued).

164	0.2085827	0.795	0.1269780	05	2.0774900	04	1.0590570	01
165	0.2110553	0.794	0.17415630	05	3.1304940	04	1.5020340	01
166	0.2135670	0.682	0.5308450	05	4.1257180	04	2.7934430	00
167	0.2160804	0.629	1.3670390	04	0.7765830	04	1.3699170	01
168	0.2185530	0.575	2.3407220	04	5.1613150	04	0.0200010	00
169	0.2211055	0.523	3.0613020	06	5.0010930	04	2.1104590	01
170	0.2236191	0.472	4.5857950	06	5.0971090	04	2.9253030	00
171	0.2261307	0.422	5.8500010	06	5.0892290	04	2.6131030	01
172	0.2286432	0.370	7.2030560	06	5.3628590	04	1.2225470	00
173	0.2311558	0.326	8.6550090	06	5.0322750	04	2.6200220	00
174	0.2336683	0.280	1.0170210	06	4.3600700	04	1.5016500	01
175	0.2361809	0.230	7.9297510	05	3.3612100	04	0.0250000	00
176	0.2386935	0.180	6.0235750	05	2.9309720	04	1.0561020	00
177	0.2412060	0.146	5.0543390	05	1.0027080	04	6.3900760-01	
178	0.2437186	0.113	0.9116500	05	1.0095490	04	2.0171770	00
179	0.2462312	0.081	0.1701710	05	1.2751930	04	9.0064730-01	
180	0.2487437	0.050	0.0000190	05	1.1109170	04	0.0055120-01	
181	0.2512563	0.000	3.0004600	05	1.0093820	04	2.5770050-01	
182	0.2537688	0.001	3.7047100	05	9.0097610	03	1.4006670-01	
183	0.2562814	0.002	3.6571300	05	9.0747120	03	0.0095000	00
184	0.2587940	0.004	3.5966600	05	9.1000380	03	2.0544940	00
185	0.2613065	0.007	3.5308270	05	0.0095750	03	2.2564910	00
186	0.2638191	0.009	3.4550100	05	9.0639330	03	2.1004790	00
187	0.2663317	0.005	3.3751100	05	0.0051200	03	2.3813930	00
188	0.2688442	0.002	3.2953900	05	0.2219760	03	3.1520050	00
189	0.2713568	0.004	3.2100000	05	0.9390000	03	1.3264620-01	
190	0.2738693	0.001	3.0001500	05	0.2030700	03	2.5100990-01	
191	0.2763819	0.001	3.5160000	05	4.7003000	03	5.5535200-01	
192	0.2788945	0.006	3.6575610	05	1.0592810	04	1.4002900	00
193	0.2814070	0.004	3.9072570	05	1.2216060	04	3.1117690	00
194	0.2839196	0.002	4.3205020	05	1.5511000	04	6.2104670	00
195	0.2864322	0.001	4.9995600	05	2.0123050	04	0.1827250	00
196	0.2889447	0.001	6.1477900	05	5.0075770	04	3.1014010	00
197	0.2914573	0.001	6.2294630	05	1.0060350	05	9.0359230	01
198	0.2939698	0.002	1.2513000	06	1.0744560	05	6.5096610	01
199	0.2964824	0.003	2.3507100	06	1.0064650	05	0.9000030	00
200	0.2989950	0.003	6.7126100	06	1.0992390	05	2.6941560	01
201	0.3015075	0.003	2.5036670	07	1.2915000	05	3.2049510	01
202	0.3040201	0.009	5.0017100	06	1.0702000	05	6.0042120	01
203	0.3065327	0.002	1.6022000	06	1.0547320	05	5.3043370	01
204	0.3090452	0.002	6.0551700	05	1.0100330	05	2.0946230	01
205	0.3115578	0.001	3.4779000	06	9.1371400	04	1.8047200	01
206	0.3140704	0.001	2.2197210	05	5.1355730	04	0.8641170	00
207	0.3165829	0.004	1.0639100	06	1.3707600	04	2.1305500	00
208	0.3190954	0.001	1.7299260	05	6.0700920	03	2.2025570	00
209	0.3216080	0.009	7.6176600	04	3.5010270	03	0.1195030-01	
210	0.3241206	0.004	5.0070750	04	2.3102500	03	1.1703990-01	
211	0.3266332	0.002	0.4707960	04	1.6625200	03	0.5213990-01	
212	0.3291457	0.001	1.0350200	04	1.2667120	03	1.8004050-01	
213	0.3316583	0.004	3.2290820	04	1.0095000	03	0.7900020-02	
214	0.3341709	0.002	2.7024630	04	0.1012710	02	0.0001330-02	
215	0.3366835	0.000	2.0090000	04	7.1107070	02	2.0610020-02	
216	0.3391960	0.000	2.1003970	04	6.2241100	02	1.9602020-01	
217	0.3417085	0.002	2.0101200	04	5.5022010	02	1.2295700-01	
218	0.3442211	0.005	1.0703500	04	5.1169920	02	0.0102310-02	
219	0.3467337	0.000	1.7095470	04	0.7071120	02	5.0505090-02	
220	0.3492462	0.003	1.7013370	04	0.5651500	02	3.0710060-02	
221	0.3517588	0.003	1.6613930	04	0.0332950	02	2.0572220-02	
222	0.3542714	0.003	1.6668570	04	0.3003500	02	2.1625770-02	
223	0.3567839	0.007	1.6566550	04	0.3999910	02	3.2200960-03	
224	0.3592965	0.003	1.6900330	04	0.9992250	02	1.0067010-02	
225	0.3618090	0.006	1.7000000	04	4.6075660	02	3.1750030-03	
226	0.3643216	0.005	1.0333100	04	0.0750260	02	2.2065120-01	
227	0.3668342	0.004	1.0059000	04	5.1700000	02	2.2095110-02	
228	0.3693467	0.007	2.0090000	04	5.0000000	02	0.0200200-02	
229	0.3718593	0.000	2.2000000	04	5.9702340	02	1.0050700-01	
230	0.3743719	0.001	2.0702320	04	4.0670560	02	1.2515370-01	
231	0.3768845	0.005	2.1606000	04	6.0055130	02	4.0210020-02	
232	0.3793970	0.006	2.9000000	04	7.5370000	02	1.6762700-01	
233	0.3819095	0.009	3.2077260	04	0.0660300	02	1.3250200-01	
234	0.3844221	0.001	1.5061300	04	0.5552570	02	2.0902120-02	
235	0.3869347	0.000	1.7903000	04	0.0707910	02	2.7070700-01	
236	0.3894472	0.000	3.0000000	04	0.3109100	02	3.5011000-02	
237	0.3919598	0.001	0.1017210	04	0.5650260	02	3.0560900-02	
238	0.3944724	0.005	0.1350000	04	0.7120000	02	5.9201100-02	
239	0.3969849	0.009	0.0970150	04	1.7665100	02	2.6017920-01	
240	0.3994975	0.001	0.7071010	04	0.7103900	02	3.0326370-01	
241	0.4020101	0.000	1.0093020	04	0.6531010	02	9.6926000-02	
242	0.4045226	0.002	1.7717750	04	0.5375300	02	6.1709300-02	
243	0.4070352	0.007	1.6079150	04	0.0270160	02	2.2000990-02	
244	0.4095477	0.002	1.5029610	04	0.3599100	02	2.0071800-02	
245	0.4120603	0.007	1.5567630	04	0.3719150	02	3.7300010-01	
246	0.4145729	0.002	1.5672500	04	0.5000900	02	3.1655570-02	
247	0.4170854	0.000	1.4322700	04	0.7031070	02	0.0279000-02	
248	0.4195980	0.003	1.7612270	04	1.0267220	03	1.5202170-01	
249	0.4221106	0.009	1.0000000	04	1.1010000	03	2.5051100-01	
250	0.4246231	0.005	0.7070700	04	1.2115000	03	0.3000000-01	
251	0.4271357	0.001	0.6912000	04	1.1710730	03	5.7015200-02	
252	0.4296482	0.007	0.2700230	04	1.6030070	03	1.0796000-01	
253	0.4321608	0.001	0.0911110	04	1.0001000	03	2.1000200-01	
254	0.4346734	0.001	7.2295000	04	2.0707000	03	0.3321500-01	
255	0.4371859	0.007	0.0505670	04	1.3000000	03	0.0525000-01	
256	0.4396985	0.002	1.1231000	04	0.6001000	03	9.2951000-01	
257	0.4422111	0.000	1.0000100	05	7.0055700	03	2.0020100	00

I= 178	0.0007236	2.209	2.0609300 05	1.0615260 00	2.0002900 00
I= 179	0.0072162	2.234	2.0202100 05	1.0000000 00	2.2698000 00
I= 180	0.0097007	2.223	0.7013560 04	1.0905200 00	0.9610700-01
I= 181	0.0522613	2.211	7.7016500 05	2.0017710 00	0.4751370-01
I= 182	0.0507739	2.199	1.2060000 04	2.2006000 00	3.0027900 00
I= 183	0.0572860	2.187	1.6772260 06	2.3109600 00	1.7279520 01
I= 184	0.0597000	2.175	1.3576500 06	2.3025000 00	0.0209100 00
I= 185	0.0623116	2.163	1.2090100 04	2.0005270 00	2.2610020 00
I= 186	0.0600201	2.151	0.8031370 05	2.3633150 00	1.0210500-01
I= 187	0.0673367	2.140	6.7009100 05	2.2200560 00	7.2515900-01
I= 188	0.0690092	2.128	5.5726970 05	1.9702710 00	6.1902900 00
I= 189	0.0723618	2.117	0.8036090 05	1.6067050 00	7.2902220 00
I= 190	0.0700700	2.106	0.5070970 05	1.5036000 00	2.3197020-01
I= 191	0.0773069	2.095	0.0032510 05	1.0007500 00	2.0236700 00
I= 192	0.0790995	2.080	0.6720060 05	1.5001360 00	3.0753010 00
I= 193	0.0820121	2.073	5.1500700 05	1.9230020 00	7.0093000-01
I= 194	0.0809206	2.062	6.0009050 05	2.6019900 00	1.0001010 01
I= 195	0.0870372	2.052	7.5776390 05	3.0750000 00	2.0906720 01
I= 196	0.0890097	2.041	1.0292220 06	0.0	0.0
I= 197	0.0920623	2.031	1.5373010 06	0.0	0.0
I= 198	0.0900709	2.020	2.5513110 06	0.0	0.0
I= 199	0.0970070	2.010	0.0525220 06	0.0	0.0
I= 200	0.5000000	2.000	5.9909250 06	0.0	0.0

FINAL PREDICTION ERROR

I	NUMBER OF FILTER COEFFICIENTS	FINAL PREDICTION ERROR
1	2	1.363000 06
2	3	1.067919 06
3	4	1.575300 06
4	5	1.731670 06
5	6	1.063570 06
6	7	1.103210 06
7	8	1.250590 06
8	9	1.372930 06
9	10	1.143670 06
10	11	1.207900 06
11	12	1.310390 06
12	13	1.515210 06
13	14	1.007330 06
14	15	2.213000 06
15	16	2.776170 06
16	17	3.700050 06
17	18	5.575000 06
18	19	1.103030 07

Fig. 6.1. (continued).

ORNL-DWG 78-5878

A. Hudson River catch per unit effort data; dashed line represents the linear trend (Source for the time series: Texas Instruments, Inc., 1977, p. 76, Table 7.2-1).

ORNL-DWG 78-5879

B. Detrended Hudson River catch per unit effort; solid line represents the axis $y = 0$.

Fig. 6.2. Analysis of the Hudson River catch per unit effort data.

ORNL-DWG 78-5880

C. Residuals versus fitted values.

ORNL-DWG 78-5881

D. Yule - Walker autocorrelation function.

Fig. 6.2. (continued).

E. Periodogram.

F. Cumulative periodogram.

Fig. 6.2. (continued).

G. Fourier amplitude spectrum.

H. Cumulative Fourier power spectrum.

I. Burg autocorrelation estimate.

J. Maximum entropy spectrum with $M = 18$.

K. Integrated maximum entropy spectrum.

L. Final prediction error

8. Setting up the Program for Execution

The time it takes to run the program depends on many factors, primarily the length of the time series, the number of frequencies used, the number of times MESA is performed and the use of other options. Table 6.2 lists a few examples indicating the time and core used. Although the list is short, it is hoped that it will serve as a guide in the classification of a program run. All the data were linearly detrended.

For users at ORNL, Table 6.3 is a typical setup of the job control cards for a Class C job with disks used for plotting. Similarly, Table 6.4 is a setup for a class E job with the use of tapes in plotting. The job initials BLKL (wherever they appear) should be changed to that of the user's.

The inclusion of the following cards is explained:

1. `///DDDSN = SYS2.DISSPLA,DISP = SHR .....` "gives access to the plotting routines.
2. `///DDDSNAME = BROOKS.LOAD.MODULES .....` "gives access to the BLSQ.
3. `///DDDSNAME = JDAIMSL1.REF 10276 .....` "gives access to DCADRE.

Table 6.2. List of programs runs; time and core used.

Number of points in time series	IPRINT	IPLOT	ITREND	IFPEPL	IYWACV	IINTSP	IAUTCV	Number of frequencies	Mcount values	Logarithmic transform analyzed	Computer time and core used
21	yes	yes	yes	no	yes	no	no	600	8-17	yes	1.04 minutes 286 K
21	yes	yes	yes	yes	no	no	no	600	19	yes	23.90 seconds 286 K
31	yes	yes	yes	no	yes	yes	no	501	15,16	no	41.88 seconds 280 K
42	yes	yes	yes	no	no	no	no	600	40	yes	28.37 seconds 320 K
70	yes	yes	yes	no	no	no	no	201	30	no	11.35 seconds 280 K
114	yes	yes	yes	no	no	no	no	600	69	no	21.26 seconds 286 K
260	yes	yes	yes	no	no	no	no	600	109 107	no	35.46 seconds 290 K
277	yes	yes	yes	no	no	no	no	600	160	yes	49.78 seconds 320 K
277	yes	yes	yes	yes	no	no	no	600	275	yes	1.06 minutes 320 K

Table 6.3. Job setup for disk plotting

```

//BLKL JOB (16064),'ENV_2001_X-10_KIRK'
//*CLASS_CPU91=5M,REGION=540
//_EXEC_FORTHCLG,PARM.FORT='XREF',
//_PARM.GO='TIME=5.0,EU=-1,DUMP=-1,S0=51',REGION.GO=540K
//FORT.SYSIN_DD_*
.
.
.
source deck
.
.
.
/*
//LKED.SYSLIB_DD
//_DD
//_DD
//_DD_DSN = SYS2.DISSPLA,DISP=SHR
//_DD_DSNAME=BROOKS.LOAD.MODULES,UNIT=2314,
//_VOLUME=SER=ZZZZZZ,DISP=SHR
//_DD_DSNAME=JDAIMSL1.REF10276,UNIT=2314,VOL=SER=CADPK1,DISP=SHR
//LKED.PLOTSUBS_DD_DSN=JGSPLOTH,DISP=SHR
//LKED.SYSIN_DD_*
__INCLUDE_PLOTSUBS
/*

```

Table 6.3. (continued)

```
//GO.F749F001_DD_UNIT=IN20U2,DISP=(NEW,KEEP),
```

```
//_SPACE=(3208,40,RLSE),DSN=PLOT00.BKL,
```

```
//_DCB=(RECFM=VS,LRECL=3204,BLKSIZE=3208)
```

```
//GO.FT05F001_DD_*
```

```
.
```

```
.
```

```
.
```

data cards

```
.
```

```
.
```

```
.
```

```
/*
```

```
//
```

Table 6.4. Job setup for tape plotting

```

/**NOSEQCARD
//BLKL_JOB_(16064),'ENV_2001_X-10_KIRK'
/**CLASS_CPU91=5M,REGION=540,LINES=40,CARDS=10,SPECIAL=TAPE
/**PLOT__NUMBER=30
//_EXEC_FOR:HCLG,PARM.FORT='XREF',
//_PARM.GO='TIME=5.0,CU=-1,DUMP=1,SO=51',REGION.GO=540K
//FORT.SYSIN_DD *
.
.
.
source deck
.
.
.
/*
//LKED.SYSLIB_DD
// DD
// DD
//_DD_DSNAME=SYS2.DISSPLA,DISP=SHR
//_DD_DSNAME=BROOKS.LOAD.MODULES,UNIT=2314,
//_VOLUME=SER=ZZZZZZ,DISP=SHR
//_DD_DSNAME=JDAI:MSL1.REF10276,UNIT=2314,VOL=SER=CADPK1,DISP=SHR
//GO.PLOTTAPE_DD_INIT=TAPE7,LABEL=(,NL),DISP=OLD,VOLUME=SER=15
//GO.FT05F001_DD_

```

Table 6.4. (continued)

.
.
.
data cards
.
.
.
/*
//

REFERENCES

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APPENDICES

APPENDIX I

DESCRIPTION OF FUNCTION DCADRE*

FUNCTION DCADRE (F,A,B,AERR,RERR,ERROR,IER)

DCADRE-----S/D-----LIBRARY 1-----

FUNCTION	- INTEGRATE F(X) FROM A TO B, USING CAUTIOUS ADAPTIVE ROMBERG EXTRAPOLATION.
USAGE	- FUNCTION DCADRE (F,A,B,AERR,RERR,ERROR,IER)
PARAMETERS	DCADRE - ESTIMATE OF THE INTEGRAL OF F(X) FROM A TO B.
	F - A SINGLE-ARGUMENT REAL FUNCTION SUBPROGRAM SUPPLIED BY THE USER. F MUST BE DECLARED EXTERNAL IN THE CALLING PROGRAM.
	A,B - THE TWO ENDPOINTS OF THE INTERVAL OF INTEGRATION. (INPUT)
	AERR - DESIRED ABSOLUTE ERROR IN THE ANSWER. (INPUT)
	RERR - DESIRED RELATIVE ERROR IN THE ANSWER. (INPUT)
	ERROR - ESTIMATED BOUND ON THE ABSOLUTE ERROR OF THE OUTPUT NUMBER, DCADRE.
	IER - ERROR PARAMETER
	WARNING ERROR(WITH FIX) = 64 + N
	N = 1 IMPLIES THAT ONE OR MORE SINGULARITIES WERE SUCCESSFULLY HANDLED.
	N = 2 IMPLIES THAT, IN SOME SUBINTERVAL(S), THE ESTIMATE OF THE INTEGRAL WAS ACCEPTED MERELY BECAUSE THE ESTIMATED ERROR WAS SMALL, EVEN THOUGH NO REGULAR BEHAVIOR WAS RECOGNIZED.
	TERMINAL ERROR = 128 + N
	N = 3 -- FAILURE DUE TO INSUFFICIENT INTERNAL WORKING STORAGE.
	N = 4 -- FAILURE. THIS MAY BE DUE TO TOO MUCH NOISE IN THE FUNCTION (RELATIVE TO THE GIVEN ERROR REQUIREMENTS) OR DUE TO AN ILL-BEHAVED INTEGRAND.
	N = 5 INDICATES THAT RERR IS GREATER THAN 0.1, OR RERR IS LESS THAN 0.0, OR RERR IS TOO SMALL FOR THE PRECISION OF THE MACHINE.
PRECISION	- SINGLE/DOUBLE
REQD. IMSL ROUTINES	- UERTST
LANGUAGE	- FORTRAN

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*Reprinted from IMSL LIBRARY 1. Reference Manual. Edition 4 (Fortran IV) S/370 - 360, 1975.

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FUNCTION DCADRE(F,A,B,AERR,RERR,ERROR,IER)

Purpose

DCADRE attempts to solve the following problem: Given the name F of a real function subprogram, two real numbers A and B, and two non-negative numbers AERR and RERR, find a number DCADRE such that

$$\left| \int_A^B F(x)dx - \text{DCADRE} \right| \leq \max \left(\text{AERR}, \text{RERR} * \left| \int_A^B F(x)dx \right| \right)$$

Algorithm

This routine uses a scheme whereby DCADRE is computed as the sum of estimates for the integral of F(x) over suitably chosen subintervals of the given interval of integration. Starting with the interval of integration itself as the first such subinterval, cautious Romberg extrapolation is used to find an acceptable estimate on a given subinterval. If this attempt fails, the subinterval is divided into two subintervals of equal length, each of which is considered separately.

See reference: de Boor, Carl, "CADRE: An algorithm for numerical quadrature", Mathematical Software (John R. Rice, Ed.), New York, Academic Press, 1971, Chapter 7.

Programming Notes

1. DCADRE can, in many cases, handle jump discontinuities and certain algebraic discontinuities. See reference for full details.
2. The relative error parameter RERR must be in the interval [0,0.1]. For example, RERR=0.1 indicates that the estimate of the integral is to be correct to one digit, whereas RERR=10 calls for four digits of accuracy. If DCADRE determines that the relative accuracy requirement cannot be satisfied, IER is set to 133 (RERR should be large enough that, when added to 100.0, the result is a number greater than 100.0).
3. The absolute error parameter, AERR, should be nonnegative. In order to give a reasonable value for AERR, the user must know the approximate magnitude of the integral being computed. In many cases it is satisfactory to use AERR=0. In this case, only the relative error requirement is satisfied in the computation.
4. We quote from the reference, "A very cautious man would accept DCADRE only if IER is 0 or 65. The merely reasonable man would keep the faith even if IER is 66. The adventurous man is quite often right in accepting DCADRE even if IER is 131 or 132". Even when IER≠0, DCADRE returns the best estimate that has been computed.

APPENDIX II

DESCRIPTION OF SUBROUTINE BLSQ*

Identification

Linear least squares solution.

Purpose

ALSQ and BLSQ are double precision FORTRAN IV subroutines to solve the linear least squares problem by applying the Householder reduction to the least squares matrix [1]. Both programs use the same technique in solving the least squares problem; however, BLSQ returns additional information useful in statistical applications. Both programs have subordinate entries, ALSQ1, BLSQ1, which enable the user to perform additional fits at very little cost once the least squares matrix has been reduced.

Method

The linear least squares problem to be solved may be formulated as follows: given $n \times m$ ($m < n$) matrix A of rank m and an n -vector y find an m -vector b such that

$$\|Ab - y\|^2 = \min. \quad (1)$$

The programs, in effect, use the Householder reduction of the matrix A to triangular form to find an $n \times n$ orthogonal matrix H such that

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*Reprinted from Westley and Watts, The Computing Technology Center
Numerical Analysis Library. CTC - 39. (Oct. 6, 1970) pp. 383-386.

$$HA = \begin{bmatrix} T \\ 0 \end{bmatrix}, \quad (2)$$

where T is an $m \times m$ upper triangular matrix and "0" is an $(n - m) \times m$ null matrix. If we let U be the $n \times m$ matrix consisting of the first m columns of H^T (in order) and V be the $n \times (n - m)$ matrix consisting of the remaining columns and consider the new least squares problem of minimizing $\|Uc - y\|^2$, then

$$c = U^T y \quad (3)$$

and

$$b = T^{-1}c. \quad (4)$$

ALSQ performs the Householder reduction (2) and uses equations (3) and (4) to calculate b . BLSQ calculates and returns the vector c and the matrices T and T^{-1} which are useful in statistical applications. The vector b may subsequently be calculated by the user from equation (4). Both programs return the approximating vector $Ab = Uc$ and the residual sum of squares $\|Ab - y\|^2$.

Under the usual statistical assumptions of regression analyses, it can be shown that the number of runs in $V^T y$ minus one has the binomial distribution with parameters $n - m - 1$ and .5. If the problem being solved involves fitting continuous functions of a real variable, a statistical test based on the above may be expected to be powerful when the

fit inadequately represents the data. Because of this, BLSQ returns the number of runs in $V^T y$.

Usage

The calling sequence for ALSQ is:

```
CALL ALSQ(A, Y, B, R2, N, M, NA),
```

where

- A is a doubly subscripted double precision array containing the least squares matrix. The elements of A are altered by the program.
- Y is a singly subscripted, double precision array containing the vector to be fit. The elements of Y are unchanged by the program.
- B is a singly subscripted, double precision array which upon return contains the coefficients of the fit.
- R2 contains, upon return, the double precision residual sum of squares of the fit.
- N is the integer number of rows in the least squares matrix.
- M is the integer number of columns in the least squares matrix.
- NA is the integer first dimension of the array A.

The program requires that the array A be dimensioned so that it will hold one extra row and one extra column in addition to the least squares matrix, i.e., the dimension of A must be at least $N + 1$ by

$M + 1$. Upon return from the program, the locations $A(I, M + 1)$ contain the components of the approximating vector Ab (see Method).

A second entry, ALSQ1, allows the user to obtain additional fits without repeating the reduction of the least squares matrix. ALSQ must be called before calling ALSQ1, and the contents of the array A must not be changed between these calls.

The calling sequence is:

CALL ALSQ1(Y, B, R2, K)

where Y , B , and $R2$ are the same as above and

K is the number of columns of the least squares matrix to fit to the vector Y ; e.g., if $K = 4$, the first four columns of A will be fit to Y returning four coefficients in the array B . K must be less than or equal to M in the call ALSQ.

This entry also returns the approximating vector in the locations $A(I, M + 1)$.

The calling sequence for BLSQ is:

CALL BLSQ(A, Y, C, R2, IRUN, T, T1, N, M, NA, NT),

where A , Y , $R2$, N , M , and NA are the same as in the calling sequence for ALSQ and

C is a singly subscripted, double precision array which upon return contains the coefficients of the transformed problem given by (3).

IRUN is an integer array whose i^{th} element contains upon return the number of runs in $V^T y$ after i columns of the least squares matrix have been fitted.

T is a doubly subscripted, double precision array which upon return contains the upper triangular transformation matrix of equations (2).

T1 is a doubly subscripted, double precision array which upon return contains the inverse of the transformation matrix.

NT is the first dimension of the arrays **T** and **T1**.

BLSQ also has a second entry, **BLSQ1**, to calculate new fits. The same restrictions apply to the use of **BLSQ1** as apply to **ALSQ1**. The calling sequence is:

CALL **BLSQ1**(**Y**, **C**, **R2**, **IRUN**, **K**),

where **Y**, **C**, **R2**, **IRUN** are the same as in the calling sequence of **BLSQ** and **K** is the number of columns of the least squares matrix to fit to the data. Both **BLSQ** and **BLSQ1** return the components of the approximating vector in the locations **A(I, M + 1)**.

Coding Information

Although these programs are coded in double precision arithmetic, they may be easily converted to single precision. Additional accuracy may be obtained in the single precision program by accumulating inner products in double precision. The places where this should be done are signaled by the appearance of the variables **SS** and **PP**.

Reference

- (1) Golub, G. (1965): "Numerical methods for solving linear least squares problems," Numer. Math. 7, 206-216.

- (2) Wampler, R. H., "An Evaluation of Linear Least Squares Computer Programs," J. of Research of NBS, Vol. 73B, No. 2, p. 59.

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[Editor's Note: For an evaluation of ALSQ in single and double precision see (2). Unfortunately, the single precision version reported in (2) did not accumulate inner products in double precision.]

APPENDIX III
PROGRAM LISTING

05/360 PORTMAN H

```

C
ISN 0002      IMPLICIT REAL*8(A-H,O-Z)
ISN 0003      REAL*4 NAN, XLBL, YLBL
ISN 0004      COMMON /SPECBL/ DELTA, YEAR(600), Y(600), XD(600),
1 VMAX,X(600),V(600),S(600),FPE(600),IDIFF(600),
2 DATE(600),I(600),ACS(600),BSW(600),AMP(600),ASP(600),
3 PERIOD(600),NPTS,NPILT,NV,NCOUNT(600),YPRINT,IPLT,ITREND,
4 IPRPFL,IINTSP,IAUTCV,IYBACY
ISN 0005      COMMON /BLKLBL/NAN(14),XLBL(2),YLBL(2)

C
C          INITIALISE VALUES.
C
C          XLABEL = IDENTIFIER (LABEL) FOR TIME AXIS  (A9)
C          YLABEL = IDENTIFIER FOR TIME SERIES DATA  (A9)
C          DELTA   = SAMPLING TIME INTERVAL
C          NPTS    = NUMBER OF DATA POINTS IN TIME SERIES
C          NPILT   = TOTAL NUMBER OF TIMES MAXIMUM ENTROPY IS
C                   TO BE DONE WITH DIFFERENT VALUES FOR
C                   THE NUMBER OF FILTER COEFFICIENTS
C          NV      = NUMBER OF FREQUENCIES NEEDED IN OUTPUT
C          VHAX    = MAXIMUM FREQUENCY FOR OUTPUT SPECTRUM
C                   (SHOULD NOT EXCEED THE NYQUIST FREQUENCY).
C          NAR     = TITLE FOR THE DATA SET  (A80)
C                   (TITLE PRINTED AT TOP OF THE PLOTS.
C
C
ISN 0006      1 CONTINUE
C
ISN 0007      READ (5,20,HUB=500) FLBL,YLBL,DELTA,NPTS,NPILT,NV,VHAX
ISN 0008      20 FORMAT (2A4,2X,2A4,2X,E10.0,3I5,E10.0)
ISN 0009      READ (5,21,HUB=500) NAR
ISN 0010      21 FORMAT ('A80')

C
C          READ NCOUNT. NCOUNT IS THE ARRAY OF NUMBERS WHICH CORRESPOND TO THE
C          NUMBER OF FILTER COEFFICIENTS DESIRED IN THE MAXIMUM ENTROPY.
C
C
ISN 0011      READ 40, (NCOUNT(I),I=1,NPILT)
ISN 0012      40 FORMAT ('16I3')

```


***** C O R T E S R E F E R E N C E S L I S T I N G *****

L A S T	D E S I G N	R E F E R E N C E S
1	0004	0026
20	0004	0307
21	0010	0009
40	0012	0011
60	0015	0014
80	0014	0013
100	0018	0017
200	0025	0007 0009

LEVEL 21.8 (JUN 76)

03/360 PORTMAN 1

COMPILED OPTIONS - NAME= MAIN,OPT=02,LINENR=40,SIZE=0000K,
SOURCE,BBCDIC,NOLIST,NOBCK,LOAD,NAP,NOEDIT,ID,XREF

```

      C
ISN 0002      SUBROUTINE TRANS1
      C
ISN 0003      IMPLICIT REAL*8 (A-H,O-Z)
ISN 0004      REAL*8 NAN, XLBL, YLBL
      C
ISN 0005      COMMON /SPECL/ DEL' ',YEAR(600),Y(600),XD(600),
1 VARI,2(600),V(600),S(600),PFR(600),XDIFF(600),
2 DATE(600),I(600),ACS(600),NSH(600),ANF(600),ASP(600),
3 PERIOD(600),NPTS,NPILY,NV,NCOUNT(600),IPRINT,IPLOT,ITREND,
4 IPPRPL,IINTSP,IAUTCV,IYBACV
ISN 0006      COMMON /BLRBL/NAN(16),XLBL(2),YLBL(2)
      C
      C
ISN 0007      ILOG = 0
ISN 0008      PRINT 10
ISN 0009      10 FORMAT (1H1,'ANALYSIS OF THE ORIGINAL INPUT ',
1 'DATA WITHOUT ANY PRIOR TRANSFORMATIONS'/)
      C
ISN 0010      CALL SPECTR
      C
ISN 0011      RETURN
      C
ISN 0012      ENTRY TRANS2
ISN 0013      ILOG = 1
ISN 0014      PRINT 30
ISN 0015      30 FORMAT (1H1,'ANALYSIS OF THE LOGARITHM OF THE ORIGINAL DATA'//)
ISN 0016      DO 50 I=1,NPTS
ISN 0017      IF (Y(I)-EQ.0.) Y(I) = 1.
ISN 0018      Y(I) = DLOG(Y(I))
ISN 0019      50 CONTINUE
ISN 0020      CALL SPECTR
ISN 0021      RETURN
      C
ISN 0022      C
ISN 0023      ENTRY TRANS3
      C
ISN 0024      RETURN
      C
ISN 0025      ENTRY TRANS4
      C
ISN 0026      RETURN
      C
ISN 0027      ENTRY TRANS5
      C
ISN 0028      RETURN
      C
ISN 0029      END

```


*****
***** PORTLAND CROSS REFERENCE LISTING *****
*****

LABEL	DEFINED	REFERENCES
10	0009	0000
30	0015	0014
50	0020	0016

ISM 0002

SUBROUTINE SECTD

MAXIMUM ENTROPY SPECTRAL ANALYSIS SUBROUTINE

191

TO COMPUTE THE SPECTRA
IF(ITREND.EQ.0) USE RAW DATA

ITYWACV = YULE-WALKER AUTOCOVARANCE FLAG.
IF(ITYWACV.NE.0) CALCULATE, PRINT, AND PLOT THE
YULE-WALKER ESTIMATE OF THE
AUTOCOVARANCE FUNCTION
IF(ITYWACT.EQ.0) DON'T CALCULATE YULE-WALKER
ESTIMATE

NCOUNT = VECTOR OF FREQ. ERROR FILTER LENGTHS TO USE

NAME = TITLE FOR THE DATA SET (TITLE FOR THE PLOTS)
MAX. OF 80 CHARACTERS (10AB) ENDING WITH \$

NFILT = NO. OF PREDICTION ERROR FILTER LENGTHS TO TRY
(LENGTH OF PRED. ERROR FILTER = NUMBER OF LAGS)

NPER = NO. OF PERIODOGRAM ORDINATES. (NPER = GREATEST INTEGER
L.E. NYA/2).

NPTS = NO. OF DATA POINTS IN TIME SERIES.

NV = NO. OF FREQUENCIES AT WHICH SPECTRUM IS TO BE CALC.

PERDC = PERIODOGRAM VALUES.

PERIOD = VECTOR OF WORKING STORAGE (SAME LENGTH AS V)
UPON RETURN THIS VECTOR CONTAINS THE PERIODS
CORRESPONDING TO THE FREQUENCIES IN V.

PPER = PERIODOGRAM PERIODS IN CORRESPONDENCE WITH PPER.

S = THE MAXIMUM ENTROPY SPECTRUM. ONE ELEMENT
FOR EACH ELEMENT IN VECTOR V.
THIS ARRAY IS WRITTEN OVER FOR EACH NEW
VALUE OF NCOUNT. UPON RETURN TO CALLING
PROGRAM, IT WILL CONTAIN THE LAST SPECTRUM
COMPUTED. I.E., THE SPECTRUM CORRESPONDING
TO THE LAST LAG VALUE IN NCOUNT ARRAY.

STATE = LABEL FOR DEP. VARS. (TIME SERIES) AXIS
MAX. OF 8 CHARACTERS (AB) ENDING WITH \$

V = OUTPUT VECTOR OF FREQUENCIES AT WHICH SPECTRUM
IS CALC. (HEREIN GENERATED). SHOULD BE
DIMENSIONED TO HAVE AT LEAST (BP+1) ELEMENTS.

VMAX = MAX. FREQ. AT WHICH SPECTRUM IS TO BE CALC.

X = UPON RETURNING FROM DETREND IS THE TREND OF THE INPUT DATA.
IT IS LATER ON USED AS A WORKING VECTOR.

XD = IF DETRENDING IS DONE, IT IS THE EXTENDED DATA.
IF DETRENDING IS NOT DONE, IT IS THE ORIGINAL DATA.
HESA AND FOURIER ANALYSIS ARE PERFORMED ON XD.

XDIF = UPON RETURNING FROM DETREND IS THE DETRENDED DATA.

```

C      IT IS LATER USED AS A WORKING VECTOR.
C
C      Y      = INPUT TIME SERIES VALUES.
C
C      YEAR   = WORKING VECTOR (CAN BE USED FOR TIME VARIABLE
C              IN SUBROUTINE DETEND.)
C
C      YTIME  = LABEL FOR THE TIME AXIS
C              MAX. OF 8 CHARACTERS (AS) ENDING WITH S
C
C      Z      = PRED. ERROR FILTER COMPPS. (HEREIN CALC. & RETURNED)
C
C
C      IMPL/CIT REAL*8(A-H,0-2)
C      REAL*8 NTPREC
C      REAL*8 NAM,YTIME,STATE
C      COMMON /SPICBL/ DELTA, YEAR(60), Y(600), XD(600),
C      1 VMAX,Z(600),V(600),Z(600),PPE(600),XDIPP(600),
C      2 DATE(600),Y(600),ACS(600),RSU(600),AMP(600),ASFP(600),PERIOD(600),
C      3 NPTS,NPTS1,NV,NCOUNT(600),IPRINT,IPLDT,IPTEND,IPTDEL,IINTSP,
C      4 XAUSCV,IYRACY
C      COMMON /VRLBL/ NAM(14),YTIME(2),STATE(2)
C      COMMON /PRCGB/ PPER(300),PPER(300),PEROD(300),CUNPER(300)
C      EXTERNAL SECVAL
C      COMPLETS=16,CSUR,CY
C      DIMENSION ACFM(600),A(600),a(600)
C      DATA FL/Z*132#3P6A085A31/
C
C
C      INITIALIZE DISSEPLA ROUTINES FOR PLOTTING.
C
C
C      CALL COPY
C
C      PRINT INPUT DATA.
C
C
C      IF (IPRINT.EQ.0) GO TO 100
C      PRINT 10, NAM
C      10 FORMAT ('NO,NAME')
C      20 PRINT 40,YTIME,STATE
C      40 FORMAT ('NO,ZE,ZA9,IZI,ZA9')
C      PRINT 60
C      60 FORMAT ('N')
C      DO 80 I=1,NPTS
C      PRINT 70, DATE(I), Y(I)
C      70 FORMAT ('N',IPEL1,a,ZI,IPEI19.6)
C      80 CONTINUE
C      100 CONTINUE
C
C
C      DETEND ORIGINAL DATA BY CALLING SUBROUTINE DETEND.
C
C      SUBROUTINE DETEND MUST NOT DESTROY THE ARRAYS DATE AND CATCH.
C      IT SHOULD SET UP ITS OWN TIME ARRAY IN YEAR.
C

```

```

C
C
ISM 0027      CALL DETEND(NPTS,DATE,YEAR,Y,XDIFF,X,SPPE)
ISM 0028      DO 170 I=1,NPTS
ISM 0029      XD(I) = XDIFF(I)
ISM 0030      170 CONTINUE

C
C
C      PLOT DATA AND ITS TREND COMPONENT.
C
C
ISM 0031      IF (IPLOT1.EQ.0) GO TO 200
ISM 0033      CALL DCFIT1
ISM 0034      200 IF (ITREND.NE.0) GO TO 300
ISM 0036      PRINT 202
ISM 0037      202 FORMAT (1H0///1H0,'THE SPECTRAL ANALYSIS WILL BE ',
        1 'APPLIED TO RAW DATA WITHOUT REMOVING THE TREND'////)
        SPPE = 0.0
ISM 0038      220 DO 240 I=1,NPTS
ISM 0039      XD(I) = Y(I)
ISM 0040      240 CONTINUE
ISM 0041      GO TO 310
ISM 0042

C
C
ISM 0043      300 CONTINUE
ISM 0044      PRINT 302
ISM 0045      302 FORMAT (1H0///1H0,'THE SPECTRAL ANALYSIS WILL BE ',
        1 'APPLIED TO THE DETRENDED DATA'////)
        PRINT 303,YTIME,STATE,STATE
ISM 0046      303 FORMAT (1H0///1H0,2X,2A0,12X,2A0,10X,'DETRENDED ',2A0/ )
ISM 0047      DO 305 I=1,NPTS
ISM 0048      PRINT 304, DATE(I), Y(I), XDIFF(I)
ISM 0049      304 FORMAT ( 1H ,1F011.0,8X,1PD14.6,8X,1PD14.6)
ISM 0050      305 CONTINUE
ISM 0051

C
C
C      PUT IN PLOT FOR DETRENDED DATA.
C
C
C
ISM 0052      CALL DCFIT2
ISM 0053      310 CONTINUE

C
C
C      COMPUTE YULI-WALKER ESTIMATES OF AUTOCOVARIANCE
C      FUNCTION IF DESIRED. LAG VALUES STORED IN X,
C      ACV FUNCTION VALUES STORED IN XDIFF.
C
C
ISM 0054      IF (IYUACV.EQ.0) GO TO 500
ISM 0056      AN = NPTS
ISM 0057      CMEAN=0.0D0
ISM 0058      DO 410 I=1, NPTS
ISM 0059      CMEAN=CMEAN+ XD(I)
ISM 0060      X(I)=(I-1)
ISM 0061      XDIFF(I)=0.0D0
ISM 0062      410 CONTINUE
ISM 0063      CMEAN=CMEAN/AN

```

```

ISN 0064      DO 440 I=1, NPTS
ISN 0065      ILAG=I-1
ISN 0066      NMI = NPTS - ILAG
ISN 0067      DO 430 J=1, NMI
ISN 0068      JLAG=J+ILAG
ISN 0069      XDIFP(I)=XDIFP(I)+(( XD(J)-CMEAN)*( XD(JLAG)-CMEAN))
ISN 0070      430 CONTINUE
ISN 0071      XDIFP(I)=XDIFP(I)/AN
ISN 0072      440 CONTINUE
ISN 0073      XDP1 = XDIFP(1)
ISN 0074      DO 445 I = 1, NPTS
ISN 0075      XDIFP(I) = XDIFP(I)/XDP1
ISN 0076      445 CONTINUE
C
ISN 0077      PRINT 450
ISN 0078      450 FORMAT (1H1, 'YULE-WALKER ESTIMATE OF AUTOCORRELATION ',
ISN 0079      1 'FUNCTION'//1H0,4X,'LAG',11X,'AUTOCORR. '/')
ISN 0080      DO 460 I=1,NPTS
ISN 0081      PRINT 455, X(I),XDIFP(I)
ISN 0082      455 FORMAT (1H ,1P210.3,5X,1P212.5)
ISN 0083      460 CONTINUE
C
C      CALL DOPLOT3
C
C      COMPUTE PERIODOGRAM AND FOURIER SPECTRUM.
C
ISN 0084      500 NTR = NPTS
ISN 0085      KPI = NTR + 1
C
C      FIRST THE PERIODOGRAM AND CUMULATIVE PERIODOGRAM.
C
ISN 0086      PTIMP = NPTS / 2
ISN 0087      NPER = PTIMP
ISN 0088      DO 502 J=1,NPER
ISN 0089      AJ = J
ISN 0090      PPER(J) = AJ / NPTS
ISN 0091      PPER(J) = 1.000 / PPER(J)
ISN 0092      502 CONTINUE
ISN 0093      CALL SPT(XI,ACS,BSN,PPER,NPER,NPTS,PI)
ISN 0094      SUMPER = 0.000
ISN 0095      TWONTA = 2.000 / NPTS
ISN 0096      DO 503 I = 1,NPER
ISN 0097      PERIG(I) = TWONTA * (ACS(I)**2 + BSN(I)**2)
ISN 0098      SUMPER = SUMPER + PERIG(I)
ISN 0099      503 CONTINUE
ISN 0100      DSUM = 0.000
ISN 0101      PERMAX = PERIG(1)
ISN 0102      INAX = 1
ISN 0103      DO 505 I=1,NPER
ISN 0104      DSUM = DSUM + PERIG(I)
ISN 0105      CUMPER(I) = DSUM / SUMPER
ISN 0106      IF (PERIG(I) .LT. PERMAX) GO TO 505
ISN 0107      PERMAX = PERIG(I)
ISN 0108      INAX = I
ISN 0109      505 CONTINUE
ISN 0110      FISHER = PERMAX / ((1.000/NPER) * SUMPER)
ISN 0111

```

```

C
C
ISM 0112      PRINT THE PERIODOGRAM AND CUMULATIVE PERIODOGRAM.
ISM 0113
C
      PRINT 507
      507 FORMAT (1H1, 'PERIODOGRAM'//1H0, 21, 'K', 5X, 'PERIOD',
1 6X, 'PERIOD', 5X, 'PERIODOGRAM', 3X, 'CUMULATIVE', 1H,
2 3X, 'ORDINATE', 4X, 'PERIODOGRAM'//)
      508 FORMAT (1H, 'K', 5X, 'PERIOD', 5X, 'PERIODOGRAM', 3X, 'CUMULATIVE', 1H,
1 6X, 'PERIOD', 5X, 'PERIODOGRAM', 3X, 'CUMULATIVE', 1H,
2 3X, 'ORDINATE', 4X, 'PERIODOGRAM'//)
      509 FORMAT (1H, 'K', 5X, 'PERIOD', 5X, 'PERIODOGRAM', 3X, 'CUMULATIVE', 1H,
1 6X, 'PERIOD', 5X, 'PERIODOGRAM', 3X, 'CUMULATIVE', 1H,
2 3X, 'ORDINATE', 4X, 'PERIODOGRAM'//)
      510 FORMAT (1H0, 'MAXIMUM ORDINATE V', 1H0, 12, 5, ' OCCURRING AT ',
1 /1H0, 'FREQUENCY =', 1H0, 12, 5, ' PERIOD =', 1H0, 12, 5, 1H0,
2 ' FISHER STATISTIC =', 1H0, 12, 5, 1H0, ' NUMBER OF ',
3 ' ORDINATES =', 1H)
C
C
C      COMPUTE THE KP+1 FREQUENCIES FOR FOURIER AND HESA SPECTRA.
C
      V(1) = 0.
      V(2) = VMAX / NV
      KPP = KP1 - 2
      DO 510 I=1, KPP
        V(I+2) = V(2) + I**V(2)
      510 CONTINUE
C
C
C      COMPUTE THE FOURIER SPECTRUM.
C
      CALL SPT(ED, ACS, BSM, V, KP1, NYR, PI)
      DO 520 I=1, KP1
        ASP(I) = (2.*ACS(I)/NYR)**2 + (2.*BSM(I)/NYR)**2
        ASF(I) = DSQRT(ASP(I))
      520 CONTINUE
C
C
C      COMPUTE THE CUMULATIVE FOURIER SPECTRUM.
C
      YSUM = 0.
      DO 535 I=1, KP1
        YEAR(I) = NYR**2 / 2 * (ASF(I)**2)
        YSUM = YSUM + YEAR(I)
      535 CONTINUE
      ZSUM = 0.0D0
      DO 540 J=1, KP1
        ZSUM = ZSUM + YEAR(J)
        BSM(J) = ZSUM / YSUM
      540 CONTINUE
      PRINT 550, YSUM
      550 FORMAT (1H1, 'SUM OF FOURIER ORDINATES =', 1PE19.6)
C
C
C      IF (FLOT-EG-0) GO TO 600
C      CALL DOP19(KP1, NYR)
C      CALL DOP19(KP1)
C
C
C      APPLY THE MAXIMUM ENTROPY METHOD. REFERENCE TO
C      ANDERSEN'S ALGORITHM.
C

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C
C
ISM 0145      600 N1 = MYR - 1
ISM 0146      SUM = 0.
ISM 0147      DO 620 I=1,MYR
ISM 0148      SUM = SUM + XD(I)**2
ISM 0149      620 CONTINUE
ISM 0150      SUM = SUM/MYR
ISM 0151      ANP(1) = SUM
ISM 0152      XDIFF(1) = ANP(1)

C
C
C      MAXIMUM ENTROPY COMPUTATIONS BEGIN HERE.
C
ISM 0153      IFLAG = 0
ISM 0154      DO 1000 IF = 1,NFILT
ISM 0155      IFLAG = IFLAG + 1
ISM 0156      IF (IFLAG.GT.1) GO TO 795
ISM 0157      N = 1
ISM 0158      ACS(1) = XD(1)
ISM 0159      BSN(N1) = XD(MYR)
ISM 0160      DO 640 I=2,N1
ISM 0161      ACS(I) = XC(I)
ISM 0162      BSN(I-1) = XD(I)
ISM 0163      640 CONTINUE
ISM 0164      660 DMCH = 0.
ISM 0165      DDEN = 0.
ISM 0166      NM = MYR-N
ISM 0167      DO 680 I=1,NM
ISM 0168      DMCH = DMCH + ACS(I)*BSN(I)
ISM 0169      DDEN = DDEN + ACS(I)**2 + BSN(I)**2
ISM 0170      680 CONTINUE
ISM 0171      Z(N) = 2 * DMCH/DDEN
ISM 0172      APP(N+1) = APP(N)*(1.-Z(N)**2)
ISM 0173      QNUM = N*N1
ISM 0174      QDEN = MYR-N
ISM 0175      IF (.TREND.EQ.0.) GO TO 690
ISM 0176      QNUM = QNUM*SPFE
ISM 0177      QDEN = QDEN - SPFE
ISM 0178      690 CONTINUE
ISM 0179      IF (QDEN.ZC.0.00) QDEN=1.0D-5
ISM 0180      FFE(N) = APP(N+1) * QNUM/QDEN
ISM 0181      IF (N.NE.1) GO TO 760
ISM 0182      XDIFF(2) = XDIFF(1) * Z(1)
ISM 0183      720 N = N+1
ISM 0184      N1 = N-1
ISM 0185      DO 730 I=1,N1
ISM 0186      YEAR(I) = Z(I)
ISM 0187      730 CONTINUE
ISM 0188      NM = MYR-N
ISM 0189      DO 740 I=1,NM
ISM 0190      ACS(I) = ACS(I) - YEAR(N-1)*BSN(I)
ISM 0191      BSN(I) = BSN(I+1) - YEAR(N-1)*ACS(I+1)
ISM 0192      740 CONTINUE
ISM 0193      GO TO 660
ISM 0194      760 N1 = N-1
ISM 0195      DO 780 I=1,N1

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ISN 0200      Z(I) = YEAR(I) - Z(N)*YEAR(N-I)
ISN 0201      780 CONTINUE
ISN 0202      SUN = 0.
ISN 0203      DO 790 I=1,N
ISN 0204      SUN = SUN + IDIFF(N+1-I)*Z(I)
ISN 0205      790 CONTINUE
ISN 0206      IDIFF(N+1) = SUN
ISN 0207      GO TO 798
ISN 0208      795 CONTINUE
ISN 0209      N = MCOUNT(IP-1)
ISN 0210      798 CONTINUE
ISN 0211      IF (N.EC-MCOUNT(IP)) GO TO 820
ISN 0212      800 GO TO 720
ISN 0213      820 DO 860 I=1,N*1
ISN 0214      CSUM = 0.
ISN 0215      DO 840 J=1,N
ISN 0216      ENT = - 2. * PI * J * V(I) * DELTA
ISN 0217      CV = ECOMPLX(0.00,ENT)
ISN 0218      CSUM = CSUM + Z(J)*CDEFP(CV)
ISN 0219      840 CONTINUE
ISN 0220      CSUM = 1.+CSUM
ISN 0221      EX = (CDABS(CSUM))**2
ISN 0222      S(I) = (ABS(N+1) * DELTA) / EX
ISN 0223      860 CONTINUE
ISN 0224      NC = MCOUNT(IP)
ISN 0225

C
C
C
C      PRINT OUT AUTOCORRELATION AND ENTROPY VALUES.

ISN 0226      NYREQ = 1.000 / (2.000 * DELTA)
ISN 0227      RESBW = NYREQ/((DFLOAT(MCOUNT(IP))/2.000)+1.000)
ISN 0228      IF (IFPRINT.EQ.0) GO TO 980
ISN 0229      IF (IAUTCV.EQ.C) GO TO 891
ISN 0230      PRINT 869,MCOUNT(IP)
ISN 0231      869 FORMAT (1H1,'MCOUNT = ',I5//)
ISN 0232      PRINT 870
ISN 0233      870 FORMAT (1H0/1H0,'AUTOCORRELATION FUNCTION VALUES'//,
ISN 0234      1 1H0,'LAG',20X,'AUTOCORR.'/)
ISN 0235      ACT = MCOUNT(IP)+1
ISN 0236      IDP1 = IDIFF(1)
ISN 0237      DO 875 I=1,ACT
ISN 0238      ACPN(I) = IDIFF(I) / IDP1
ISN 0239      875 CONTINUE
ISN 0240      DO 890 I=1,ACT
ISN 0241      X(I) = I-1.
ISN 0242      PRINT 880, X(I), ACPN(I)
ISN 0243      880 FORMAT (1H ,1P2819.6)
ISN 0244      890 CONTINUE
ISN 0245      891 CONTINUE
ISN 0246      PRINT 895, NYREQ, RESBW, FPE(NC)
ISN 0247      895 FORMAT (1H1,'MAXIMUM ENTROPY AND FOURIER SPECTRA',///,
ISN 0248      1 1H0,'NYQUIST FREQUENCY = ',1PD10.3//,
ISN 0249      2 1H0,'RESOLUTION BANDWIDTH = ',1PD10.3//,
ISN 0250      3 1HC,'FINAL PREDICTION ERROR = ',1PD10.3//)
ISN 0249      PRINT 900, MCOUNT(IP)
ISN 0250      900 FORMAT (1H0,'MCOUNT=',I3,6X,'FREQUENCY',6X,'PERIOD',10X,
ISN 0251      1'HAXIMUM ENTROPY',7X,'FOURIER AMPLITUDE')

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ISN 0251      DO 960 I=1,NP1
ISN 0252      IF (V(I).EQ.0) GO TO 920
ISN 0254      910 PERIOD(I) = L/V(I)
ISN 0255      GO TO 930
ISN 0256      920 PERIOD(I) = 0.
ISN 0257      930 PRINT 940,I,V(I),PERIOD(I),S(I),ASF(I)
ISN 0258      940 FORMAT (1H,'I=',I4,10X,P9.7,6X,P10.3,6X,1P19.6,6X,1P19.6)
ISN 0259      960 CONTINUE

```

```

C
C
C      PLCT MAXIMUM ENTROPY AND AUTOCORRELATION FUNCTION.
C
C

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```

ISN 0260      980 IF (IFLOT.EQ.0) GO TO 91
ISN 0262      IF (IAUTCV.EQ.0) GO TO 90
ISN 0264      CALL DOP1T5(NCT,NC,ACFN)
ISN 0265      90 CONTINUE
ISN 0266      CALL TCF1T6(KF1,NC)
ISN 0267      91 CONTINUE

```

```

C
C
C      COMPUTE INTEGRATED SPECTRUM IF DESIRED
C
C

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ISN 0268      IF (IINTSI.EQ.0) GO TO 1030

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ISN 0270      WINDOW = PESBW/2.000
ISN 0271      PRINT 991,WINDOW,NCOUNT(IP)
ISN 0272      991 FORMAT (1H1,'INTEGRATED MAXIMUM ENTROPY SPECTRUM',///,
1 1H0,'WIDTH OF INTEGRATION WINDOW = ',1PD10.3,///,
2 1H0,'NCOUNT=',I3,6X,'FREQUENCY',6X,'PERIOD',10X,
3 1H0,'MAX. ENT. SPEC.',7X,'INTEGRATED SPECTRUM',7X,
4 1H0,'ERROR ED. (INTEG. SP.)',//)
ISN 0273      WINDOW2=WINDOW/2.000
ISN 0274      SPCVAL = SPCVAL1 (ANP,2,DELTA,KF1,N)
ISN 0275      DO 995 I=1,NP1
ISN 0276      VLC=V(I)-WINDOW2
ISN 0277      VUP=V(I)+WINDOW2
ISN 0278      IF ((VLC.GE.0.000).AND.(VUP.LE.NYFREQ)) GO TO 992
ISN 0280      A(I) = 0.000
ISN 0281      B(I) = 0.000
ISN 0282      GO TO 995
ISN 0283      992 CONTINUE

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ISN 0284      C
1  A(I) =DLADRE(SPCVAL,VLO,VUP,0.000,0.00100,
1  ERROR,IER)
ISN 0285      B(I) = PRFOR
ISN 0286      IF (IER.LE.66) GO TO 995
ISN 0288      PRINT 993, I, IER, V(I), A(I), B(I)
ISN 0289      993 FORMAT (1H0/1H0,'I =',I4,5X,'IER =',I3,5X,'FREQ. =',
1 1P14.6,5X,'INTEGRAL =',1PD14.6,5X,'ERROR PARAMETER =',
2 1P14.6//)
ISN 0290      995 CONTINUE
ISN 0291      DO 997 I=1,NP1
ISN 0292      PRINT 996, I,V(I),PERIOD(I),S(I),A(I),B(I)
ISN 0293      996 FORMAT(1H,'I=',I4,10X,P9.7,6X,P10.3,6X,1P19.6,6X,
1 1P19.6,6X,1P19.6)

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ISM 0294      997 CONTINUE
ISM 0295      CALL DOP177(RF1,MC,A)
C
C
ISM 0296      1000 CONTINUE
C
C
ISM 0297      IF (1/PPEFL.EQ.0) RETURN
ISM 0299      PRINT 1010
ISM 0300      1010 FORMAT (1H1,'FINAL PREDICTION ERROR',//1H2,2X,
        1 '1',3X,'NUMBER OF FILTER',4X,'FINAL PREDICTION',//1H ,
        1 8X,'COEFFICIENTS',11X,'ERROR',//)
ISM 0301      DO 1020 I=1,MC
ISM 0302      I1 = I+1
ISM 0303      PRINT 1015,I,I1,PPE(I)
ISM 0304      1015 FORMAT (1H ,13,115,12X,1PD12.5)
ISM 0305      1020 CONTINUE
C
C
C      PLOT THE FINAL PREDICTION ERROR.
C
C
ISM 0306      CALL LUP18
ISM 0307      RETURN
ISM 0308      END

```

***** F O R T R A N   C R O S S   R E F E R E N C E   L I S T I N G *****

SYMBOL	INTERNAL STATEMENT NUMBERS
A	0011 0280 0284 0288 0292 0295
B	0011 0281 0285 0288 0292
I	0022 0023 0023 0028 0029 0039 0339 0040 0040 0048 0049 0049 0049 0058 0059 0060 0060 0061 0064 0065 0069 0069 0071 0071 0074 0075 0075 0079 0080 0080 0096 0097 0097 0097 0098 0103 0104 0105 0106 0108 0109 0121 0122 0122 0125 0126 0126 0126 0127 0127 0130 0131 0131 0132 0147 0148 0161 0162 0162 0163 0163 0168 0169 0169 0170 0170 0189 0190 0190 0193 0194 0194 0194 0195 0195 0195 0199 0200 0200 0200 0203 0204 0204 0218 0217 0223 0238 0239 0241 0242 0242 0243 0243 0251 0252 0254 0254 0256 0257 0257 0257 0257 0257 0275 0276 0277 0280 0281 0284 0285 0288 0288 0288
J	0067 0068 0069 0074 0089 0090 0091 0091 0135 0136 0137 0216 0217 0219
K	0114 0114 0114 0114 0114 0114 0114 0114
M	0158 0167 0172 0173 0173 0173 0174 0175 0183 0183 0184 0187 0187 0188 0192 0194 0195 0198 0200 0200 0203 0204 0206 0209 0211 0216 0223 0274
S	0006 0223 0257 0292
V	0006 0118 0119 0122 0122 0122 0124 0217 0252 0254 0257 0276 0277 0288 0292
X	0006 0027 0060 0080 0242 0243
Y	0006 0023 0027 0040 0049
Z	0006 0172 0173 0186 0190 0200 0200 0204 0219 0274
AJ	0089 0090
AM	0056 0063 0071
CV	0010 0218 0219
EX	0222 0223
IP	0154 0209 0211 0225 0227 0232 0236 0249 0271
II	0302 0303
MC	0225 0247 0264 0266 0295 0301
NI	0188 0189 0198 0199
NR	0167 0168 0192 0193
NV	0006 0085 0119
NI	0145 0160 0161
PI	0012 0093 0124 0217
XD	0006 0029 0040 0059 0069 0093 0124 0148 0159 0160 0162 0163
ACS	0006 0093 0097 0124 0126 0154 0162 0169 0170 0194 0194 0195
AMP	0006 0151 0152 0173 0173 0183 0223 0274
ASF	0006 0126 0127 0127 0131 0257
BSN	0006 0093 0097 0124 0126 0137 0160 0163 0169 0170 0194 0195 0195
BWT	0217 0218
FPE	0006 0183 0247 0303
IEE	0284 0286 0288
KPF	0120 0121
KPI	0085 0120 0124 0125 0130 0135 0143 0144 0214 0251 0266 0274 0275 0291 0295
MCI	0236 0238 0241 0264
NAR	0005 0007 0016
NRI	0066 0067
NTR	0084 0124 0126 0126 0145 0147 0150 0160 0167 0174 0175 0192
SPT	0093 0124
SUR	0146 0148 0148 0150 0150 0151 0202 0204 0204 0206
VLO	0276 0274 0284
VUP	0277 0275 0284
ACFN	0011 0239 0243 0264
CSGN	0010 0215 0219 0219 0221 0221 0222
DATE	0006 0023 0027 0049
DDEN	0166 0170 0170 0172
DNOR	0165 0169 0169 0172
DSUN	0100 0104 0104 0105
FPER	0008 0090 0091 0093 0114 0116
ILAG	0065 0066 0068
IMAX	0104 0109 0116 0116

***** C E T A N C R O S S R E F E R E N C E L I S T I N G *****

SYMBOL	INTERNAL	STATEMENT	NUMBERS
JLAG	0068	0069	
NPFR	0087	0088	0093
NPFS	0306	0022	0027
PPRN	0008	0091	0114
QDEN	0175	0179	0179
QNR	0174	0178	0183
SPPE	0027	0038	0178
YBAI	0006	0119	
YFPI	0073	0075	0237
YFAR	0006	0027	0131
YSDR	0129	0132	0137
YSDR	0134	0136	0137
CDABS	0222		
CDENP	0219		
CREAN	0057	0059	0063
DELTA	0006	0217	0223
DOPLT	0013		
DSQRT	0127		
ERRON	0201	0285	
IPLAG	0153	0155	0156
IPLOT	0006	0031	0141
NP'LT	0006	0154	
P DG	0008	0097	0098
PIENP	0086	0087	
RESB4	0227	0247	0270
STATE	0005	0007	0018
XDIFP	0006	0027	0029
YTIME	0005	0007	0018
CUMPER	0008	0105	0114
DCADRE	0284		
DCMPLX	0218		
DETEND	0027		
DFLOAT	0227		
DOPLT1	0033		
DOPLT2	0052		
DOPLT3	0083		
DOPLT4	0144		
DOPLT5	0264		
DOPLT6	0266		
DOPLT7	0295		
DOPLT8	0306		
DOPLT9	0143		
FISHER	0111	0116	
IAUTCV	0006	0230	0262
IPPEPL	0006	0297	
IINTSP	0006	0268	
IPRINT	0006	0014	0228
ITREVD	0006	0034	0176
ITNACY	0004	0054	
BCOUNT	0006	0209	0211
HYPERQ	0004	0226	0227
PERIOD	0006	0254	0257
PERMAX	0101	0106	0111
SPCYAL	0009	0284	

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***** C I T R A N   C R O S S   R E F E R E N C E   L I S T I N G *****

SYMBOL	INTERNAL STATEMENT NUMBERS					
SPCVIB	0276					
SPCVL1	0276					
SPECTR	0002					
SUPER	0096	0098	0098	0105	0111	0116
TWOPTA	0095	0097				
WINDO1	0270	0271	0273			
WINDO2	0273	0276	0277			

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***** U P I R A M C R O S S R E F E R E N C E L I S T I N G *****

*****
REFERENCES

DERIVED

LABEL

10	0017	0016
20	0018	0018
40	0019	0020
60	0021	0020
70	0024	0021
80	0025	0022
90	0025	0022
91	0026	0020
100	0026	0019
170	0030	0028
200	0034	0031
202	0037	0036
220	0035	
240	0041	0039
300	0042	0038
302	0045	0044
303	0047	0046
304	0050	0049
305	0051	0048
310	0053	0042
410	0062	0058
430	0070	0067
440	0072	0064
445	0076	0074
450	0078	0077
455	0081	0080
460	0082	0079
500	0084	0084
502	0092	0088
503	0095	0096
505	0110	0103
507	0113	0112
508	0115	0114
509	0117	0116
510	0123	0121
520	0128	0125
535	0133	0130
540	0138	0135
550	0140	0139
600	0145	0141
620	0149	0147
640	0164	0161
660	0165	0167
680	0171	0168
690	0180	0176
700	0186	0186
710	0187	0189
730	0191	0193
740	0196	0193
760	0198	0188
780	0201	0199
790	0205	0203
795	0208	0196
798	0210	0207
800	0213	
820	0214	0211
840	0220	0216

0106

ORNL-DWG 78-5864

***** O R T R A M C R O S S R E F E R E N C E L I S T I N G *****

LABEL	DELETED	REFERENCES
860	0224	0214
869	0233	0232
870	0234	0238
875	0240	0238
880	0244	0243
890	0245	0241
891	0246	0230
895	0248	0247
900	0250	0249
910	0254	
920	0256	0252
930	0257	0255
940	0258	0257
940	0259	0251
980	0260	0220
991	0272	0271
992	0283	0278
993	0289	0288
995	0290	0275
996	0293	0292
997	0294	0291
1000	0296	0154
1010	0300	0266
1015	0300	0299
1020	0305	0303
		0301


```

LEVEL 21.0 ( JOB '0' )
                                OS/360  FORTRAN 8
COMPILE OPTIONS - . . . . .  AIN,OPT=02,LIST=40,SIZE=00001,
SOURCE,BCDIC,NOLIST,NOBCE,LOAD,MAP,MODIT,10,INP
C
C
C
100 0002      SUBROUTINE SP7(I,A,B,V,K,N,PI)
100 0003      IMPLICIT REAL*8(A-H,O-Z)
100 0004      DIMENSION I(1),A(1),S(1),V(1)
100 0005      DO 40 I=1,K
100 0006      A(I) = 0.
100 0007      S(I) = 0.
100 0008      DO 20 J=1,N
100 0009      ALPHA = 2.0*PI*(J-1)*V(I)
100 0010      ASIN = ASIN(I(1)*COS(ALPHA))
100 0011      BSIN = SIN(I(2)*COS(ALPHA))
100 0012      CONTINUE
100 0013      A(I) = ASIN
100 0014      S(I) = BSIN
100 0015      CONTINUE
100 0016      RETURN
100 0017      END
20
40
END

```

----- CROSS REFERENCE LISTING-----

SYMBOL INTERNAL STATEMENT NUMBERS

A	0002 0004 0013
B	0002 0004 0014
C	0005 0009 0013 0014
D	0004 0009 0010 0011
E	0002 0005
F	0002 0006
G	0002 0006 0009
H	0002 0006 0010
I	0002 0006 0010 0011
J	0002 0009
K	0002
L	0006 0010 0013
M	0007 0011 0011 0016
N	0010
O	0011
P	0009 0010 0011

***** O B T A I N C R O S S R E F E R E N C E L I S T I N G *****

NAME: DUTCHMAN
20 0012
00 0015

DATE: 0012
00 0015

REFERENCE
0000
0005

03/360 PORTMAN 1

```

COMPILER OPTIONS - NAME= RAW,OPT=02,LINCHIT=60,SIZB=0000K,  
SOURCEC,BSCDIC,NOLIST,MODBACK,LOAD,RAP,MORDR7,ID,XBBP  
  
C  
C  
C DOUBLE PRECISION FUNCTION SPCVAL (PRNO)  
  
    IMPLICIT REAL*8 (A-H,O-S)  
    COMMON/16 CSUM,CV  
    DIMENSION AMP(1),Z(1)  
    DATA PI /3.1415926535897931/  
  
C  
    CSUM=0.000  
    DO 100 J=1,N  
        RT = -2.0D+0E+CJ*PRNQ*DTIME  
        CV = DCPLX1 (0.0D0,RT)  
        CSUM = CSUM + Z(J)*CDIEXP(CV)  
    100 CONTINUE  
    CSUM = 1.0D+CSUM  
    XI = CABS(CSUM))**2  
    SPCVAL = (AMP(N+1)*DTIME)/XI  
  
C  
C RETURN  
  
C  
C RETURN SPCVAL (ARP,Z,DTIME,KF1,H)  
  
C INITIAL ENTRY FOR DEFINING PARAMETERS IN  
C ARGUMENT LIST  
C  
C SPCVAL1 = 777777.  
C  
C RETURN  
  
C  
END
```

STUDID	***** O B T E I N C R O S S R E P R E M C B L I S T I N D O O O O O
J	INTERNAL STATEMENT NUMBERS
0000	0009 0011
0000	0015 0017
0005	0011 0017
0005	0010 0011
0010	0015 0015
0010	0009 0009
0015	0005 0015 0017
0015	0009 0010
0017	0007 0011 0013 0013 0014
0007	0007 0009
0002	0010
0010	0011 0015 0017
0011	0004 0010 0015
0004	0002 0015
0010	0017 0010
0004	0007 0011 0013 0013 0014
0007	0009
0002	0010
0010	0011 0015 0017
0011	0004 0010 0015
0004	0002 0015
0010	0017 0010
0004	0007 0011 0013 0013 0014
0007	0009
0002	0010
0010	0011 0015 0017
0011	0004 0010 0015
0004	0002 0015
0010	0017 0010
0004	0007 0011 0013 0013 0014
0007	0009
0002	0010
0010	0011 0015 0017
0011	0004 0010 0015
0004	0002 0015
0010	0017 0010
0004	0007 0011 0013 0013 0014
0007	0009
0002	0010
0010	0011 0015 0017
0011	0004 0010 0015
0004	0002 0015
0010	0017 0010
0004	0007 0011 0013 0013 0014
0007	0009
0002	0010
0010	0011 0015 0017
0011	0004 0010 0015
0004	0002 0015
0010	0017 0010
0004	0007 0011 0013 0013 0014
0007	0009
0002	0010
0010	0011 0015 0017
0011	0004 0010 0015
0004	0002 0015
0010	0017 0010
0004	0007 0011 0013 0013 0014
0007	0009
0002	0010
0010	0011 0015 0017
0011	0004 0010 0015
0004	0002 0015
0010	0017 0010
0004	0007 0011 0013 0013 0014
0007	0009
0002	0010
0010	0011 0015 0017
0011	0004 0010 0015
0004	0002 0015
0010	0017 0010
0004	0007 0011 0013 0013 0014
0007	0009
0002	0010
0010	0011 0015 0017
0011	0004 0010 0015
0004	0002 0015
0010	0017 0010
0004	0007 0011 0013 0013 0014
0007	0009
0002	0010
0010	0011 0015 0017
0011	0004 0010 0015
0004	0002 0015
0010	0017 0010
0004	0007 0011 0013 0013 0014
0007	0009
0002	0010
0010	0011 0015 0017
0011	0004 0010 0015
0004	0002 0015
0010	0017 0010
0004	0007 0011 0013 0013 0014
0007	0009
0002	0010
0010	0011 0015 0017
0011	0004 0010 0015
0004	0002 0015
0010	0017 0010
0004	0007 0011 0013 0013 0014
0007	0009
0002	0010
0010	0011 0015 0017
0011	0004 0010 0015
0004	0002 0015
0010	0017 0010
0004	0007 0011 0013 0013 0014
0007	0009
0002	0010
0010	0011 0015 0017
0011	0004 0010 0015
0004	0002 0015
0010	0017 0010
0004	0007 0011 0013 0013 0014
0007	0009
0002	0010
0010	0011 0015 0017
0011	

***** ORTRAN CROSS REFERENCE LISTING*****

LABEL	DEFINED	REFERENCES
000	0012	0000

LEVEL 21.8 (JUN 74)

OS/360 FORTRAN H

CCMPILER OPTIONS - NAME= MAIN,OPT=02,LINENCT=60,SIZE=0000K,
SOURCE,IBCDIC,NOLIST,MODECK,LOAD,NAP,NOEDIT,ID,XREF

```

C
C
ISN 0002      SUBROUTINE DCPIT
ISN 0003      IMPLICIT REAL*8(A-H,O-Z)
ISN 0004      REAL*8 XORIG,YSTEP,XORIG,YCYCLE,XAX,YAX
ISN 0005      REAL*4 NAB,YTIME,STATE
ISN 0006      REAL*4 XS(600),YS(600)
ISN 0007      COMMON /SPICBL/ DTIME,YEAR(600),CATCH(600),XE(600),
1 VMAX,Z(600),V(600),S(600),PPE(600),XDIPP(600),
2 DATL(600),X(600),ACS(600),BSH(600),ANP(600),ASF(600),PERIOD(600),
3 NYA,NFILT,KF,NCOUNT(600),IPRINT,IPLT,ITPENTL,ITPEPL,IINTSP,
4 IAUTCV,TYMACV
ISN 0008      COMMON /ELELEL/ NAB(14),YTIME(2),STATE(2)
ISN 0009      COMMON /PRDGRM/ PPAR(300),PPER(300),PERDG(300),CUNYER(300)
ISN 0010      COMMON /QOEXIT/ XORIG, YSTEP, XORIG, YCYCLE, XAX, YAX, IDUN(3)
ISN 0011      DIMENSION ACEN(600),A(600)

C
C
ISN 0012      CALL CALCHF
ISN 0013      XLONG = 7.
ISN 0014      YLONG = 5.
ISN 0015      RETURN

C
C
ISN 0016      ENTRY DOP11
ISN 0017      CALL GPM3GN(XS,YS,DATE,CATCH,NYA,XORIG,XNAX,YORIG,YNAX)
ISN 0018      CALL TITLE(NAB,100,YTIME,100,STATE,100,XLONG,YLONG)
ISN 0019      CALL ANGLIC
ISN 0020      XORIG = XS(1)
ISN 0021      XNAX = XS(NYA)
ISN 0022      CALL GRAP(XORIG,'SCALE',XNAX,YORIG,'SCALE',YNAX)
ISN 0023      CALL NABREF(5)
ISN 0024      CALL CURVE(XS,YS,NYA,1)
ISN 0025      CALL BSET('MARKER')
ISN 0026      IF (ITRND.IQ-0) GO TO 199
ISN 0028      CALL SINGLE(XS,YS,DATE,X,NYA)
ISN 0029      CALL EASE
ISN 0030      CALL CURVE(XS,YS,NYA,0)
ISN 0031      CALL BSET('DASH')
ISN 0032      199 CONTINUE
ISN 0033      CALL FRAMB
ISN 0034      CALL ENDPL(0)
ISN 0035      RETURN

C
C
ISN 0036      ENTRY DOP12
ISN 0037      CALL GPM3GN(XS,YS,DATE,XDIPP,NYA,XORIG,XNAX,YORIG,YNAX)
ISN 0038      CALL TITLE('DWTREPLD DATAS',100,YTIME,100,STATE,100,
1 XLONG,YLONG)
ISN 0039      CALL ANGLIC
ISN 0040      XORIG = XS(1)
ISN 0041      XNAX = XS(NYA)
ISN 0042      CALL GRAP(XORIG,'SCALE',XNAX,YORIG,'SCALE',YNAX)
ISN 0043      CALL CURVE(XS,YS,NYA,0)
ISN 0044      DO 300 I=1,NYA
ISN 0045      YS(I) = 0.

```

```

ISN 0046      300 CONTINUE
ISN 0047      CALL CURVE(XS,YS,NYA,0)
ISN 0048      CALL FRAMF
ISN 0049      CALL ENDPL(0)

C
C
C
C
ISN 0050      CALL GPHDGN(XS,YS,X,DIFF,NYA,XORIG,INAX,YORIG,YMAX)
ISN 0051      CALL TITLE('RESIDUALS VS. FITTED VALUES',100,'FITTED VALUES',
1             100,'RESIDUALS',100,7,5.)
ISN 0052      CALL ANOTIC
ISN 0053      CALL GRAP(XORIG,'SCALE',XMAX,YORIG,'SCALE',YMAX)
ISN 0054      CALL MARKS(5)
ISN 0055      CALL CURVE(XS,YS,NYA,-1)
ISN 0056      CALL RESH('MARKER')
ISN 0057      CALL FRAMF
ISN 0058      CALL ENDPL(3)
ISN 0059      RETURN

C
C
ISN 0060      ENTRY DOPLT3
ISN 0061      CALL GPHDGN(XS,YS,X,DIFF,NYA,XORIG,INAX,YORIG,YMAX)
ISN 0062      CALL TITLE('TOLE-WALKER AUTOCORRELATIONS',100,'LAGS',100,
1             100,'ACV. FN. S',100,XLONG,YLONG)
ISN 0063      CALL GPHDGN(XORIG,YORIG,INAX,YMAX,XS,YS,NYA)
ISN 0064      RETURN

C
C
ISN 0065      ENTRY DOPLT4(KF1)
ISN 0066      CALL GPHDGN(XS,YS,V,ASF,KF1,XORIG,INAX,YORIG,YMAX)
ISN 0067      CALL TITLE('FOURIER AMPLITUDE SPECTRUMS',100,'FREQUENCIES',100,
1             100,'AMPLITUDES',100,XLONG,YLONG)
ISN 0068      CALL GPHDGN(XORIG,YORIG,INAX,YMAX,XS,YS,KF1)
ISN 0069      CALL GPHDGN(XS,YS,V,BSE,KF1,XORIG,INAX,YORIG,YMAX)
ISN 0070      CALL TITLE('COS. FOURIER POWER SPECTRUMS',100,'FREQUENCIES',100,
1             100,'NORMALIZED POWERS',100,XLONG,YLONG)
ISN 0071      CALL GRAP(XORIG,'SCALE',INAX,YORIG,'SCALE',YMAX)
ISN 0072      CALL CURVE(XS,YS,KF1,0)
ISN 0073      IF ((2*(NYA/2)-80-NYA) QQQ = (NYA - 2) / 2
ISN 0074      IF ((2*(NYA/2)-80-NYA) QQQ = (NYA - 1) / 2
ISN 0075      RATIO = 1.35 / DSQRT(QQQ)
ISN 0076      YR(1) = 0.0
ISN 0077      YR(2) = RATIO
ISN 0078      YR(3) = 0.5 + (1.0-RATIO)
ISN 0079      YR(4) = 1.0
ISN 0080      CALL CURVE(XS,YS,2,0)
ISN 0081      YS(1) = 0.5 * RATIO
ISN 0082      YS(2) = 0.0
ISN 0083      YS(3) = 0.5
ISN 0084      YS(4) = (1.0 - RATIO)
ISN 0085      CALL CURVE(XS,YS,2,0)
ISN 0086      CALL FRAMF
ISN 0087      CALL ENDPL(0)
ISN 0088      RETURN
ISN 0089
ISN 0090

```



```

ISN 0091      ENTRY DOPLT5(HCT,HC,ACFM)
ISN 0092      CALL GPNBGN(XS,YS,X,ACFM,HCT,XORIG,XMAX,YORIG,YMAX)
ISN 0093      CALL TITLE(1H,1,'FREQUENCIES',100,'BURG AUTOCORRELATIONS',
1 100,XLONG,YLONG)
ISN 0094      CALL MESSAG('NO. OF FILTER COEFS. ISS',100,.5,YLONG+.25)
ISN 0095      CALL INTWO(HC,3.5,YLONG+.25)
ISN 0096      CALL GPHEND(XORIG,YORIG,XMAX,YMAX,XS,YS,HCT)
ISN 0097      RETURN

C
ISN 0098      ENTRY DOPLT6(KP,HC)
ISN 0099      CALL GPNBGN(XS,YS,V,S,KP,XORIG,XMAX,YORIG,YMAX)
ISN 0100      CALL TITLE(1H,1,'FREQUENCIES',100,'RESA VALUES',100,
1 XLONG,YLONG)
ISN 0101      CALL MESSAG('NO. OF FILTER COEFS. ISS',100,.5,YLONG+.25)
ISN 0102      CALL INTWO(HC,3.5,YLONG+.25)
ISN 0103      CALL GPHEND(XORIG,YORIG,XMAX,YMAX,XS,YS,KP)
ISN 0104      RETURN

C
ISN 0105      ENTRY DOPLT7(KP1,HC,A)
ISN 0106      CALL GPNBGN(XS,YS,V,A,KP1,XORIG,XMAX,YORIG,YMAX)
ISN 0107      CALL TITLE(1H,1,'FREQUENCIES',100,'INTEGRATED SPECTRUMS',100,
1 XLONG,YLONG)
ISN 0108      CALL MESSAG('NO. OF FILTER COEFS. ISS',100,.5,YLONG+.25)
ISN 0109      CALL INTWO(HC,3.5,YLONG+.25)
ISN 0110      CALL GPHEND(XORIG,YORIG,XMAX,YMAX,XS,YS,KP1)
ISN 0111      RETURN

C
ISN 0112      ENTRY DOPLT8
ISN 0113      NYA2 = NYA - 3
ISN 0114      CALL GPNBGN(XS,YS,NYA2,PPER)
ISN 0115      XAX = XICRG
ISN 0116      XSTEP = FLOAT(NYA) / XAX
ISN 0117      NSTEP = XSTEP
ISN 0118      XSTEP = FLCAT(NSTEP) + 1.
ISN 0119      CALL TITLE(NYA,100,'NO. OF FILTER COEFS. S',100,
1 'FINAL PREDICTION ERRORS',100,XLONG,YLONG)
ISN 0120      CALL XTICKS(5)
ISN 0121      CALL YLOG(XORIG,XSTEP,YORIG,YCYCLE)
ISN 0122      CALL CURVE(XS,YS,NYA2,0)
ISN 0123      CALL FRAME
ISN 0124      CALL ENDPL(0)
ISN 0125      RETURN

C
ISN 0126      ENTRY DOPLT9(KP1,NPER)
ISN 0127      CALL GPNBGN(XS,YS,PPER,PERDG,NPER,XORIG,XMAX,YORIG,YMAX)
ISN 0128      CALL TITLE('PERIODOGRAMS',100,'FREQUENCIES',100,
1 'POWERS',100,XLONG,YLONG)
ISN 0129      CALL GPHEND(XORIG,YORIG,XMAX,YMAX,XS,YS,NPER)
ISN 0130      CALL GPNBGN(XS,YS,PPER,CUMPER,NPER,XORIG,XMAX,
1 SIG,YMAX)
ISN 0131      CALL TITLE('CUMULATIVE PERIODOGRAMS',100,'FREQUENCIES',100,
1 'NORMALIZED POWERS',100,XLONG,YLONG)
ISN 0132      CALL GRAP(XORIG,'SCALE',XMAX,YORIG,'SCALE',YMAX)
ISN 0133      CALL CURVE(XS,YS,NPER,0)
ISN 0134      IF ((2*(NYA/2)).EQ.NYA) QQQ = (NYA - 2) / 2

```

```

ISN 0136
ISN 0138
ISN 0139
ISN 0140
ISN 0141
ISN 0142
ISN 0143
ISN 0144
ISN 0145
ISN 0146
ISN 0147
ISN 0148
ISN 0149
ISN 0150
ISN 0151

      IF ((12*(MTA/2))-WE-NYA) QQQ = (MTA - 1) / 2
      RATIO = 1.36 / DSQRT(QQQ)
      XS(1) = 0.0
      YS(1) = RATIO
      XS(2) = 0.5 * (1.0-RATIO)
      YS(2) = 1.0
      CALL CURVE(XS,YS,2.0)
      XS(1) = 0.5 * RATIO
      YS(1) = 0.0
      XS(2) = 0.5
      YS(2) = (1.0 - RATIO)
      CALL CURVE(XS,YS,2.0)
      CALL FRAME
      CALL ENDPL(0)
      RETURN
C
C
ISN 0152
      END

```

***** C I T R A N C R O S S R E P E R E N C E L I S T : N G 0 0 0 0 0

INTERNAL STATEMENT NUMBERS										
SYMBOL	INTERNAL STATEMENT NUMBERS									
A	0011	0105	C106							
I	0084	0085								
S	0007	0099	0069	0099	0106					
V	0307	0066	0050	0061	0092					
X	0007	0028								
Z	0007									
KP	0007									
MC	0091		0098	0102	C105	0109				
ID	0007									
IS	0006	0017	0020	0021	0024	0028	0030	0037	0040	0041
	0072	0078	0080	0082	0083	0085	0087	0092	0096	0099
	0139	0141	0143	0144	0146	0148				
	0006	0017	0024	0028	0030	0037	0043	0045	0047	0050
YS	0082	0084	0086	0087	0092	0096	0099	0103	0106	0110
	0145	0147	0148							
ACS	0007									
ANP	0007									
ASP	0007	0066								
BSH	0007	0069								
FPE	0007	0114								
KRP1	0065	0066	0068	0069	0072	0098	0099	0103	0105	0106
MC1	0091	0092	0096							0126
NAB	0005	0008	0018	0114						
NFA	0007	0017	0021	0024	0028	0030	0037	0041	0043	0044
	0075	0075	0113	0116	C134	0134	0134	0136	0136	0136
QQ	0073	0075	0077	0134	0136	0138				
XAZ	0004	0004	0010	0115	0116					
YAX	0004	0010								
ACFN	0011	0091	C092							
DASH	0029									
DATE	0007	0017	0028	0037						
PPEH	0009	0127	0130							
GRAP	0022	0042	0053	C071	0132					
IDUH	0010									
MPER	C126	0127	C129	0130	0133					
NFA2	0113	0114	0122							
PRR2	0009									
VRAN	0007									
YAX	0017	0021	0022	0037	0041	0042	0050	0053	0061	0063
YXAX	0130	0127	C129	0130	C132					
YEAR	0007									
YLOG	0121									
YAX1	0017	0022	0037	0042	0050	0053	0061	0063	0066	0068
	0129	0130	0132							
CATCH	0007	0017								
CURVE	0028	0030	0043	0047	0055	0072	0082	0087	0122	0133
DOPLT	0002									0143
DOPLT	0002									0148
DSORT	0077	0138								
DTIME	0007									
ENDPL	0034	0049	0058	C089	C124	0150				
FLOAT	0116	0118								
FRANZ	0033	0048	0057	0088	0123	0149				
INTNO	0035	0102	0109							
INLOT	0007									
INVILT	0007									
VSXEP	0117	0118								
FEEDG	0009	00127								

*****PORTMAN CROSS REFERENCE LISTING*****

SYMBOL	INTERNAL STATEMENT NUMBERS
RATIO	0077 0079 0080 0083 0086 0138 0140 0141 0144 0147
RESET	0025 0031 0056
STATE	0005 0008 0018 0038
TITLE	0018 0038 0051 0062 0067 0070 0073 0100 0107 0119 0128 0131
XDIFP	0007 0037 0050 0061
XLONG	0013 0018 0038 0062 0067 0070 0093 0100 0107 0115 0119 0128 0131
XORIG	0004 0010 0017 0020 0022 0037 0040 0042 0050 0053 0061 0063 0066 0068 0069 0071 0092 0096 0099
	0103 0106 0110 0121 0127 0129 0130 0132
XSTEP	0004 0010 0116 0117 0118 0121
YLONG	0014 0018 0038 0062 0067 0070 0093 0094 0095 0100 0101 0102 0107 0108 0109 0119 0128 0131
YORIG	0004 0010 0017 0022 0037 0042 0050 0053 0061 0063 0066 0068 0069 0071 0092 0096 0099 0103 0106
	0110 0121 0127 0129 0130 0132
YTIME	0005 0008 0018 0038
ANGTIC	0019 0039 0052
CALCRP	0012
CUMPER	0009 0130
DOPLT1	0016
DOPLT2	0036
DOPLT3	0060
DOPLT4	0065
DOPLT5	0091
DOPLT6	0096
DOPLT7	0105
DOPLT8	0112
DOPLT9	0126
GPMBGM	0017 0037 0050 0061 0066 0069 0092 0099 0106 0127 0130
GPMBG1	0114
GPMBG2	0063 0068 0096 0103 0110 0129
JAUTCV	0007
IFPEPL	0007
INTSP	0007
IPRINT	0007
ITEND	0007 0076
LYWACV	0007
MARKER	0023 0054
RCOUNT	0007
MESSAG	0094 0101 0108
PERIOD	0007
SINGLE	0028
ITICKS	0120
TCYCLE	0004 0010 0121

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*****FORTRAN CROSS REFERENCE LISTING*****

LABEL	DEFINED	REFERENCES
199	0032	0026
300	0046	0044

```

LEVER 21.0 ( JUN 79 )
05/760 PORTMAN B
COMPILE OPTIONS - NAME= NAME,OPT=02,LIB=00,SIZE=0000,
30000,RECDIC,PCLIST,NOCHK,LOAD,N,P,NOEDIT,LD,NAME
SUPPORTE SIMPL (KS,IS,KD,IB,M)
C
NAME= KD(1),YM(1)
C
DIRECTION IS(1),YS(1)
C
DO 100 I=1,N
  IS(1) = IS(1)
  YS(1) = YS(1)
  100 CONTINUE
C
NAME=
END
150 0002
150 0003
150 0004
150 0005
150 0006
150 0007
150 0008
C
150 0009
150 0010
END

```

```

00000 D B T A N C B C S S R R P U M M C E L I S Y I N G O O O O
INTERNAL STATEMENT NUMBERS
I 0005 0006 0006 0007 0007
B 0002 0005
XB 0002 0003 0006
XS 0002 0004 0006
YB 0002 0003 0007
YS 0002 0004 0007
ZIGGLE 0002

```

DATA ON TRAN CROSS REFERENCE LISTING 000000

NAME	DEFINED	REFERENCES
100	0000	0003


```

LEVEL 21.0 ( JUL 74 )
                                05/360 FORTRAN B
COMPILE OPTIONS - NAME= RATE,OPT=03,LIBR=40,X1B=0000K,
SOURCE,INCDIC,NOIST,ROBCK,LOAD,MAP,ROBKIT,LD,REBP
SUBROUTINE DPNDN(TS,YS,X,Y,N,TORIG,THAT,TCHIS,THAZ)
C
C
C
COMMON /OBS1-2/ XCA,ISTP,TOR,TSPP,TAX,TAE,IBDN(3)
REAL*8 X,Y
DIMENSION X(1),Y(1),X(1),Y(1)
CALL STORLX(TS,YS,X,Y,N)
CALL STORLX(-1)
CALL PAGE(9,5,10,5)
CALL SLIPIT
CALL SLIPIT(TS,YS,')
TORIG = TOR
THAZ = TOR + ISTP * TAX
THAZ = TOR + TSPP * TAX
C
C
C
END

```

***** CROSS REFERENCE LIST *****

INTERNAL STATEMENT NUMBERS

SYMBOL

1	0001	0001	
2	0002	0002	
3	0003	0003	0004
4	0004	0004	0005
5	0005	0005	0006
6	0006	0006	0007
7	0007	0007	0008
8	0008	0008	0009
9	0009	0009	0010
10	0010	0010	0011
11	0011	0011	0012
12	0012	0012	0013
13	0013	0013	0014
14	0014	0014	0015
15	0015	0015	0016
16	0016	0016	0017
17	0017	0017	0018
18	0018	0018	0019
19	0019	0019	0020
20	0020	0020	0021
21	0021	0021	0022
22	0022	0022	0023
23	0023	0023	0024
24	0024	0024	0025
25	0025	0025	0026
26	0026	0026	0027
27	0027	0027	0028
28	0028	0028	0029
29	0029	0029	0030
30	0030	0030	0031
31	0031	0031	0032
32	0032	0032	0033
33	0033	0033	0034
34	0034	0034	0035
35	0035	0035	0036
36	0036	0036	0037
37	0037	0037	0038
38	0038	0038	0039
39	0039	0039	0040
40	0040	0040	0041
41	0041	0041	0042
42	0042	0042	0043
43	0043	0043	0044
44	0044	0044	0045
45	0045	0045	0046
46	0046	0046	0047
47	0047	0047	0048
 

1131100000

DEPT OF COMMERCE

CROSS REFERENCE

1131100000

REFERENCE

DEPT OF COMMERCE

1131100000

SCHOOL	INTERNAL STATEMENT NUMBERS
Z	0002 0009
ZS	0003 0003 0009
ZS	0002 0003 0009
GRAP	0000
GRAY	0002 0008
GRAY	0002 0008
CUTTER	0011
FRANK	0010
XORIG	0002 0008
YOLITE	0002 0008
SPRING	0002
KAZAM	0006
TITLES	0004
FAX/WG	0007
TITLES	0005

```

LEVEL 21.0 ( JUN 74 )
                                OS/360 FORTRAN E
                                COMPILER OPTIONS - NAME= MAIN,OPT=02,LINCHT=40,SIZE=0000R,
                                SOURCE=,ESCDIC,MOLIST,MORSECH,LOAD,MIP,MORDIT,ID,YRNP
                                SUBROUTINE AMOTIC
ISB 0002      C
ISB 0003      C
ISB 0004      CALL YTHICKS(3)
ISB 0005      CALL YTHICKS(5)
ISB 0006      CALL YTHICKS(45.)
ISB 0007      CALL YTHICKS(0.)
ISB 0008      C
ISB 0009      C
ISB 0010      RETURN
ISB 0011      END

```


SYMBOL INTERNAL STATEMENT NUMBERS
 ANOTIC 0002
 YANAB 0005
 YICKS 0003
 YANAB 0006
 YICKS 0004

*****
 CROSS REFERENCE LISTING

APPENDIX IV
LISTING OF SUBROUTINE DETRND

OS/160 PORTMAN N

ISB 0002

1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 19 20 21 22 23 24 25 26 27 28 29 30 31 32 33 34 35 36 37 38 39 40 41 42 43 44 45 46 47 48 49 50 51 52 53 54 55 56 57 58 59 60 61 62 63 64 65 66 67 68 69 70 71 72 73 74 75 76 77 78 79 80 81 82 83 84 85 86 87 88 89 90 91 92 93 94 95 96 97 98 99 100 101 102 103 104 105 106 107 108 109 110 111 112 113 114 115 116 117 118 119 120 121 122 123 124 125 126 127 128 129 130 131 132 133 134 135 136 137 138 139 140 141 142 143 144 145 146 147 148 149 150 151 152 153 154 155 156 157 158 159 160 161 162 163 164 165 166 167 168 169 170 171 172 173 174 175 176 177 178 179 180 181 182 183 184 185 186 187 188 189 190 191 192 193 194 195 196 197 198 199 200 201 202 203 204 205 206 207 208 209 210 211 212 213 214 215 216 217 218 219 220 221 222 223 224 225 226 227 228 229 230 231 232 233 234 235 236 237 238 239 240 241 242 243 244 245 246 247 248 249 250 251 252 253 254 255 256 257 258 259 260 261 262 263 264 265 266 267 268 269 270 271 272 273 274 275 276 277 278 279 280 281 282 283 284 285 286 287 288 289 290 291 292 293 294 295 296 297 298 299 300 301 302 303 304 305 306 307 308 309 310 311 312 313 314 315 316 317 318 319 320 321 322 323 324 325 326 327 328 329 330 331 332 333 334 335 336 337 338 339 340 341 342 343 344 345 346 347 348 349 350 351 352 353 354 355 356 357 358 359 360 361 362 363 364 365 366 367 368 369 370 371 372 373 374 375 376 377 378 379 380 381 382 383 384 385 386 387 388 389 390 391 392 393 394 395 396 397 398 399 400 401 402 403 404 405 406 407 408 409 410 411 412 413 414 415 416 417 418 419 420 421 422 423 424 425 426 427 428 429 430 431 432 433 434 435 436 437 438 439 440 441 442 443 444 445 446 447 448 449 450 451 452 453 454 455 456 457 458 459 460 461 462 463 464 465 466 467 468 469 470 471 472 473 474 475 476 477 478 479 480 481 482 483 484 485 486 487 488 489 490 491 492 493 494 495 496 497 498 499 500 501 502 503 504 505 506 507 508 509 510 511 512 513 514 515 516 517 518 519 520 521 522 523 524 525 526 527 528 529 530 531 532 533 534 535 536 537 538 539 540 541 542 543 544 545 546 547 548 549 550 551 552 553 554 555 556 557 558 559 560 561 562 563 564 565 566 567 568 569 570 571 572 573 574 575 576 577 578 579 580 581 582 583 584 585 586 587 588 589 590 591 592 593 594 595 596 597 598 599 600 601 602 603 604 605 606 607 608 609 610 611 612 613 614 615 616 617 618 619 620 621 622 623 624 625 626 627 628 629 630 631 632 633 634 635 636 637 638 639 640 641 642 643 644 645 646 647 648 649 650 651 652 653 654 655 656 657 658 659 660 661 662 663 664 665 666 667 668 669 670 671 672 673 674 675 676 677 678 679 680 681 682 683 684 685 686 687 688 689 690 691 692 693 694 695 696 697 698 699 700 701 702 703 704 705 706 707 708 709 710 711 712 713 714 715 716 717 718 719 720 721 722 723 724 725 726 727 728 729 730 731 732 733 734 735 736 737 738 739 740 741 742 743 744 745 746 747 748 749 750 751 752 753 754 755 756 757 758 759 760 761 762 763 764 765 766 767 768 769 770 771 772 773 774 775 776 777 778 779 780 781 782 783 784 785 786 787 788 789 790 791 792 793 794 795 796 797 798 799 800 801 802 803 804 805 806 807 808 809 810 811 812 813 814 815 816 817 818 819 820 821 822 823 824 825 826 827 828 829 830 831 832 833 834 835 836 837 838 839 840 841 842 843 844 845 846 847 848 849 850 851 852 853 854 855 856 857 858 859 860 861 862 863 864 865 866 867 868 869 870 871 872 873 874 875 876 877 878 879 880 881 882 883 884 885 886 887 888 889 890 891 892 893 894 895 896 897 898 899 900 901 902 903 904 905 906 907 908 909 910 911 912 913 914 915 916 917 918 919 920 921 922 923 924 925 926 927 928 929 930 931 932 933 934 935 936 937 938 939 940 941 942 943 944 945 946 947 948 949 950 951 952 953 954 955 956 957 958 959 960 961 962 963 964 965 966 967 968 969 970 971 972 973 974 975 976 977 978 979 980 981 982 983 984 985 986 987 988 989 990 991 992 993 994 995 996 997 998 999 1000 1001 1002 1003 1004 1005 1006 1007 1008 1009 1010 1011 1012 1013 1014 1015 1016 1017 1018 1019 1020 1021 1022 1023 1024 1025 1026 1027 1028 1029 1030 1031 1032 1033 1034 1035 1036 1037 1038 1039 1040 1

A = MATRIX OF DIMENSION NYA*1 BY NLSCOL+1. USED IN BLSQ.
CATCH = INPUT TIME SERIES VALUES.
COEF = VECTOR CONTAINING THE COEFFICIENTS OF THE LINEAR FIT.
COEF2 = WORKING VECTOR OF SAME LENGTH AS COEF.
DATE = TIME VALUES FOR INPUT SERIES.
NLSCOL = NO. OF COEFFICIENTS OF THE LINEAR FIT.
NRNW = ROW DIMENSION OF MATRIX A.
NYA = NO. OF DATA POINTS IN TIME SERIES.
X = LINEAR TREND OF TIME SERIES (HEREIN CALCULATED AND RETURNED).
XDIPP = DETRENDED TIME SERIES (HEREIN CALCULATED AND RETURNED).
YEAR = TIME VARIABLE (HEREIN CALCULATED).

ISN 0003
ISN 0004
ISN 0005
ISN 0006

```

IMPLICIT REAL*8(A-H,O-Z)
DIMENSION DATE(1),YEAR(1),CATCH(1),XDIFF(1),K(1)
DIMENSION A(601,3),IRON(601),CORP(10),T(2,2),T1(2,2),VAN(2,2)
DIMENSION CORP2(10)

```

cccc

INITIALIZE VALUES.

ISN 0007
ISN 0008
ISN J009
ISN 0010
ISN 0011
ISN 0012
ISN 0013
ISN 0014
ISN 0015

```

      NROW = 601
      NLSCOL = 2
      NT = 2
      SPPE = NLSCOL
      DO 50 I=1,NYA
      YEAR(I) = I - .5
      A(I,1) = 1.0
      A(I,2) = YEAR(I)
50 CONTINUE

```

CCCC

CALL LEAST SQUARES ROUTINE.

ISN 0016
ISN 0017
ISN 0018

```
CALL NLSC (A,CATCH,CORP,RESID,IRUN,T,T1,NYA,NLSCOL,NBOW,NLSCOL)
DO 80 I=1,NLSCOL
SUM = 0.
```

```

158 0019      DO 70 J=1,NLSCOL
159 0020      SUM = SUM + Y(I,J)*CORP(J)
160 0021      70 CONTINUE
161 0022      CORP2(I) = SUM
162 0023      80 CONTINUE
163 0024      DO 90 I=1,NLSCOL
164 0025      CORP(I) = CORP2(I)
165 0026      90 CONTINUE
166 0027      C
167 0028      DO 100 I=1,NYA
168 0029      X(I) = A(I,NLSCOL+1)
169 0030      X(I) = CATCH(I) - X(I)
170 0031      100 CONTINUE
171 0032      C
172 0033      PRINT THE COEFFICIENTS OF THE TEND LINE.
173 0034      C
174 0035      PRINT 130, RESID
175 0036      130 FORMAT (180,'RESIDUAL SUM OF SQUARES = ', 1P19.6)
176 0037      PRINT 150
177 0038      150 FORMAT (180//180,'COEFFICIENTS OF THE LINEAR FIT'//)
178 0039      PRINT 200, (CORP(I),I=1,NLSCOL)
179 0040      200 FORMAT (180/180,'INTERCEPT = ',1P19.6,3X,'SLOPE = ',1P19.6//)
180 0041      C
181 0042      COMPUTE THE VARIANCE-COVARIANCE MATRIX.
182 0043      C
183 0044      RESLT = RESID / (NYA-NLSCOL)
184 0045      DO 400 I=1,NLSCOL
185 0046      DO 400 J=1,NLSCOL
186 0047      SUM = 0.
187 0048      DO 300 K=1,NLSCOL
188 0049      SUM = SUM + Y(I,K)*Y(J,K)
189 0050      300 CONTINUE
190 0051      VAR(I,J) = RESLT * SUM
191 0052      400 CONTINUE
192 0053      PRINT 500
193 0054      500 FORMAT (180//180,'VARIANCE-COVARIANCE MATRIX'//)
194 0055      CALL ABRTD(VAR,NLSCOL,NLSCOL,NY,NY)
195 0056      C
196 0057      RETURN
197 0058      END
198 0059      C
199 0060
200 0061
201 0062
202 0063
203 0064
204 0065
205 0066
206 0067
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211 0072
212 0073
213 0074
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348 0209
349 0210
350 0211
351 0212
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353 0214
354 0215
355 0216
356 0217
357 0218
358 0219
359 0220
360 0221
361 0222
362 0223
363 0224
364 0225
365 0226
366 0227
 

```

```

***** O R T B A N   E R O S S   R E P E R E N C E   L I S T I N G *****

SYMBOL  INTERNAL STATEMENT NUMBERS
A 0005 0013 0016 0028
1 0011 0012 0012 0013 0016
0035 0035 0038 0042 0046
2 0019 0020 0020 0039 0042 0044
K 0041 0042 0042
T 0005 0016 0021 0029
BT 0002 0004 0009 0008
T1 0005 0016 0020 0042 0042
BTA 0002 0011 0016 0027 0037
SBB 0018 0020 0020 0042 0042 0044
VAR 0005 0044 0044
BL30 0016 0016 0020 0025 0035
COMP 0005 0004 0002 0004
DATE 0002 0004
INOP 0005 0016
RCOM 0007 0016
3773 0002 0010
TEAR 0002 0004 0012 0014
CATCH 0002 0004 0016 0025
COMP2 0006 0022 0025
BAGLT 0037 0044
RESID 0016 0031 0037
IDIFF 0002 0004 0029
ARRATO 0048
DEFEND 0002
BLSCOL 0006 0010 0016 0016 0017 0019 0024 0024 0028 0035 0037 0038 0039 0041 0048 0048

*****

```

***** C R O S S R E F E R E N C E S *****

1.332L	DEPT 000	REFERENCES
50	0015	0011
70	0021	0019
80	0023	0017
90	0026	0020
100	0030	0027
130	0032	0031
150	0030	0033
200	0036	0035
300	0043	0041
400	0045	0039
500	0047	0046

***** C R O S S R E F E R E N C E L I S T *****

SYMBOL	INTERNAL STATEMENT NUMBERS
B	0002 0003 0030 0030 0030 0030
J	0023 0023 0023 0030 0030 0030
L	0027 0028 0029 0030 0030 0030
LI	0015 0016 0018 0030 0030 0030
LV	0007 0013 0018 0030 0030 0030
BN	0009 0010 0012 0018 0023 0023
BL	0008 0009 0013 0023 0023 0023
BC	0004 0004 0010 0012 0012 0012
BB	0017 0018 0020 0027 0027 0027
BL	0016 0017 0027 0027 0027 0027
BB	0005 0015 0018 0020 0020 0020
LVV	0002 0007 0018 0020 0020 0020
BBB	0029 0030 0030 0030 0030 0030
BAL	0020 0030 0030 0030 0030 0030
BCC	0002 0004 0004 0004 0004 0004
BPR	0042 0005 0005 0005 0005 0005
BMM1	0018 0029 0029 0029 0029 0029
BMM1	0013 0029 0029 0029 0029 0029
BMM2	0002 0021 0021 0021 0021 0021
ARRATO	0002 0004 0004 0004 0004 0004
COLUMB	0003 0003 0003 0003 0003 0003

***** O T R A N C R O S S R E F E R E N C E L I S T I N G *****

Label	DEFINITION	REFERENCE
1	0013	0010
2	0021	0010
3	0030	0000 0015 0027
100	0022	0021
200	0020	0023
300	0026	0025
600	0031	0030